

(1) Law of Large #s.

X_1, X_2, X_3, \dots (iid).

$$E X_i = \mu$$

Say we play N times

→ Avg winnings: $\frac{1}{N} \sum_{k=1}^N X_k$

→ Expected winnings = μ $\downarrow N \rightarrow \infty$

8.1. **Law of large numbers.** Now consider infinitely many i.i.d. random variables X_1, X_2, \dots .

Theorem 8.1 (Weak law of large numbers). Suppose $\underline{E}X_n = \underline{\mu}$ and $\text{Var } X_n = \underline{\sigma}^2 < \infty$, and let $S_n = \sum_1^n X_k$. Then $\text{Var}(\underline{S}_n/n) \rightarrow 0$, and hence for any $\underline{\varepsilon} > 0$, $\lim_{n \rightarrow \infty} \underline{P}\left(\left|\frac{S_n}{n} - \underline{\mu}\right| > \underline{\varepsilon}\right) = 0$.

Lemma 8.2 (Chebyshev's inequality). For any $\underline{\varepsilon} > 0$, $\underline{P}(X > \underline{\varepsilon}) \leq \frac{1}{\underline{\varepsilon}} \underline{E}|X|$.

total wings after n flights.

Pf: Note $\varepsilon \mathbb{1}_{\{X > \varepsilon\}} \leq |X|$

$$\Rightarrow \varepsilon \underline{E} \mathbb{1}_{\{X > \varepsilon\}} \leq \underline{E}|X|$$

$$\Rightarrow \underline{P}(X > \varepsilon) \leq \frac{1}{\varepsilon} \underline{E}|X|. \quad \text{QED.}$$

Proof of Theorem 8.1

$$E X_n = \mu, \quad E (X_n - \mu)^2 = \sigma^2. \quad X_n \text{'s are iid.}$$

$$S_n = \sum_{k=1}^n X_k$$

① NTS $\text{Var}\left(\frac{S_n}{n}\right) \rightarrow 0.$

Pf: Note $\text{Var}\left(\frac{S_n}{n}\right) = \frac{1}{n^2} \text{Var}(S_n) = \frac{1}{n^2} \sum_{k=1}^n \text{Var}(X_k) = \frac{\sigma^2}{n}.$

$$\Rightarrow \text{Var}\left(\frac{S_n}{n}\right) = \frac{\sigma^2}{n} \xrightarrow{n \rightarrow \infty} 0$$

(\because X & Y are indep
 $\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$)

② NTS $\forall \varepsilon > 0, \quad P\left(\left|\frac{S_n}{n} - \mu\right| > \varepsilon\right) \xrightarrow{n \rightarrow \infty} 0$

$$\text{Pf: } P\left(\left|\frac{S_M}{n} - \mu\right| > \varepsilon\right) = P\left(\left(\frac{S_M}{n} - \mu\right)^2 > \underline{\underline{\varepsilon^2}}\right)$$

$$\stackrel{\text{Chebyshev}}{\leq} \frac{1}{\varepsilon^2} \underbrace{E\left(\frac{S_M}{n} - \mu\right)^2}_{= \text{Var}\left(\frac{S_M}{n}\right)}$$

$$= \frac{1}{\varepsilon^2} \cdot \frac{\sigma^2}{n} \longrightarrow 0$$

QED.

Theorem 8.3 (Strong law of large numbers). Under the same assumptions as Theorem 8.1, $\lim_{n \rightarrow \infty} \frac{S_n}{n} = \mu$ almost surely.

avg winnings after n plays
Expected return

(Pf is harder)

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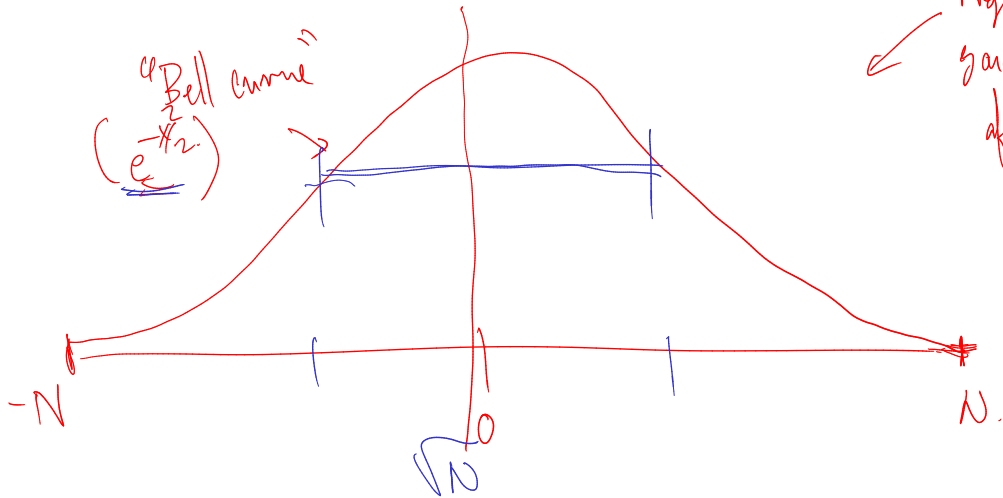
Play a fair slot machine N times.

$X_k \rightarrow$ outcome of k^{th} slot.

(iid)

$$E X_k = 0$$

$$X_k = 1 \quad \text{with prob } \frac{1}{2}$$
$$X_k = -1 \quad \text{with prob } \frac{1}{2}$$



8.2. Central limit theorem.

Theorem 8.4. Let X_n be a sequence of \mathbb{R}^d valued, i.i.d. random variables be such that $\mathbf{E}X_n^i = \underline{\mu}_i$ and $\text{cov}(X_n^i, X_n^j) = \underline{\Sigma}_{i,j}$. Let $S_N = \sum_1^N X_n$. Then $(S_N - N\underline{\mu})/\sqrt{N}$ converges weakly to $\mathcal{N}(\underline{\mu}, \underline{\Sigma})$.

Definition 8.5. We say a sequence of random variables Y_n converges weakly to a random variable Z if $\mathbf{E}f(Y_n) \rightarrow \mathbf{E}f(Z)$ for every bounded continuous function f .

Definition 8.6. Let $\underline{\mu} \in \mathbb{R}^d$, and $\underline{\Sigma}$ be a $d \times d$ covariance matrix (positive semi-definite, symmetric). $((\underline{\Sigma}) \cdot x \geq 0 \forall x \in \mathbb{R}^d)$

(1) $\mathcal{N}(\underline{\mu}, \underline{\Sigma})$ denotes a normally distributed random variable with mean $\underline{\mu}$ and covariance matrix $\underline{\Sigma}$.

(2) When $\underline{\Sigma}$ is invertible, the probability density function of $\mathcal{N}(\underline{\mu}, \underline{\Sigma})$ is $\frac{1}{(2\pi)^{d/2} (\det(\underline{\Sigma}))^{1/2}} \exp\left(-\frac{1}{2}(x - \underline{\mu}) \cdot \underline{\Sigma}^{-1}(x - \underline{\mu})\right)$

(3) When $d = 1$, $\underline{\Sigma} = \sigma^2$ the PDF of $\mathcal{N}(\underline{\mu}, \sigma^2)$ is $\frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(x-\underline{\mu})^2}{2\sigma^2}}$.

(4) When $\underline{\mu} = 0$, $\sigma = 1$, $\mathcal{N}(0, 1)$ is called the standard normal, and its PDF is the Gaussian $G(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$.

Definition 8.7. We say p is the probability density function (PDF) of a d -dimensional random variable X if $\mathbf{P}(X \in A) = \int_A p(x) dx$ for all cubes $A \subseteq \mathbb{R}^d$.

Remark 8.8. Equivalently, p is the PDF of X if $\mathbf{E}f(X) = \int_{\mathbb{R}^d} f(x)p(x) dx$ for every bounded continuous function f .

Remark 8.9. We will prove Theorem 8.4 during the course of the construction of Brownian motion.

S.L.L. : $EX_n = 0$.

$$\frac{S_n}{n} \longrightarrow \mu = 0 \quad \text{a.s.}$$

CLT : $EX_n = 0$:

$$\frac{S_n}{\sqrt{n}} \xrightarrow{\text{weakly}} N(0, \Sigma)$$

Normally dist RV.
mean 0, cov. Σ

$Y_n \rightarrow Z$ weakly if $\underbrace{E f(Y_n)} \rightarrow E f(Z)$ \forall odd ds
fns f .

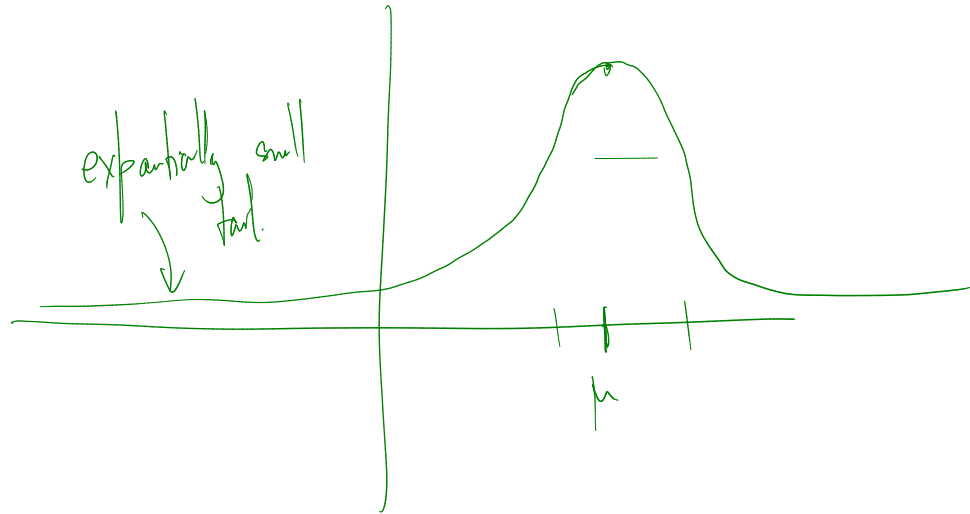
\Leftrightarrow for cts RVS,

$$\lim_{n \rightarrow \infty} P(Y_n \in A) = P(Z \in A)$$

$(\forall A \subseteq \mathbb{R})$
intervals

↙

$$G_{\mu, \sigma^2}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad (\text{Gaussian})$$



✓

X is a \mathbb{R} valued RV.

f is the PDF of X if $\forall a, b \in \mathbb{R}, P(X \in (a, b)) = \int_a^b f(x) dx$

$\Leftrightarrow \forall$ hold etc for f , $E f(X) = \int_{\mathbb{R}} f(x) f(x) dx$

(Note: For disc RV's, f is the PMF.

Y is a RV.

$$E f(Y) = \sum f(y_i) P(Y = y_i)$$

(Range $Y = \{y_1, \dots, y_n\}$)

8.3. Brownian motion.

- Suppose now X_1, X_2, \dots are i.i.d. \mathbb{R} valued random variables.
- Use $\tilde{\mathbf{P}}$ to denote the probability measure, and $\tilde{\mathbf{E}}, \tilde{\mathbf{E}}_n$ to denote the associated expectation / conditional expectation.
- Assume $\tilde{\mathbf{E}}X_n = 0$, and $\tilde{\mathbf{E}}X_n^2 = 1$.

Theorem 8.10. *Let $W_n^N = \frac{1}{\sqrt{N}}S_n = \frac{1}{\sqrt{N}}\sum_1^n X_k$. Then $\lim_{N \rightarrow \infty} W_{[Nt]}^N$ exists almost surely.*

Theorem 8.11. (1) *The function $t \mapsto W_t$ is continuous almost surely, and $W_0 = 0$.*

(2) *If $0 = t_0 < t_1 < \dots < t_n$, then $W_{t_1} - W_{t_0}, W_{t_2} - W_{t_1}, \dots, W_{t_n} - W_{t_{n-1}}$ are independent and $W_{t_i} - W_{t_{i-1}} \sim \mathcal{N}(0, t_i - t_{i-1})$.*

Remark 8.12. Typically one changes the probability space to ensure the function $t \mapsto W_t$ is continuous *surely*.

Definition 8.13. The process W above is called a standard (one dimensional) Brownian motion.