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9. Multi-dimensional Itô calculus

- Let X and Y be two Itô processes.
- $P = \{0 = t_0 < t_1 < \dots < t_n = T\}$ is a partition of $[0, T]$.



$$\|P\| = \max_i (t_{i+1} - t_i)$$

Definition 9.1. The joint quadratic variation of $\underline{X}, \underline{Y}$, is defined by

$$\underline{[X, Y]}_T = \lim_{\|P\| \rightarrow 0} \sum_{i=0}^{n-1} \underbrace{(X_{t_{i+1}} - X_{t_i})}_{\Delta_i X} \underbrace{(Y_{t_{i+1}} - Y_{t_i})}_{\Delta_i Y}, \quad \lim_{\|P\| \rightarrow 0} \sum_i (\Delta_i X) (\Delta_i Y)$$

Remark 9.2. The joint quadratic variation is sometimes written as $d[X, Y]_t = dX_t dY_t$.

$$QV: \quad \underline{[X, X]}_T = \lim_{\|P\| \rightarrow 0} \sum_{i=0}^n \underbrace{(\Delta_i X)^2}$$

$$\Delta_i X = X_{t_{i+1}} - X_{t_i}$$

Lemma 9.3. $[X, Y]_T = \frac{1}{4}([X + Y, X + Y]_T - [X - Y, X - Y]_T)$

\uparrow
 Joint QV of X & Y
 $\underbrace{\hspace{10em}}_{\text{QV of } X+Y}$ $\underbrace{\hspace{10em}}_{\text{QV of } X-Y}$

Pf: $(a+b)^2 - (a-b)^2 = 4ab$

$$\Rightarrow \lim_{\|P\| \rightarrow 0} \sum_i \Delta_i(\overset{X}{\cancel{X}}) \Delta_i(\overset{Y}{\cancel{X}}) = \lim \sum_i \frac{1}{4} \left[(\Delta_i(X+Y))^2 - (\Delta_i(X-Y))^2 \right]$$

\downarrow $[X, Y]_T$ \downarrow $\frac{1}{4} ([X+Y, X+Y]_T - [X-Y, X-Y]_T)$

Proposition 9.4 (Product rule). $d(\underline{XY})_t = X_t dY_t + Y_t dX_t + d[X, Y]_t$

If X & Y are diff,

$$\frac{d}{dt}(XY) = X \frac{dY}{dt} + \frac{dX}{dt} Y$$

$$\Rightarrow d(XY) = X dY + Y dX$$

If X, Y are stoch processes (not diff)

$$d(XY) = X dY + Y dX + d[X, Y]$$

$$\text{Pf: } 4XY = (X+Y)^2 - (X-Y)^2$$

$$\text{Ito: } d\left(\frac{(X+Y)^2}{t}\right) = 2\left(\frac{X+Y}{t}\right) d\left(\frac{X+Y}{t}\right) + \frac{1}{2} \cdot 2 d[X+Y, X+Y]$$

$$d\left(\frac{(X-Y)^2}{t}\right) = 2\left(\frac{X-Y}{t}\right) d\left(\frac{X-Y}{t}\right) + d[X-Y, X-Y]_t$$

$$= 2X_t dx_t + 2Y_t dy_t - 2Y_t dx_t - 2X_t dy_t$$

$$+ d[X-Y, X-Y]_t.$$

$$\begin{aligned} \Rightarrow d\left((x+y)^2 - (x-y)^2\right) \\ = 4X_t dY_t + 4Y_t dX_t + d\left(\begin{array}{l} [x+y, x+y]_t \\ - [x-y, x-y]_t \end{array}\right) \end{aligned}$$

$$\Rightarrow 4 d(xy) = 4X dY + 4Y dX + 4 d[x, Y] \Rightarrow \text{done!}$$

Proposition 9.5. Say X, Y are two semi-martingales.

- Write $X = \underline{X}_0 + \underline{B} + \underline{M}$, where B has bounded variation and M is a martingale.
- Write $Y = \underline{Y}_0 + \underline{C} + \underline{N}$, where C has bounded variation and N is a martingale.
- Then $d[\underline{X}, \underline{Y}]_t = d[\underline{M}, \underline{N}]_t$.

Remark 9.6. Recall, all processes are implicitly assumed to be adapted and continuous.

$$\begin{aligned} \text{Pf: } [X, Y] &= \frac{1}{4} \left([X+Y, X+Y] - [X-Y, X-Y] \right) \\ &= \frac{1}{4} \left([M+N, M+N] - [M-N, M-N] \right) \quad (\because \text{BV part} \\ &\quad \text{does not change } \otimes V) \end{aligned}$$

$$= [M, N] //$$

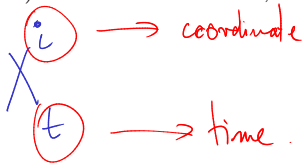
Corollary 9.7. If X is a semi-martingale and B has bounded variation then $[X, B]$ = 0.

$$\begin{aligned} [X, B] &= [\text{mg part of } X, \text{mg part of } B] \\ &= [M, 0] = 0 \end{aligned}$$

Notation.

- d -dimensional vectors: Write $x = (x_1, \dots, x_d) \in \mathbb{R}^d$.
- d -dimensional random vectors: $X = (X_1, \dots, X_d)$, where each X_i is a random variable.
- d -dimensional stochastic processes: $X_t = (X_t^1, \dots, X_t^d)$, where each X_t^i is a stochastic process.
 - ▷ For scalars (or random variables): X^i denotes the i -th power of X .
 - ▷ For vectors (or random random vectors): X^i denotes the i -th coordinate of X .
 - ▷ There is no ambiguity (can't take powers of vectors, or coordinates of scalars)
- Alternate notation used in many books: Use $X(t)$ for the d -dimensional stochastic process, and $X_i(t)$ for the i -th coordinate.
- May sometimes write $X = (X^1, \dots, X^d)$ for random vectors, instead of (X_1, \dots, X_d) .

Notation for d -dim Processes



Remark 9.8 (Chain rule). If \underline{X} is a differentiable function of t , then

$$d(\underline{f}(t, X_t)) = \partial_t f(t, X_t) dt + \sum_{i=1}^d \partial_i f(t, X_t) dX_t^i$$


Remark 9.9 (Notation). $\underline{\partial}_t f = \frac{\partial f}{\partial t}$, $\underline{\partial}_i f = \frac{\partial f}{\partial x_i}$.

$$f = f(\underline{t}, x), \quad x \in \mathbb{R}^d$$

$$\frac{d}{dt} \left(f(\underline{t}, X_t) \right) \stackrel{\text{Chain Rule}}{=} \partial_t f \Big|_{(t, X_t)} \frac{dt}{dt} + \sum_{i=1}^d \partial_i f \Big|_{(t, X_t)} \frac{dX_t^i}{dt}$$

$$\Rightarrow \frac{d}{dt} f(t, X_t) = \partial_t f(t, X_t) + \sum_{i=1}^d \partial_{i0} f(t, X_t) \frac{dX^i}{dt}$$

"Multiply by dt " & get

$$df(t, X_t) = \partial_t f(t, X_t) dt + \sum_{i=1}^d \partial_{i0} f(t, X_t) dX^i$$


Theorem 9.10 (Multi-dimensional Itô formula).

- Let \underline{X} be a d -dimensional Itô process. $\underline{X}_t = (X_t^1, \dots, X_t^d)$.
- Let $f = f(\underline{t}, \underline{x})$ be a function that's defined for $\underline{t} \in \mathbb{R}$, $\underline{x} \in \mathbb{R}^d$.
- Suppose $f \in C^{1,2}$. That is:
 - ▷ f is once differentiable in t
 - ▷ f is twice in each coordinate x_i (includes $\partial_i \partial_j f$)
 - ▷ All the above partial derivatives are continuous. Then:

$$\boxed{d(f(\underline{t}, \underline{X}_t))} = \underline{\partial_t f(t, X_t)} dt + \sum_{i=1}^d \underline{\partial_i f(t, X_t)} d\underline{X}_t^i + \frac{1}{2} \sum_{i,j} \underline{\partial_i \partial_j f(t, X_t)} \underline{d[X^i, X^j]_t}$$

chain Rule

extra!

↑
Joint Q.V.

Remark 9.11 (Integral form of Itô's formula).

$$f(T, X_T) - f(0, X_0) = \int_0^T \underline{\partial_t f}(t, X_t) dt + \sum_{i=1}^d \int_0^T \partial_i f(t, X_t) dX_t^i + \frac{1}{2} \sum_{i,j} \int_0^T \partial_i \partial_j f(t, X_t) d[X^i, X^j]_t$$

Remark 9.12. As with the 1D Itô, will drop the arguments (t, X_t) . Remember they are there.

$$\int_0^T a_t dt + \sum_{i=1}^d \int_0^T a_i dX_t^i + \frac{1}{2} \sum_{i,j} \int_0^T a_i a_j d[X^i, X^j]_t$$

Intuition behind Theorem 9.10.



$$\mathcal{P} = \{0 = t_0 < t_1 \dots t_n = T\}$$

$$f(T, X_T) - f(0, X_0) = \sum_{k=1}^n f(t_{k+1}, X_{t_{k+1}}) - f(t_k, X_{t_k})$$

$$\begin{aligned} \text{Taylor} \quad & \sum_{k=0}^{n-1} \frac{\partial f}{\partial t} (t_{k+1} - t_k) + \sum_{k=0}^{n-1} \sum_{i=1}^d \frac{\partial f}{\partial x^i} (X_{t_{k+1}}^i - X_{t_k}^i) \\ & \xrightarrow{\text{IPI} \Rightarrow} \int_0^T \frac{\partial f}{\partial t} dt \quad \xrightarrow{\text{IPI} \Rightarrow} \int_0^T \frac{\partial f}{\partial x^i} dX^i \end{aligned}$$

$$+ \sum_{k=0}^{n-1} \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d \sigma_i \sigma_j \left(X_{t_{k+1}}^i - X_{t_k}^i \right) \left(X_{t_{k+1}}^j - X_{t_k}^j \right)$$

inc of Joint O.V.

IPF \rightarrow

$$\frac{1}{2} \int_0^T \sum_{i=1}^d \sum_{j=1}^d \sigma_i \sigma_j d[X^i, X^j]_t$$

To use the d -dimensional Itô formula, we need to compute joint quadratic variations.

Proposition 9.13. Let M, N be continuous martingales, with $EM_t^2 < \infty$ and $EN_t^2 < \infty$.

- (1) $MN - [M, N]$ is also a continuous martingale.
- (2) Conversely if $MN - B$ is a continuous martingale for some continuous adapted, bounded variation process B with $B_0 = 0$, then $B = [M, N]$.

Proof. ①
$$d(\underline{MN} - [M, N]) = M dN + N dM + \cancel{d[M, N]} - \cancel{d[M, N]}$$
$$= \underbrace{M dN}_{Mg} + \underbrace{N dM}_{Mg}$$

(Recall: If M is a mg then $M^2 - [M, M]$ is also a mg)

Proposition 9.14. (1) (Symmetry) $[X, Y] = [Y, X]$

(2) (Bi-linearity) If $\alpha \in \mathbb{R}$, X, Y, Z are semi-martingales, $[X, Y + \alpha Z] = [X, Y] + \alpha[X, Z]$.

Proof.

$$\begin{aligned} \text{Just Q.V. } [X, Y + \alpha Z]_T &= \lim_{\|P\| \rightarrow 0} \sum (\Delta_i X) (\Delta_i (Y + \alpha Z)) \\ &= \lim_{\|P\| \rightarrow 0} \sum \Delta_i X \Delta_i Y + \alpha \sum (\Delta_i X) (\Delta_i Z) \\ &= [X, Y]_T + \alpha [X, Z]_T // \end{aligned}$$

Proposition 9.15. Let $\underline{M}, \underline{N}$ be two martingales, σ, τ two adapted processes.

• Let $\underline{X}_t = \int_0^t \underline{\sigma}_s d\underline{M}_s$ and $\underline{Y}_t = \int_0^t \underline{\tau}_s d\underline{N}_s$.

• Then $\underline{[X, Y]}_t = \int_0^t \underline{\sigma}_s \underline{\tau}_s d\underline{[M, N]}_s$.

$$[X, Y]_t = \int_0^t \sigma_s \tau_s d[M, N]_s$$

Remark 9.16. Alternately, if $dX_t = \sigma_t dM_t$ and $dY_t = \tau_t dN_t$, then $d[X, Y]_t = \sigma_t \tau_t d[M, N]_t$.

Intuition.

Recall $[X, X]_t = \int_0^t \sigma_s^2 d[M, M]_s$

$$\rightarrow X_T = \lim_{\|P\| \rightarrow 0} \sum \sigma_{t_i} \Delta_i M \quad \left. \vphantom{\sum} \right\} \Delta_i X = X_{t_{i+1}} - X_{t_i}$$

$$Y_T = \lim_{\|P\| \rightarrow 0} \sum \tau_{t_i} \Delta_i N \quad) \quad \approx \sum \sigma_{t_i} (M_{t_{i+1}} - M_{t_i})$$

$$\Delta_i Y \approx \tau_{t_i} (N_{t_{i+1}} - N_{t_i})$$

$$[X, Y] \approx \sum (\Delta_i X) (\Delta_i Y)$$

$$= \sum (\sigma_i \Delta_i M) (\tau_i \Delta_i N) \longrightarrow \int_0^T \sigma_t \tau_t d[M, N]_t$$

Proposition 9.17. If M, N are continuous martingales, $\mathbf{E}M_t^2 < \infty$, $\mathbf{E}N_t^2 < \infty$ and M, N are independent, then $[M, N] = 0$.

Remark 9.18 (Warning). Independence implies $\mathbf{E}(M_t N_t) = \mathbf{E}M_t \mathbf{E}N_t$. But it *does not* imply $\mathbf{E}_s(M_t N_t) = \mathbf{E}_s M_t \mathbf{E}_s N_t$. So you can't use this to show MN is a martingale, and hence conclude $[M, N] = 0$.

Correct proof.

WRONG
(Not fixable!).

$$\mathbf{E}_s(M_t N_t) \stackrel{\text{indep}}{=} \mathbf{E}_s M_t \mathbf{E}_s N_t \quad \leftarrow \text{FALSE}$$

$$= M_s N_s$$

$\Rightarrow M_t N_t$ is a mg $\Rightarrow [M, N] = 0$

Comment P_f: Claim $E [M, N]_{\downarrow T}^2 = 0$

$$E [M, N]_{\downarrow T}^2 \approx E \left(\sum (\Delta_i M) (\Delta_i N) \right)^2$$

$$= E \left[\sum_{i \neq j} (\Delta_i M) (\Delta_i N) (\Delta_j M) (\Delta_j N) \right]$$
$$= E \sum_{i=0}^{n-1} (\Delta_i M)^2 (\Delta_i N)^2 + 2E \sum_{j=1}^{n-1} \sum_{i=0}^{j-1} (\Delta_i M) (\Delta_j M) (\Delta_i N) (\Delta_j N)$$

$$= \parallel + 2 \sum_{j=1}^{n-1} \sum_{i=0}^{j-1} E(\Delta_i M \Delta_j M) E((\Delta_i N) \Delta_j N)$$

$$= \parallel + 2 \sum \sum E E_{t_{i+1}} (\Delta_i M \Delta_j M) \cdot E()$$

$$= E \sum_{i=0}^{n-1} (\Delta_i M)^2 (\Delta_i N)^2 +$$

$$\begin{aligned} & \because E_{t_{i+1}} (\Delta_j M) \\ &= E_{t_{i+1}} (M_{t_{j+1}} - M_{t_j}) \\ &= M_{t_{i+1}} - M_{t_{i+1}} = 0 \end{aligned}$$

$$= \sum_{i=0}^{n-1} \underbrace{E(\Delta_i M)^2}_{\text{wavy line}} E(\Delta_i N)^2.$$

$$\leq \underbrace{\max_i E(\Delta_i M)^2}_{\text{wavy line}} \cdot \underbrace{\sum_i E(\Delta_i N)^2}_{\text{wavy line}}$$

↓ (since Mis ds!)
0

↓
 $E[N, N]_T < \infty$

= 0

Remark 9.19. $[M, N] = 0$ does not imply M, N are independent. For example:

- Let $M_t = \int_0^t \mathbf{1}_{W_s < 0} dW_s$
- Let $N_t = \int_0^t \mathbf{1}_{W_s \geq 0} dW_s$

$$\left. \begin{aligned} M_t &= \int_0^t \mathbf{1}_{\{W_s < 0\}} dW_s \\ N_t &= \int_0^t \mathbf{1}_{\{W_s \geq 0\}} dW_s \end{aligned} \right\} d[M, N] = \mathbf{1}_{\{W_s < 0\}} \mathbf{1}_{\{W_s \geq 0\}} ds = 0 ds$$

But $M_t + N_t = \int_0^t \mathbf{1} dW_s = W_t \leftarrow \Rightarrow M \& N \text{ are Not ind.}$

Definition 9.20 (d -dimensional Brownian motion). We say a d -dimensional process $\underline{W} = (W^1, \dots, W^d)$ is a Brownian motion if:

- (1) Each coordinate \underline{W}^i is a standard 1-dimensional Brownian motion.
- (2) For $\underline{i} \neq \underline{j}$, the processes \underline{W}^i and \underline{W}^j are independent.

Remark 9.21. If W is a d -dimensional Brownian motion then

$$d[\underline{W}^i, \underline{W}^j]_t = \begin{cases} \underline{dt} & i = j, \\ \underline{0 dt} & i \neq j. \end{cases}$$

Theorem 9.22 (Lévy). Let M be a d -dimensional process such that:

(1) M is a continuous martingale.

(2) The joint quadratic variation satisfies: $d[W^i, W^j]_t = \begin{cases} dt & i = j, \\ 0 dt & i \neq j. \end{cases}$

Then M is a d -dimensional Brownian motion.

Proof. Find $\mathbf{E}_s e^{\lambda M_t^i + \mu M_t^j}$ using Itô's formula, similar to Problem 7.5. □

Example 9.23. Let $f \in C^{1,2}$, W be a d -dimensional Brownian motion, and set $X_t = f(t, W_t)$. Find the Itô decomposition of X .

Question 9.24. *Let W be a 2-dimensional Brownian motion. Let $X_t = \ln(|W_t|^2) = \ln((W_t^1)^2 + (W_t^2)^2)$. Is X a martingale?*

10. Risk Neutral Pricing

Goal.

- Consider a market with a bank and one stock.
- The interest rate R_t is some adapted process.
- The stock price satisfies $dS_t = \alpha_t S_t dt + \sigma_t S_t dW_t$. (Here α, σ are adapted processes).
- Find the risk neutral measure and use it to price securities.

Definition 10.1. Let $D_t = \exp(-\int_0^t R_s ds)$ be the discount factor.

Remark 10.2. Note $\partial_t D = -R_t D_t$.

Remark 10.3. D_t dollars in the bank at time 0 becomes \$1 in the bank at time t .

Theorem 10.4. *The (unique) risk neutral measure is given by $d\tilde{\mathbf{P}} = Z_T d\mathbf{P}$, where*

$$Z_T = \exp\left(-\int_0^T \theta_t dW_t - \frac{1}{2} \int_0^T \theta_t^2 dt\right), \quad \theta_t = \frac{\alpha_t - R_t}{\sigma_t}.$$

Theorem 10.5. *Any security can be replicated. If a security pays V_T at time T , then the arbitrage free price at time t is*

$$V_t = \frac{1}{D_t} \tilde{\mathbf{E}}_t(D_T V_T).$$

Remark 10.6. We will explain the notation $d\tilde{\mathbf{P}} = Z_T d\mathbf{P}$ and prove both the above theorems later.

Definition 10.7. We say $\tilde{\mathbf{P}}$ is a risk neutral measure if:

- (1) $\tilde{\mathbf{P}}$ is equivalent to \mathbf{P} (i.e. $\tilde{\mathbf{P}}(A) = 0$ if and only if $\mathbf{P}(A) = 0$)
- (2) $D_t S_t$ is a $\tilde{\mathbf{P}}$ martingale.

Remark 10.8. As before, if $\tilde{\mathbf{P}}$ is a new measure, we use $\tilde{\mathbf{E}}$ to denote expectations with respect to $\tilde{\mathbf{P}}$ and $\tilde{\mathbf{E}}_t$ to denote conditional expectations.

Example 10.9. Fix $T > 0$. Let Z_T be a \mathcal{F}_T -measurable random variable.

- Assume $Z_T > 0$ and $\mathbf{E}Z_T = 1$.
- Define $\tilde{\mathbf{P}}(A) = \mathbf{E}(Z_T \mathbf{1}_A) = \int_A Z_T d\mathbf{P}$.
- Can check $\tilde{\mathbf{E}}X = \mathbf{E}(Z_T X)$. That is $\int_{\Omega} X d\tilde{\mathbf{P}} = \int_{\Omega} X Z_T d\mathbf{P}$.
- Notation: Write $d\tilde{\mathbf{P}} = Z_T d\mathbf{P}$.

Lemma 10.10. Let $Z_t = \mathbf{E}_t Z_T$. If X_t is \mathcal{F}_t -measurable, then $\tilde{\mathbf{E}}_s X = \frac{1}{Z_s} \tilde{\mathbf{E}}_s(Z_t X_t)$.

Proof. You will see this in the proof of the Girsanov theorem in part 2 of this course. □

Theorem 10.11 (Cameron, Martin, Girsanov). *Fix $T > 0$, and define:*

- $b_t = (b_t^1, \dots, b_t^d)$ a d -dimensional adapted process.
- W a d -dimensional Brownian motion.
- $\tilde{W}_t = W_t + \int_0^t b_s ds$ (i.e. $d\tilde{W}_t = b_t dt + dW_t$).
- $d\tilde{\mathbf{P}} = Z_T d\mathbf{P}$, where

$$Z_t = \exp\left(-\int_0^t b_s \cdot dW_s - \frac{1}{2} \int_0^t |b_s|^2 ds\right).$$

If Z is a martingale, then $\tilde{\mathbf{P}}$ is an equivalent measure under which \tilde{W} is a Brownian motion up to time T .