



## 9. Multi-dimensional Itô calculus

- Let  $X$  and  $Y$  be two Itô processes.
- $P = \{0 = t_1 < t_1 \cdots < t_n = T\}$  is a partition of  $[0, T]$ .



**Definition 9.1.** The joint quadratic variation of  $X, Y$ , is defined by

$$\underline{[X, Y]}_T = \lim_{\|P\| \rightarrow 0} \sum_{i=0}^{n-1} (X_{t_{i+1}} - X_{t_i})(Y_{t_{i+1}} - Y_{t_i}),$$

*Remark 9.2.* The joint quadratic variation is sometimes written as  $\underline{d[X, Y]}_t = \underline{dX}_t \underline{dY}_t$ .

$$4ab = \underline{(a+b)^2} - \underline{(a-b)^2}$$

**Lemma 9.3.**  $[X, Y]_T = \frac{1}{4}([X + Y, X + Y]_T - [X - Y, X - Y]_T)$

Joint QV

QV of  $X+Y$

QV of  $X-Y$

**Proposition 9.4.** Say  $X, Y$  are two semi-martingales.

- Write  $X = X_0 + B + M$ , where  $B$  has bounded variation and  $M$  is a martingale.
- Write  $\underline{Y} = \underline{Y}_0 + \underline{C} + \underline{N}$ , where  $C$  has bounded variation and  $N$  is a martingale.
- Then  $\underline{d}[X, \underline{Y}]_t = \underline{d}[M, \underline{N}]_t$ .

Remark 9.5. Recall, all processes are implicitly assumed to be adapted and continuous.

**Corollary 9.6.** If  $X$  is a semi-martingale and  $B$  has bounded variation then  $[X, B] = 0$ .

$$\hookrightarrow \textcircled{1} \quad [X, B] = \frac{1}{4} \left( \underbrace{[X+B, X+B]}_{[X, X]} - \underbrace{[X-B, X-B]}_{[X, X]} \right) \Bigg\} = 0$$

$$\textcircled{2} \quad X = \underbrace{X_0}_{\text{wavy}} + \underbrace{C}_{\text{wavy}} + \underbrace{M}_{\text{red box}} \quad \Bigg\} [X, B] = [M, 0] = 0$$
$$\hat{=} \quad B = \underbrace{B_0}_{\text{wavy}} + \underbrace{(B-B_0)}_{\text{wavy}} + \underbrace{0}_{\text{red box}} \quad \Bigg\}$$

Remark 9.7 (Chain rule). If  $X, Y$  are differentiable functions of  $t$ , then

$$d(f(t, X_t, Y_t)) = \partial_t f(t, X_t, Y_t) dt + \partial_x f(t, X_t, Y_t) dX_t + \partial_y f(t, X_t, Y_t) dY_t$$

Remark 9.8 (Notation).  $\partial_t f = \frac{\partial f}{\partial t}$ ,  $\partial_x f = \frac{\partial f}{\partial x}$ ,  $\partial_y f = \frac{\partial f}{\partial y}$ .

$$f = f(t, x, y)$$

$$\frac{d}{dt} f(t, X_t, Y_t) \stackrel{\text{Chain Rule}}{=} \partial_t f(t, X_t, Y_t) dt + \partial_x f(t, X_t, Y_t) \frac{dX_t}{dt} dt + \partial_y f(t, X_t, Y_t) \frac{dY_t}{dt} dt$$

If  $X$  &  $Y$  are not diff, then

there is an extra term  $\circ \frac{1}{2} \left[ \partial_x^2 f d[X, X] + \partial_y^2 f d[Y, Y] + 2\partial_x \partial_y f d[X, Y] \right]$

**Theorem 9.9** (Two-dimensional Itô formula).

- Let  $X, Y$  be two processes.
- Let  $f = f(t, x, y)$  be a  $C^{1,2}$  function. That is:
  - ▷  $f$  is once differentiable in  $t$
  - ▷  $f$  is twice in both  $x$ , and  $y$ .
  - ▷ All the above partial derivatives are continuous. Then:

$$d(f(t, X_t, Y_t)) = \partial_t f(t, X_t, Y_t) dt + \partial_x f(t, X_t, Y_t) dX_t + \partial_y f(t, X_t, Y_t) dY_t + \frac{1}{2} \left[ \partial_x^2 f(t, X_t, Y_t) d[X, X]_t + \partial_y^2 f(t, X_t, Y_t) d[Y, Y]_t + 2\partial_x \partial_y f(t, X_t, Y_t) d[X, Y]_t \right]$$

from chain rule

$2\partial_x \partial_y f$

*Remark 9.10.* As with the 1D Itô, will drop the arguments  $(t, X_t, Y_t)$ . Remember they are there.

*Remark 9.11* (Integral form of Itô's formula).

$$\underbrace{f(T, X_T, Y_T)} - f(0, X_0, Y_0) = \int_0^T \partial_t f dt + \int_0^T \partial_x f dX_t + \int_0^T \partial_y f dY_t + \frac{1}{2} \int_0^T (\partial_x^2 f d[X, X]_t + \partial_y^2 f d[Y, Y]_t + 2\partial_x \partial_y f d[X, Y]_t)$$

Intuition behind Theorem 9.9.

① 1D  $f: \mathbb{R} \rightarrow \mathbb{R}$ :  $\partial_x^2 f \approx [X, X]$  come from Taylor

$$f(x+h) = f(x) + h f'(x) + \frac{1}{2} h^2 f''(x) + \text{small}$$

② In 2D: Taylor's formula is

$$f(x+h, y+k) \approx f(x, y) + h \partial_x f(x, y) + k \partial_y f(x, y)$$

$$+ \frac{1}{2} \left[ \underbrace{\partial_x^2 f}_{\text{QV}(x)} h^2 + \underbrace{\partial_y^2 f}_{\text{QV}(y)} k^2 + \underbrace{2 \partial_x \partial_y f}_{\text{Joint QV.}} hk \right]$$



**Proposition 9.12** (Product rule).  $d(XY)_t = X_t dY_t + Y_t dX_t + d[X, Y]_t$

If  $f$  &  $g$  are diff func,

$$\frac{d}{dt}(fg) = \frac{df}{dt} g + f \frac{dg}{dt}$$

Check product rule using Ito!

let  $f(x, y) = xy$

$$\frac{\partial^2 f}{\partial t^2} = 0$$

$$\frac{\partial^2 f}{\partial x^2} = 0$$

$$\frac{\partial^2 f}{\partial y^2} = 0$$

$$\frac{\partial^2 f}{\partial x \partial y} = 1$$

$$\frac{\partial f}{\partial x} = y$$

$$\frac{\partial f}{\partial y} = x$$

$$\text{By Ito: } d\left(\frac{X_t Y_t}{t}\right) = d\left(\frac{f(X_t, Y_t)}{t}\right)$$

$$= \frac{\partial f}{\partial t} dt + \frac{\partial f}{\partial x} dX_t + \frac{\partial f}{\partial y} dY_t$$

$$+ \frac{1}{2} \left[ \frac{\partial^2 f}{\partial x^2} d[X, X] + \frac{\partial^2 f}{\partial y^2} d[Y, Y] + 2 \frac{\partial^2 f}{\partial x \partial y} d[X, Y] \right]$$

$$= 0 + \frac{Y_t}{t} dX_t + X_t d\frac{Y_t}{t} + \frac{1}{t} \cancel{1} \underline{d[X, Y]}$$

$\Rightarrow$  Product Rule.

To use the multi-dimensional Itô formula, we need to compute joint quadratic variations.

**Proposition 9.13.** Let  $\underline{M}, \underline{N}$  be continuous martingales, with  $\mathbf{E}M_t^2 < \infty$  and  $\mathbf{E}N_t^2 < \infty$ .

(1)  $\underline{MN} - [\underline{M}, \underline{N}]$  is also a continuous martingale.

(2) Conversely if  $\underline{MN} - \underline{B}$  is a continuous martingale for some continuous adapted, bounded variation process  $B$  with  $B_0 = \underline{0}$ , then  $B = [\underline{M}, \underline{N}]$ .

Proof.

~~NTS~~ NTS  $\underline{MN} - [\underline{M}, \underline{N}]$  is a mg

$$\begin{aligned} d(\underline{MN} - [\underline{M}, \underline{N}]) & \stackrel{\text{Product rule}}{=} M dN + N dM + d[\underline{M}, \underline{N}] - d[\underline{M}, \underline{N}] \\ & = \underbrace{M dN + N dM}_{\text{mg}} \end{aligned}$$

**Proposition 9.14.** (1) (Symmetry)  $[X, Y] = [Y, X]$

(2) (Bi-linearity) If  $\alpha \in \mathbb{R}$ ,  $X, Y, Z$  are semi-martingales,  $[X, Y + \alpha Z] = [X, Y] + \alpha [X, Z]$ .

*Proof.*

**Proposition 9.15.** Let  $\underline{M}, \underline{N}$  be two martingales,  $\sigma, \tau$  two adapted processes.

• Let  $\underline{X}_t = \int_0^t \underline{\sigma}_s d\underline{M}_s$  and  $\underline{Y}_t = \int_0^t \underline{\tau}_s d\underline{N}_s$ .

• Then  $[X, Y]_t = \int_0^t \sigma_s \tau_s d[M, N]_s$ . ← Use of our rule to compute QV of Ito int.

Remark 9.16. In differential form, if  $dX_t = \sigma_t dM_t$  and  $dY_t = \tau_t dN_t$ , then  $d[X, Y]_t = \sigma_t \tau_t d[M, N]_t$ .

Intuition.

$$\Delta_i X = X_{t_{i+1}} - X_{t_i} \approx \sigma_{t_i} \Delta_i M$$

$$\Delta_i Y = \tau_{t_i} \Delta_i N$$

$$\Rightarrow \underbrace{\Delta_i X \Delta_i Y}_{d[X, Y]} \approx \underbrace{\sigma_{t_i} \tau_{t_i} \Delta_i M \Delta_i N}_{\approx d[M, N]}$$

**Proposition 9.17.** If  $M, N$  are continuous martingales,  $EM_t^2 < \infty$ ,  $EN_t^2 < \infty$  and  $M, N$  are independent, then  $[M, N] = 0$ .

*Remark 9.18 (Warning).* Independence implies  $E(M_t N_t) = EM_t EN_t$ . But it does not imply  $E_s(M_t N_t) = E_s M_t E_s N_t$ . So you can't use this to show  $MN$  is a martingale, and hence conclude  $[M, N] = 0$ .

Correct proof.

→ Tricky

$$\text{Intuition: } [M, N]_T = \lim_{\|P\| \rightarrow 0} \sum \Delta_i M \Delta_i N$$

$$\text{Trick: Compute } E \left( \sum \Delta_i M \Delta_i N \right)^2 \xrightarrow[\|P\| \rightarrow 0]{\text{Will show}} 0$$

$$\textcircled{1} E\left(\sum (\Delta_i M)(\Delta_i N)\right)^2 = E \sum (\Delta_i M)^2 (\Delta_i N)^2 + 2E \sum_{i < j} (\Delta_i M)(\Delta_j M) (\Delta_i N)(\Delta_j N)$$

$$= \underbrace{\sum E(\Delta_i M)^2 E(\Delta_i N)^2}_{\downarrow} + 2 \sum_{i < j} \underbrace{E(\Delta_i M \Delta_j M) E(\Delta_i N \Delta_j N)}_{\text{Use tower \& } M \text{ is a mg check} = 0}$$

Use tower &  $M$  is a mg  
check = 0

$$\left\langle \left[ \max_i E(\Delta_i, M)^2 \right] \right\rangle \quad \left\langle \sum_i E(\Delta_i, N)^2 \right\rangle$$

$$\downarrow |P| \rightarrow 0$$

$$0$$

( $\Delta$  is ds)

$$\downarrow |P| \rightarrow 0$$

$$E(N, N)$$



Remark 9.19.  $[M, N] = 0$  does not imply  $M, N$  are independent. For example:

- Let  $M_t = \int_0^t \mathbf{1}_{\{W_s < 0\}} dW_s$
- Let  $N_t = \int_0^t \mathbf{1}_{\{W_s \geq 0\}} dW_s$

$$\left. \begin{array}{l} \bullet \text{ Let } M_t = \int_0^t \mathbf{1}_{\{W_s < 0\}} dW_s \\ \bullet \text{ Let } N_t = \int_0^t \mathbf{1}_{\{W_s \geq 0\}} dW_s \end{array} \right\} M + N = \int_0^t (\mathbf{1}_{\{W_s < 0\}} + \mathbf{1}_{\{W_s \geq 0\}}) dW_s \\ = \int_0^t 1 dW_s = W_t$$

$$d[M, N] = \mathbf{1}_{\{W_s < 0\}} \mathbf{1}_{\{W_s \geq 0\}} d[W, W]_t$$

$$= 0 dt$$

$$\Rightarrow M_t + N_t = W_t$$

**Question 9.20.** Let  $W^1$  and  $W^2$  be two independent Brownian motions, and let  $W = (W^1, W^2)$ . Define the process  $X$  by  $X_t = \ln(|W_t|^2) = \ln((W_t^1)^2 + (W_t^2)^2)$ . Is  $X$  a martingale?

$W_t^1 \rightarrow 1^{\text{st}}$  BM at time  $t$

$W_t^2 \rightarrow 2^{\text{nd}}$  BM at time  $t$

$$W = (W^1, W^2) \quad |W| = \sqrt{(W^1)^2 + (W^2)^2}$$

$$X_t = \ln \left[ (W_t^1)^2 + (W_t^2)^2 \right]$$

Is  $X$  a martingale?

Ito's:  $f(t, x, y) = \ln(x^2 + y^2)$

Compute  $\partial_x f$ ,  $\partial_y f$ ,  $\partial_x^2 \dots$

Compute  $dX_t = d f(t, W'_t, W_t^2)$

Ito's

$$\partial_x f dW'_t + \partial_y f dW_t^2 + \frac{1}{2} \left[ \overbrace{\partial_x^2 f}^{dt} d[W'_t, W'_t] + \overbrace{\partial_y^2 f}^{dt} d[W_t^2, W_t^2] + 2 \partial_x \partial_y f d[W'_t, W_t^2] \right]$$

0 (imdep)

$$\begin{aligned} &= \partial_x f \, dW_t^1 + \partial_y f \, dW_t^2 + \\ &\quad + \frac{1}{2} \left[ \partial_{xx}^2 f + \partial_{yy}^2 f \right] dt \end{aligned}$$

Can compute & check  $\partial_{xx}^2 f + \partial_{yy}^2 f = 0$  (when  $f(x, y) = \ln(x^2 + y^2)$ )

$$\Rightarrow dx_t = \partial_x f \, dW_t^1 + \partial_y f \, dW_t^2 \Rightarrow X \text{ is a } mg^{??}$$

u

Claim: Even though there are no dt terms  
here  $X$  is NOT a mg!!!

(Reason: \_\_\_\_\_)

## 10. Risk Neutral Pricing

### Goal.

- Consider a market with a bank and one stock.
- The interest rate  $R_t$  is some adapted process.
- The stock price satisfies  $dS_t = \alpha_t S_t dt + \sigma_t S_t dW_t$ . (Here  $\alpha, \sigma$  are adapted processes).
- Find the risk neutral measure and use it to price securities.

**Definition 10.1.** Let  $D_t = \exp\left(-\int_0^t R_s ds\right)$  be the discount factor.

*Remark 10.2.* Note  $\partial_t D = -R_t D$ .

*Remark 10.3.*  $D_t$  dollars in the bank at time 0 becomes \$1 in the bank at time  $t$ .

Recall:  $C_t =$  cash in bank at time  $t$

$$\partial_t C_t = R_t C_t \Rightarrow C_t = C_0 \exp\left(\int_0^t R_s ds\right)$$

**Theorem 10.4.** The (unique) risk neutral measure is given by  $d\tilde{\mathbf{P}} = Z_T d\mathbf{P}$ , where

$\rightarrow Z_T = \exp\left(-\int_0^T \theta_t dW_t - \frac{1}{2} \int_0^T \theta_t^2 dt\right)$ ,  $\theta_t = \frac{\alpha_t - R_t}{\sigma_t}$ .

Market Price of risk.

**Theorem 10.5.** Any security can be replicated. If a security pays  $V_T$  at time  $T$ , then the arbitrage free price at time  $t$  is

$$V_t = \frac{1}{D_t} \tilde{\mathbf{E}}_t(D_T V_T) = \tilde{\mathbf{E}}_t\left(\exp\left(\int_t^T -R_s ds\right) V_T\right).$$

**Remark 10.6.** We will explain the notation  $d\tilde{\mathbf{P}} = Z_T d\mathbf{P}$  and prove both the above theorems later.

Disc time: RNP formula  $V_n = \frac{1}{D_n} \tilde{\mathbf{E}}_n(D_N V_N)$

(disc time version of

**Definition 10.7.** We say  $\tilde{P}$  is a risk neutral measure if:

- (1)  $\tilde{P}$  is equivalent to  $P$  (i.e.  $\tilde{P}(A) = 0$  if and only if  $P(A) = 0$ )
- (2)  $D_t S_t$  is a  $\tilde{P}$  martingale.

*Remark 10.8.* As before, if  $\tilde{P}$  is a new measure, we use  $\tilde{E}$  to denote expectations with respect to  $\tilde{P}$  and  $\tilde{E}_t$  to denote conditional expectations.

*Example 10.9.* Fix  $T > 0$ . Let  $Z_T$  be a  $\mathcal{F}_T$ -measurable random variable.

- Assume  $Z_T > 0$  and  $E Z_T = 1$ .
- Define  $\tilde{P}(A) = E(Z_T \mathbf{1}_A) = \int_A Z_T dP$ .
- Can check  $\tilde{E}X = E(Z_T X)$ . That is  $\int_{\Omega} X d\tilde{P} = \int_{\Omega} X Z_T dP$ .
- Notation: Write  $d\tilde{P} = Z_T dP$ .

(Need  $Z_T > 0$  to ensure  $P(A) = 0 \Leftrightarrow \tilde{P}(A) = 0$ )

**Lemma 10.10.** Let  $Z_t = E_t Z_T$ . If  $X_t$  is  $\mathcal{F}_t$ -measurable, then  $\tilde{E}_s X_t = \frac{1}{Z_s} E_s(Z_t X_t)$ .

*Proof.* You will see this in the proof of the Girsanov theorem. □

**Corollary 10.11.**  $M$  is martingale under  $\tilde{P}$  if and only if  $ZM$  is a martingale under  $P$ .

↖ RN cond Exp.



$$Q: \hat{P}(A) \in [0, 1]$$

$$\text{Check: } \textcircled{1} \hat{P}(A) = E(\mathbb{1}_A z_T) \geq 0 \quad (\because z_T > 0)$$

$$\textcircled{2} \hat{P}(A) = E(\mathbb{1}_A z_T) \leq E z_T = 1$$

$$Q: \text{Need } \hat{P}(\Omega) = \underline{1}$$

$$1 \quad \text{Note } \hat{P}(\Omega) = E(\mathbb{1}_\Omega z_T) = E z_T = 1$$

Prop:  $X_t \rightarrow F_t$  meas

$$Z_t \geq E_t Z_T$$

Then 
$$E_s X_t = \frac{1}{Z_s} E_s (Z_t X_t)$$

Claim:  $M$  is a  $\mathcal{P}$  mg  $\Leftrightarrow ZM$  is a  $\mathcal{P}$  mg.

① Say  $ZM$  is a  $P$ -mg

$$\text{Compute } \sum_{s=0}^{\infty} E_s M_t = \frac{1}{z_s} E_s (z_t M_t)$$

$$= \frac{1}{z_s} (z_s M_s) \quad (\because ZM \text{ is a } P\text{-mg})$$

$$= M_s \Rightarrow M \text{ is a } \emptyset \text{ mg.}$$