

- No ~~anti~~ arbitrage.

$$X_0 = 0, X_N \geq 0 \Rightarrow \mathbb{P}(X_N = 0) = 1$$

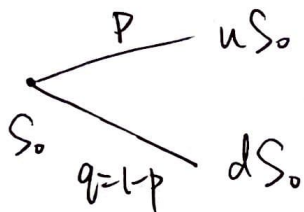
- Arbitrage free price (AFP)

Security V_N

Starts X_0 , $X_N = V_N$

$$\Rightarrow \text{AFP} = X_0.$$

- Binomial model



no arbitrage $\Leftrightarrow d < 1+r < u.$

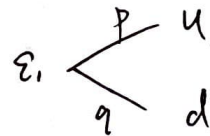
e.g. $1+r = d.$

$t=0$ long 1 Stock

$t=1$

short S_0 cash

$$S_1 = \varepsilon_1 S_0$$

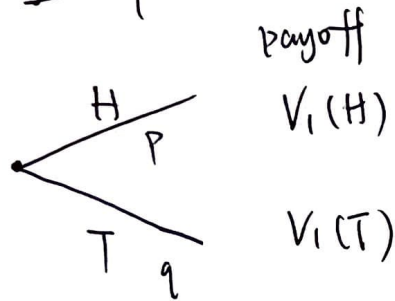


$$X_0 = 1 \cdot S_0 - S_0 = 0$$

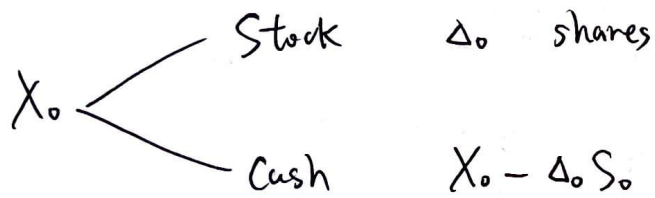
$$X_1 = S_1 - (1+r)S_0 = (\varepsilon_1 - \cancel{1+r} d) S_0 \geq 0$$

$$\mathbb{P}(X_1 > 0) = \mathbb{P}(\varepsilon_1 = u) = p > 0.$$

• 1-period



$$\tilde{p} = \frac{1+r-d}{u-d}, \quad \tilde{q} = \frac{u-(1+r)}{u-d}$$



$$X_1 = \Delta_0 S_1 + (1+r)(X_0 - \Delta_0 S_0) = V_1$$

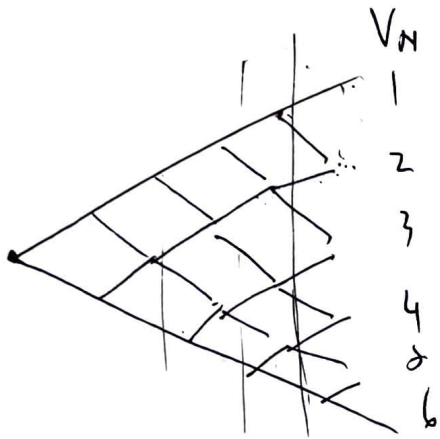
$$\left. \begin{array}{l} V_1(H) = \Delta_0 u S_0 + (1+r)(X_0 - \Delta_0 S_0) \\ V_1(T) = \Delta_0 d S_0 + (1+r)(X_0 - \Delta_0 S_0) \end{array} \right\}$$

$$\left. \begin{array}{l} V_1(H) = \Delta_0 u S_0 + (1+r)(X_0 - \Delta_0 S_0) \\ V_1(T) = \Delta_0 d S_0 + (1+r)(X_0 - \Delta_0 S_0) \end{array} \right\}$$

$$V_0 = \frac{1}{1+r} \left(\tilde{p} V_1(H) + \tilde{q} V_1(T) \right)$$
$$= \frac{1}{1+r} \tilde{\mathbb{E}} V_1$$

$$\Delta_0 = \frac{V_1(H) - V_1(T)}{(u-d)S_0}$$

- Multi-Period



$$\tilde{p} = \frac{1+r-d}{u-d}, \quad \tilde{q} = \frac{u-(1+r)}{u-d}$$

$$V_n = \frac{1}{(1+r)^{N-n}} \mathbb{E}_n^{\tilde{p}, \tilde{q}} V_N$$

- Discrete probability space. (finite PS)

$$\Omega = \{ \omega = (w_1, \dots, w_N) \mid w_i = H \text{ or } T \}$$

sample point $\omega = (w_1, \dots, w_N) \in \Omega$.

- Probability mass function (PMF)

$$P: \Omega \rightarrow [0, 1] \text{ s.t. } \sum_{\omega \in \Omega} P(\omega) = 1.$$

e.g. $p(\omega) = p^{H(\omega)} q^{T(\omega)}$, $H(\omega) = \#$ of heads in (w_1, \dots, w_N) .

- Event. $A \subseteq \Omega$. $P(A) = \sum_{\omega \in A} P(\omega)$. $P(\{w_i = H\}) = p$

• Random variable

$$X: \Omega \rightarrow \mathbb{R}$$

$$\text{Range}(X) = \{X(\omega) \mid \omega \in \Omega\}.$$

$$\sigma(X) = \{ \{X \in A\} \mid A \subseteq \mathbb{R} \} = \{ \{X \in A\} \mid A \subseteq \text{Range}(X) \}.$$

$$\{X \in A\} = \{ \omega \in \Omega \mid X(\omega) \in A \}$$

$$\{ \omega \in \Omega \mid X(\omega) \notin \text{Range}(X) \} = \emptyset.$$

$$X(HH) = X(HT) = 1, \quad X(TH) = X(TT) = 0.$$

\emptyset, Ω always in $\sigma(X)$.
 ($A = \emptyset$) ($A = \mathbb{R}$
 or $A = \text{Range}(X)$)

$$\begin{aligned} \sigma(X) &= \{ \emptyset, \Omega, \{X=1\}, \{X=0\} \} \\ &= \{ \emptyset, \Omega, \{HH, HT\}, \{TH, TT\} \} \end{aligned}$$

~~$\{ \emptyset, \Omega \}$~~

$$\{ \emptyset, \{0\}, \{1\}, \{0, 1\} \} \quad X$$

$$\sigma(X) \subseteq \mathcal{P}(\Omega) = \{ \text{all subsets of } \Omega \}.$$

• Expectation

$$\mathbb{E}X = \sum_{\omega \in \Omega} X(\omega) P(\omega) = \sum_{x \in \text{Range}(X)} x \underline{P(X=x)}$$

2^N

$$\{X=x\} = \{ \omega \in \Omega \mid X(\omega) = x \}.$$

• Independence of Events.

A_1, \dots, A_n events,

A_1, \dots, A_n are independent, if $\forall 1 \leq i_1 < i_2 < \dots < i_k \leq n$,

$$P(A_{i_1} \cap \dots \cap A_{i_k}) = P(A_{i_1}) \cdot \dots \cdot P(A_{i_k}).$$

• $P(A_i \cap A_j) = P(A_i)P(A_j), \forall i < j$ ~~X~~

• $P(A_1 \cap \dots \cap A_n) = P(A_1) \cdot \dots \cdot P(A_n)$ ~~X~~ . $P(A_1^c \cap A_2 \cap \dots \cap A_n)$

$$= P(A_1^c) P(A_2) \cdot \dots \cdot P(A_n)$$

Ex 1. $\Omega = \{1, 2, 3, 4\}$, $P(\{\omega\}) = \frac{1}{4}$, $\omega \in \Omega$.

$$A_1 = \{1, 2\}, A_2 = \{1, 3\}, A_3 = \{1, 4\}.$$

$$P(A_i) = \frac{1}{2}, \forall i$$

$$P(A_i \cap A_j) = P(\{1\}) = \frac{1}{4} \quad \forall i \neq j$$

$$= P(A_i) P(A_j)$$

$$P(A_1 \cap A_2 \cap A_3) = P(\{1\}) = \frac{1}{4} \neq \frac{1}{8} = P(A_1) P(A_2) P(A_3)$$

Ex 2. $\Omega = \{1, 2, 3, 4, 5, 6, 7, 8\}$, $P(\{w\}) = \frac{1}{8}$.

$A = B = \{1, 2, 3, 4\}$, $C = \{1, 5, 6, 7\}$. $P(A) = P(B) = P(C) = \frac{1}{2}$.

$P(A \cap B \cap C) = P(\{1\}) = \frac{1}{8} = P(A)P(B)P(C)$

$P(A \cap B) = P(A) = \frac{1}{2} \neq \frac{1}{4} = P(A)P(B)$

~~###~~

• X_1, \dots, X_n R.V. are independent if

$P(X_1 = x_1, \dots, X_n = x_n) = P(X_1 = x_1) \cdot \dots \cdot P(X_n = x_n)$, $\forall x_i \in \text{Range}(X_i)$.

~~Q~~ Fix x_2, x_3, \dots, x_n ,

$$\sum_{x_1 \in \text{Range}(X_1)} P(X_1 = x_1, X_2 = x_2, \dots, X_n = x_n) = \sum_{x_1 \in \text{Range}(X_1)} \underbrace{P(X_1 = x_1)}_{\text{"}} \underbrace{\left(P(X_2 = x_2) \dots P(X_n = x_n) \right)}_{\text{"}}$$

$= P\left(\bigcup_{x_1 \in \text{Range}(X_1)} \{X_1 = x_1, X_2 = x_2, \dots, X_n = x_n\} \right)$

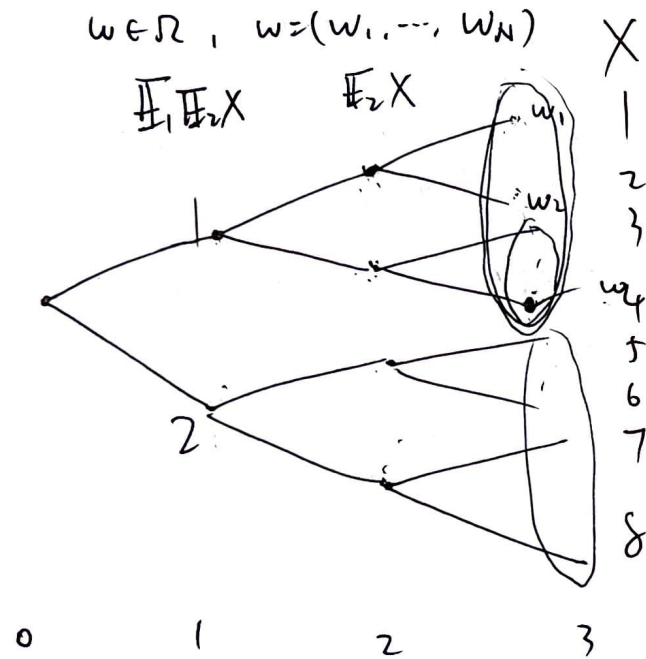
$= P(X_2 = x_2, \dots, X_n = x_n) \quad \underline{\underline{P(X_2 = x_2), \dots, P(X_n = x_n)}}$

• Filtration

$$\{\emptyset, \Omega\} = \mathcal{F}_0 \subseteq \mathcal{F}_1 \subseteq \dots \subseteq \mathcal{F}_N = \mathcal{P}(\Omega).$$

X_n is \mathcal{F}_n -measurable if X_n only depends on w_1, \dots, w_n .

$$X_n: \Omega \rightarrow \mathbb{R}.$$



$$\frac{\mathbb{P}(\{w_1, 3\})}{\mathbb{P}(\{w_1, w_2\})} \quad \frac{\mathbb{P}(\{w_2, 3\})}{\mathbb{P}(\{w_1, w_2\})}$$

e.g. X is \mathcal{F}_1 -measurable

$$\pi_1(w) \quad \pi_2(w)$$

$$\mathbb{E}_1 X$$

X is independent of \mathcal{F}_n if

$$\mathbb{P}(\{X \in A\} \cap B) = \mathbb{P}(X \in A) \mathbb{P}(B), \quad A \subseteq \mathbb{R}, B \in \mathcal{F}_n.$$

- Conditional expectation

$$\mathbb{E}_n X(\omega) = \mathbb{E}(X | \mathcal{F}_n)(\omega) = \sum_{x \in \text{Range}(X)} x \mathbb{P}(X=x | \mathbb{T}_n(\omega)), \quad \begin{aligned} &= \frac{\mathbb{P}(\{X=x\} \cap \mathbb{T}_n(\omega))}{\mathbb{P}(\mathbb{T}_n(\omega))} \\ &\text{it is a R.V.} \end{aligned}$$

where $\mathbb{T}_n(\omega) = \{ \omega' \in \Omega \mid \omega'_1 = \omega_1, \omega'_2 = \omega_2, \dots, \omega'_n = \omega_n \}$.

- $\mathbb{E}_n X$ is \mathcal{F}_n -measurable

$$\mathbb{E}_n(\alpha X + \beta Y) = \alpha \mathbb{E}_n X + \beta \mathbb{E}_n Y$$

$$m \leq n, \quad \mathbb{E}_m(X) = \mathbb{E}_m(\mathbb{E}_n X) = \mathbb{E}_n(\mathbb{E}_m X)$$

$$X \text{ is } \mathcal{F}_n\text{-measurable, } \mathbb{E}_n(XY) = X \mathbb{E}_n Y \quad (Y=1 \Rightarrow \mathbb{E}_n X = X)$$

$$\mathbb{E} \mathbb{E}_n X = \mathbb{E} X$$

- X_0, X_1, \dots, X_N is a stochastic process.

Adapted: $\forall n, X_n$ is \mathcal{F}_n -measurable.

Martingale: adapted + $\mathbb{E}_n X_{n+1} = X_n, \forall n \geq 0$.

Ex. symmetric random walk, $X_n = \xi_1 + \dots + \xi_n, \xi_i$ iid. $\mathbb{P}(\xi_i = \pm 1) = \frac{1}{2}$.

- Brownian motion W_t $\Omega = C([0, \infty))$
 W_t is a continuous process (fix $\omega \in \Omega$, $W_t(\omega)$ as function of t is continuous on $[0, \infty)$).

s.t. (1) $W_0 = 0$

(2) $\forall 0 \leq s < t$, $W_t - W_s \sim N(0, t-s)$ ($\text{Var}(W_t - W_s) = t-s \neq (t-s)^2$)

(3) $W_t - W_s$ is independent of \mathcal{F}_s ($\mathbb{P}(\{W_t - W_s \in A\} \cap B) = \mathbb{P}(W_t - W_s \in A) \mathbb{P}(B)$).

$\forall 0 \leq s < t$

- Sample space $\Omega = C([0, \infty))$. not finite, not even countable.

$\Omega \neq \{\omega_1, \omega_2, \dots, \omega_n, \dots\}$

$\mathcal{G} \subseteq \mathcal{P}(\Omega)$ ($1^\circ \emptyset \in \mathcal{G}$ $2^\circ A \in \mathcal{G}$ implies $A^c \in \mathcal{G}$ $3^\circ A_1, A_2, \dots, A_n, \dots \in \mathcal{G}$ implies $\bigcup_{n=1}^{\infty} A_n \in \mathcal{G}$)

Define $\mathbb{P} : \mathcal{G} \rightarrow [0, 1]$ ($1^\circ \mathbb{P}(\emptyset) = 0$, 2° if $A \cap B = \emptyset$, $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B)$)

3° if $A_1 \subseteq A_2 \subseteq \dots \subseteq A_n \subseteq \dots$, $\mathbb{P}(\bigcup_{n=1}^{\infty} A_n) = \lim_{n \rightarrow \infty} \mathbb{P}(A_n)$

X is a R.V. $X : \Omega \rightarrow \mathbb{R}$

$\mathbb{E}X = \int_{\Omega} X(\omega) d\mathbb{P}(\omega)$

in general, no formula. (if X has a pdf p
 $\mathbb{E}X = \int_{\mathbb{R}} x p(x) dx$)

$X = \mathbb{1}_A$ $\mathbb{1}_A(\omega) = \begin{cases} 1, & \omega \in A \\ 0, & \omega \notin A. \end{cases}$

$\mathbb{E}\mathbb{1}_A = \int_{\Omega} \mathbb{1}_A(\omega) d\mathbb{P}(\omega) = \mathbb{P}(A)$. $\mathbb{E}\left(\sum_{i=1}^n a_i \mathbb{1}_{A_i}\right) = \sum_{i=1}^n a_i \mathbb{P}(A_i)$

• $\mathcal{F}_t = \{ \text{events that can be described by } (B_s)_{0 \leq s \leq t} \}$.

$$\{ W_1 \in (1, 2) \} \in \mathcal{F}_2.$$

$$\{ W_3 > 0 \} \notin \mathcal{F}_2$$

• martingale: X_t

1° X is adapted

2° $\mathbb{E}_s X_t = X_s, \forall 0 \leq s < t$

$$\left(\mathbb{E}_t X_{t+1} = X_t, \forall t \geq 0 \quad X \right)$$

$$\underline{\mathbb{E}_1 X_{3/2} = ?}$$

• W is a BM

$\forall 0 < t_1 < t_2 < \dots < t_n$

$(W_{t_1}, W_{t_2}, \dots, W_{t_n})$ is joint normal

$$\mathbb{E}(W_{t_i} W_{t_j}) = t_i \wedge t_j = \min \{ t_i, t_j \}.$$

$$\begin{pmatrix} W_{t_1} \\ W_{t_2} \\ \vdots \\ W_{t_n} \end{pmatrix} = \begin{pmatrix} 1 & 0 & \dots & 0 \\ 1 & 1 & 0 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 1 & 1 & \dots & 1 \end{pmatrix} \begin{pmatrix} W_{t_1} \\ W_{t_2} - W_{t_1} \\ \vdots \\ W_{t_n} - W_{t_{n-1}} \end{pmatrix}$$

joint normal

Ex. Not every two normal r.v.'s form a joint normal distribution.

$X \sim N(0, 1), \varepsilon \begin{cases} 1 \\ -1 \end{cases} \perp X, \varepsilon$ independent, $Y = \varepsilon X$

$$P(Y \leq y) = P(\varepsilon X \leq y) = P(\varepsilon X \leq y, \varepsilon = 1) + P(\varepsilon X \leq y, \varepsilon = -1)$$

$$\begin{aligned}
&= \mathbb{P}(X \leq y, \varepsilon=1) + \mathbb{P}(-X \leq y, \varepsilon=-1) \\
&= \mathbb{P}(X \leq y) \mathbb{P}(\varepsilon=1) + \mathbb{P}(-X \leq y) \mathbb{P}(\varepsilon=-1) \\
&= \frac{1}{2} \left(\mathbb{P}(X \leq y) + \mathbb{P}(-X \leq y) \right) \quad (X \sim N(0,1), -X \sim N(0,1)) \\
&= \frac{1}{2} (\Phi(y) + \Phi(y)) = \Phi(y).
\end{aligned}$$

$$\Rightarrow Y = \varepsilon X \sim N(0,1)$$

(X, Y) is not joint normal.

$$= (X, \varepsilon X)$$

$$\text{Range}(X, \varepsilon X) = \{ (x, y) \in \mathbb{R}^2 \mid x=y \} \cup \{ (x, y) \in \mathbb{R}^2 \mid x=-y \} \neq \mathbb{R}^2.$$

$(X, \varepsilon X)$ not joint normal.

$$\begin{aligned}
\text{Cov}(X, Y) &= \mathbb{E}XY = \mathbb{E}(\varepsilon X^2) \\
&= \mathbb{E}\varepsilon \cdot \mathbb{E}X^2 = 0
\end{aligned}$$

X, Y uncorrelated.

X, Y not independent.

