

Today:

- Multi Dimensional Ito
- Change of Measure.
- Girsanov Theorem
- RN R.

Finish off with some
exam tips

Defn:

$$f(t, x_1, \dots, x_n) = f(t, x) \in C^{1,2}$$

then if $X \rightsquigarrow X = (X^1, X^2, \dots, X^n)$ is an Ito process

$$df(t, X_t) = \partial_t f(t, X_t) + \sum_{i=1}^n \partial_i f(t, X_t) dX_i$$

$$+ \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n \partial_{ij} f(t, X_t) d[X_i, X_j]_t.$$

Ex

$$\begin{cases} dX_t = \mu_1 X_t dt + \sigma_1 X_t dW_{1,t} \\ dY_t = \mu_2 Y_t dt + \sigma_2 Y_t dW_{2,t} \end{cases}$$

$$d[W_1, W_2]_t = pdt \quad ; \quad -1 \leq p \leq 1.$$

Let $Z_+ = \frac{X_+}{Y_+}$ (a) find dZ_+ .

$$f(t, X, Y) = \frac{X}{Y}.$$

$$f_x = \frac{1}{Y}, \quad f_{xx} = 0.$$

$$f_y = -\frac{X}{Y^2}, \quad f_{yy} = \frac{2X}{Y^3}.$$

$$f_{xy} = f_{yx} = -\frac{1}{Y^2}.$$

$$\text{by Itô: } dZ_+ = \frac{dX_+}{Y_+} - \frac{X_+}{Y_+} \frac{dY_+}{Y_+} + \frac{X_+}{Y_+} \frac{d[X, Y]_+}{Y_+^2} \\ - \frac{d[X, Y]_+}{Y^2}.$$

$$= \frac{X_+}{Y_+} (\mu_1 dt + \sigma_1 dW_{1,+}) - \frac{X_+}{Y_+} (\mu_2 dt + \sigma_2 dW_{2,+}).$$

$$+ \frac{X_t}{Y_t} \sigma_2^2 dt = \frac{\sigma_1 \sigma_2 X_t + Y_t p dt}{Y_t^2}$$

$$= Z_t [M_1 - M_2 + \sigma_2^2 - \sigma_1 \sigma_2 p)dt + \sigma_1 dW_{1,t} + \sigma_2 dW_{2,t}]$$

(6) Compute QV of $Z \cdot [dZ, d\epsilon]_+$

$$[dZ, d\epsilon]_+ = [Z_t ((\dots)dt + \sigma_1 dW_{1,t} + \sigma_2 dW_{2,t}), Z_t (\dots)dt + \sigma_1 dW_{1,t} + \sigma_2 dW_{2,t}]$$

symmetry

$$= Z_t^2 [\sigma_1 dW_{1,t} - \sigma_2 dW_{2,t}, \sigma_1 dW_{1,t} - \sigma_2 dW_{2,t}]$$

$$= Z_t^2 ([\sigma_1 dW_{1,t}, \sigma_1 dW_{1,t} - \sigma_2 dW_{2,t}] - [\sigma_2 dW_{2,t}, \sigma_1 dW_{1,t} - \sigma_2 dW_{2,t}])$$

$$= Z_t^2 ([\sigma_1 dW_{1,t}, \sigma_1 dW_{1,t}] - [\sigma_1 dW_{1,t}, \sigma_2 dW_{2,t}] - [\sigma_2 dW_{2,t}, \sigma_1 dW_{1,t}] + [\sigma_2 dW_{2,t}, \sigma_2 dW_{2,t}])$$

$$= \sum_{i=1}^d (\sigma_i^2 dt - 2\sigma_i \sigma_2 pdt + \sigma_2^2 dt)$$

σ_i from the Merton soln.

↳ called instantaneous volatility
w.r.t X and Y .

$$\text{If } dX_t = b_t dt + \sum_{i=1}^d \sigma_i(t) dW_i(t).$$

$$d[X, X]_t = d[b_t dt + \sum_{i=1}^d \sigma_i(t) dW_i(t), b_t dt + \sum_{i=1}^d \sigma_i(t) dW_i(t)]$$

$$= d \left[\sum_{i=1}^d \sigma_i(t) dW_i, \sum_{j=1}^d \sigma_j(t) dW_j \right]$$

$$= \sum_{i=1}^d \sum_{j=1}^d [\sigma_i(t) dW_i, \sigma_j(t) dW_j]$$

$$= \sum_{i=1}^d \sum_{j=1}^d [\sigma_i(t) dW_i, \sigma_j(t) dW_j]$$

$$= \sum_{i=1}^d \sum_{j=1}^d \sigma_i(t) \sigma_j(t) d[w_{ij+}, w_{ij+}].$$

still holds if $x_+ = b_+ dx + \sum_{i=1}^d \sigma_i(t) \cancel{dx_i}$.

\hookrightarrow To prove

$$d[x_+, x_+] = \sum_{i=1}^d \sum_{j=1}^d \cancel{\sigma_i(t) \sigma_j(t)} d[x_{ij+}, x_{ij+}]$$

Change of Measure:

P prob measure. Z RV s.t.

(i) $Z > 0$ a.s and (ii) $E[Z] = 1$

Then we define a new prob. measure \tilde{P} .

where if A is an event then

$$(\tilde{P}(A) = E[Z \mathbb{1}_A] \stackrel{\text{need}}{=} 1)$$

$$\tilde{P}(A) := E[Z \mathbb{1}_A]$$

we say \tilde{P} is equivalent to P if

$$P(A) = 0 \iff \tilde{P}(A) = 0 \quad \text{and} \quad \tilde{P}(A) = 1 \iff P(A) = 1.$$

Ex Suppose $X \sim \mathcal{N}(0, 1)$, under (P) .

define $d\tilde{P} = e^{\lambda X + \beta} dP$.

$\lambda \in \mathbb{R}$ is fixed. Find $\beta, \gamma \in \mathbb{R}$ s.t.

$X + \gamma \sim \mathcal{N}(0, 1)$ under \tilde{P} .

Sol'n $\quad z > 0 \quad \checkmark$

We need $\mathbb{E}[z] = 1$.

$$\mathbb{E}[z] = e^\beta \mathbb{E}[e^{\lambda z}] = e^\beta + \frac{\lambda^2}{2}$$

$$\text{so } \mathbb{E}[z] = 1 \iff \beta = -\frac{\lambda^2}{2}$$

under IP $X+\delta \sim N(\delta, 1)$

We need $X+\delta \sim N(0, 1)$ under $\tilde{\text{IP}}$.

$$\tilde{\text{IP}} \underset{N(0, 1)}{\sim} (X+\delta \leq x) = \tilde{\text{IP}} (X \leq x-\delta).$$

$$= \int_{-\infty}^{x-\delta} f_X(y) dy \underset{\tilde{\text{IP}}}{\sim} \int_{\{X \leq x-\delta\}} d\tilde{\text{IP}}.$$

$$= \int_{\{X \leq x-\delta\}} z d\text{IP} = \int_{-\infty}^{x-\delta} z f_X(y) dy.$$

$$= \int_{-\infty}^{x-\delta} \left(e^{2y} - \frac{y^2}{2} \right) \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \right) dy.$$

$$= \int_{-\infty}^{x-\delta} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y^2 - 2y + 2^2)} dy.$$

$$= \int_{-\infty}^{x-\delta} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y-\alpha)^2} dy.$$

Let $z = y - \alpha$. $dz = dy$.

$$y \downarrow -\infty, z \downarrow -\infty. \quad y \rightarrow x-\delta, z \rightarrow x-\delta-\alpha.$$

$$= \int_{-\infty}^{x-\delta-\alpha} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2} dz = N(x-\delta-\alpha).$$

We see that $x+\tau \sim N(0,1)$ under \tilde{P}

$$\Leftrightarrow \delta = -\alpha$$

$\therefore x - \alpha \sim N(0,1)$ under \tilde{P}

$$z = e^{\alpha x} - \frac{1}{2}\alpha^2$$

If $\mathbb{E}[z_+] = 1$ then.

\tilde{w}_+ is a BM under $\tilde{\mathbb{P}}$.

$$\text{where } d\tilde{\mathbb{P}} = z_+ d\mathbb{P}$$

i.e. if X_+ is an adapted process then

$$\mathbb{E}[X_+] = \tilde{\mathbb{E}}[X_+ z_+]$$

Comment if $\mathbb{E}[z_+] = 1 \forall 0 \leq t \leq T$.

then \tilde{w}_t is a BM for $t \in [0, T]$
under $\tilde{\mathbb{P}}$.

$$\begin{aligned} \mathbb{P}(X \leq x-\delta) &= N(x-\delta) \\ &= \int_{-\infty}^{x-\delta} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2} dy. \end{aligned}$$

Givs anu.:

$$b = (b_1, \dots, b_d)$$

$w = (w_1, \dots, w_d)$ as standard BM.

$$Z_t = e^{- \int_0^t b(s) \cdot dw_s - \frac{1}{2} \int_0^t |b(s)|^2 ds}.$$

$$\int_0^t b(s) \cdot dw_s = \sum_{i=1}^d b_i(s) dw_i s.$$

$$|b(s)|^2 = \sum_{i=1}^d b_i(s)^2.$$

$$\tilde{w}_+ = w_+ + \int_0^t b(s) ds.$$

if $\alpha_s = -\int_0^s b(s) \cdot dW_s$, "

and " $\frac{1}{2}x^2 = \frac{1}{2}\int_0^t (b(s))^2 ds$ " then this looks similar to our previous calculation.

Ex: $dX_t = \gamma dt + 2\sqrt{X_t} dW_t$. \rightarrow similar processes appear in the Heston Model.

define $Z_t = e^{-\frac{\alpha^2}{2}\int_0^t \sqrt{X_s} dW_s - \frac{\alpha^2}{8}\int_0^t X_s ds}$.

You are told $E[Z_t] = 1$.

Find dX_t under \tilde{P} .

Sol'n. Notice $b_+ = -\frac{\alpha}{2} \sqrt{x_+}$. ($d=1$).

so $dX_+ = r dt + 2\sqrt{x_+} (dw_+ + \underbrace{b_+ dt}_{\tilde{w}_+})$.

$$\begin{aligned} &= (r - 2\sqrt{x_+} b_+) dt + 2\sqrt{x_+} d\tilde{w}_+ \\ &= (r + \alpha x_+) dt + 2\sqrt{x_+} d\tilde{w}_+. \end{aligned}$$

Risk-Neutral Pricing.

Black-Scholes Model:

European option with payoff ~~$V(S_T, t)$~~ V_T .

then time + price is.

$$V_t = \tilde{E} [e^{-r(T-t)} V_T | \mathcal{F}_t]$$

↳ RNM. $\frac{dF_t - r_t}{\sigma_t}$

$\$ S \sim GBM(\mu, \sigma)$ under P
then $S \sim GBM(r, \sigma)$ under \tilde{P} .

If $r = r_t$ can vary over time then.

$$V_t = \tilde{E} [e^{-\int_t^T r_s ds} V_T | \mathcal{F}_t]$$

Ex: Forward Start option.: B-S model.

Fix t_1, T with $0 < t_1 < T$. the payoff
of our option at time T is

$$(S_T - S_{t_1})^+$$

\hookrightarrow strike decided at time t_1 .

but option entered into today. directly.

Note: Cannot just use B-S call formula to compute
time 0 price. since S_{t_1} is unknown today.

Note: typically we try to use independence to "lose" the conditional expectation in the RNP Formula:

Exam Tips.

- 1) Know your formulas and how to use them
 - ↳ Ito (multidim.)
 - ↳ Girsanov
 - ↳ RNP
 - ↳ Levy's Criteria for BM.
 - ↳ Ito Isometry
 - etc...

Regular Call.

$$C_{t_1} = \mathbb{E}^{\tilde{\mathbb{P}}} [e^{-r(T-t_1)} \left(\frac{S_T - K}{S_{t_1}} \right)^+]$$

$$= S_{t_1} \mathbb{E}^{\tilde{\mathbb{P}}} \left[e^{-r(T-t_1)} \left(\frac{S_T}{S_{t_1}} - \frac{K}{S_{t_1}} \right)^+ \right]$$

back to this calculation.

$$= S_0 C(T-t_1, 1)$$

$\mathbb{E}^{\tilde{\mathbb{P}}}$ $\begin{matrix} \text{x strike price.} \\ \text{time initial price of stock.} \\ \text{remaining} \end{matrix}$

$$= C(T-t_1, S_0, S_0).$$

Of course from ⑥ you could use fancy statistician
to get it, but this is lengthy.

We will compute time 0 price.

sol'n. $V_0 = \mathbb{E} [e^{-rT} (S_T - S_{t_i})^+]$

$$= \mathbb{E} [e^{-rt_i} S_{t_i} \left(\frac{S_T}{S_{t_i}} - 1 \right)^+]$$

$$\frac{S_{T_i}}{S_{t_i}} = e^{\sigma (\tilde{w}_T - \tilde{w}_{t_i}) + (r - \frac{1}{2}\sigma^2)(T-t_i)}$$

LL S_{T_i}

$$= \mathbb{E} [e^{-rt_i} S_{t_i}] \mathbb{E} [e^{-r(T-t_i)} \left(\frac{S_T}{S_{t_i}} - 1 \right)^+] \textcircled{O}$$

$$= \underbrace{S_0}_{\text{not } +} \mathbb{E} [e^{-r(T-t_i)} \left(e^{\sigma(\tilde{w}_T - \tilde{w}_{t_i}) + (r - \frac{1}{2}\sigma^2)(T-t_i)}} - 1 \right)^+]$$

~~the~~

2). Think about the answers you get and if they make sense.

E.g. If you are pricing an option and you get a price that can be negative or that it depends on future knowledge then clearly it is incorrect.

We understand that it may be hard to find your mistake in an exam, but if you write down ~~for~~ ~~that~~ that you know the answer cannot be correct, you are more likely to get partial credit.

people spot

e.g. $\mathbb{E}[X_+^2] = 0$. On Q1 midterms, which
is an avoidable mistake.