

last time: Conditional Expectation.

(Ω, \mathcal{G}, P) . $\mathcal{F} \subseteq \mathcal{G}$ σ -subalg.

X is a \mathcal{G} -meas RV,

$E(X|\mathcal{F}) =$ "best approx of X given \mathcal{F} "

\uparrow R.V. \mathcal{F} -meas, & partial av.

~~**~~ $EX \rightarrow$ number. ($\in \mathbb{R}$, non random).

$E(X|\mathcal{F}) \rightarrow$ R.V.

Claim: $E(E(X|\mathcal{F})) = EX$

Proof: Partial av, for every $F \in \mathcal{F}$,

$$\int_F E(X|\mathcal{F}) dP = \int_F X dP$$

Choose $F = \Omega$: $\int_{\Omega} E(X|\mathcal{F}) dP = \int_{\Omega} X dP$

$$\underbrace{\int_{\Omega} E(X|\mathcal{F}) dP}_{E(E(X|\mathcal{F}))} = \underbrace{\int_{\Omega} X dP}_{EX}$$

QED.

Martingales:

① filtration: We say a family of σ -alg
 $\{\mathcal{F}_t \mid t \geq 0\}$ is a filtration
if whenever $s \leq t$, $\mathcal{F}_s \subseteq \mathcal{F}_t$.

Intuition: $\mathcal{F}_t \longrightarrow$ information up to time t

Eg: If X is a process, the filtration generated by X

is defined by $\mathcal{F}_t = \sigma\left(\bigcup_{s \leq t} \sigma(X(s))\right)$.

Adapted: ~~Given~~ A process X is ~~said~~ called
adapted to the filtration $\{\mathcal{F}_t\}$ if.

for every t , $X(t)$ is \mathcal{F}_t measurable.

Clearly: If $\{\mathcal{F}_t | t \geq 0\} =$ filtration generated by X
then X is \mathcal{F}_t meas.

Martingale: We say a process M is a martingale w.r.t
the filtration $\{\mathcal{F}_t | t \geq 0\}$ if

① M is adapted (i.e. $M(t)$ is \mathcal{F}_t meas).

② ~~\mathcal{F}_t~~ Whenever $s \leq t$, $E(M(t) | \mathcal{F}_s) = M_s(s)$

(Martingale — fair game)

① ~~\mathcal{F}_t~~ (a) Sub martingale : $E(M(t) | \mathcal{F}_s) \geq M_s(s)$

② ~~\mathcal{F}_t~~ Super martingale : $E(M(t) | \mathcal{F}_s) \leq M_s(s)$.

Eg 1: BM is a mg. cont the filt gen by BM.

Check: let W be a BM. (std).

$\{\mathcal{F}_t | t \geq 0\}$ Brownian Filtration. $\mathcal{F}_t = \sigma(U_{s \leq t}, W_s)$.

NTS: (1) Adapted (Free). ✓✓

(2) If $s \leq t$, $E(W(t) | \mathcal{F}_s) = W(s)$.

Check: $E(W(t) | \mathcal{F}_s) = E(W(t) - W(s) + W(s) | \mathcal{F}_s)$
 $= \underbrace{E(W(t) - W(s) | \mathcal{F}_s)}_0 + \underbrace{E(W(s) | \mathcal{F}_s)}$

{ By independence $W(t) - W(s)$ is ind of \mathcal{F}_s
& $W(t) - W(s) \sim N(0, t-s)$.

$$\Rightarrow E(W(t) - W(s) | \mathcal{F}_s) = \cancel{0} \\ = E(W(t) - W(s)) = 0$$

Also, $E(W(s) | \mathcal{F}_s) = W(s)$ ($\because W(s)$ is \mathcal{F}_s meas).

$$\Rightarrow E(W(t) | \mathcal{F}_s) = W(s),$$

Q.E.D.

Ex 2°. Let $\{\mathcal{F}_t | t \geq 0\}$ some filt.

Let $\mathcal{F}_\infty = \sigma(\cup \mathcal{F}_t)$. Let Y be any \mathcal{F}_∞ meas RV.

Define $M(t) = E(Y | \mathcal{F}_t)$.

Claim M is a mg.

Pr°. $E(M(t) | \mathcal{F}_s) = E(E(Y | \mathcal{F}_t) | \mathcal{F}_s)$

$\stackrel{\text{tower}}{=} E(Y | \mathcal{F}_s) = M(s)$. QED.

Stochastic Integration:

$S(t)$ \longrightarrow asset spot price at time t

$\Delta(t)$ \longrightarrow Your position on the asset.

Say you only trade at times $t_0 = 0, t_1, t_2, \dots, t_n = T$

Change in wealth/winnings

$$X(T) - X(0) = \sum_{i=0}^{n-1} \Delta(t_i) (S(t_{i+1}) - S(t_i)).$$

Want to trade ctsly in time.

$$\text{Guess } X(T) - X(0) = \lim_{\|P\| \rightarrow 0} \sum_{i=0}^{n-1} \Delta(t_i) (S_{t_{i+1}} - S(t_i))$$

$P \rightarrow$ partition of $[0, T]$. = $\{t_0=0 < t_1 < t_2 \dots < t_n=T\}$

$$\|P\| = \max t_{i+1} - t_i.$$

$$\lim_{\|P\| \rightarrow 0} \sum \Delta(t_i) (S(t_{i+1}) - S(t_i)) = \text{Riemann integral}$$

$$\left(\text{Riemann-Stieltjes. } \int_0^T \Delta(t) dS(t) \right).$$

DOESNT WORK when $S =$ anything reasonable.

The above will only work if S has finite first variation.

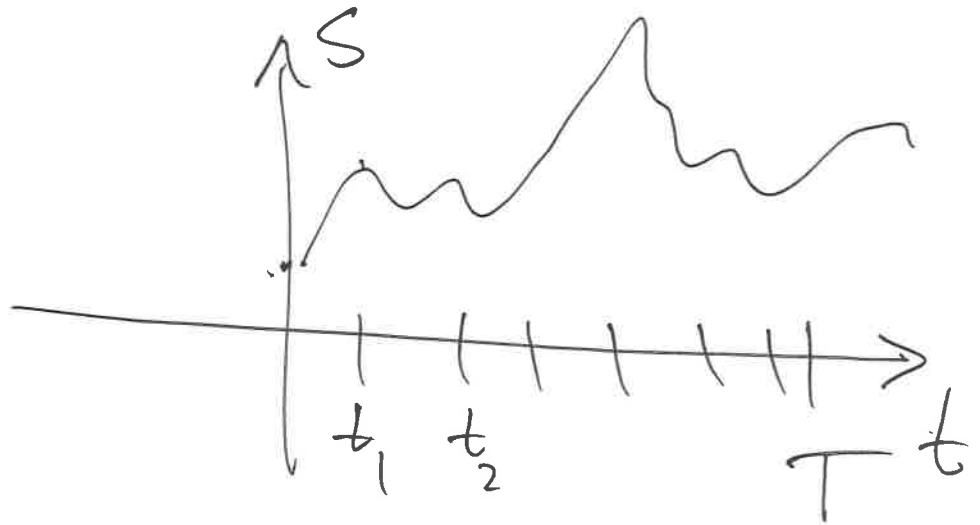
Def: $V_{[0,T]}(S) =$ First variation of S on $[0,T]$.

$$\xrightarrow{\text{def}} \lim_{\|P\| \rightarrow 0} \sum_{i=0}^{n-1} |S(t_{i+1}) - S(t_i)|.$$


You check S diff

$\Rightarrow S$ has finite
first variation.

(Jordan variation),



Claim: B.M does NOT have finite first variation.
(bounded variation).

Pf: $W \rightarrow$ std BM. Fix $T > 0$. 

let $N \in \mathbb{N}$. say $t_i = \frac{iT}{N}$ ($i = 0, 1, \dots, N$).

$$\lim_{N \rightarrow \infty} \mathbb{E} \sum_{i=0}^{N-1} |W(t_{i+1}) - W(t_i)|.$$

$$\begin{aligned} \text{Note } \mathbb{E} \sum |W(t_{i+1}) - W(t_i)| &= \sum \mathbb{E} |W(t_{i+1}) - W(t_i)| \\ &= \sum \mathbb{E} \left| W\left(\frac{(i+1)T}{N}\right) - W\left(\frac{iT}{N}\right) \right|. \end{aligned}$$

$$\text{Know } W\left(\frac{(i+1)T}{N}\right) - W\left(\frac{iT}{N}\right) \sim N\left(0, \frac{T}{N}\right).$$

$$\Rightarrow E \left| W\left(\frac{(i+1)T}{N}\right) - W\left(\frac{iT}{N}\right) \right| = E \left| N\left(0, \frac{T}{N}\right) \right|.$$

$$\stackrel{\text{You check}}{=} C \left(\frac{T}{N}\right)^{1/2}.$$

($C \rightarrow$ some constant).

$$\Rightarrow \sum_{i=0}^{N-1} E \left| W\left(\frac{(i+1)T}{N}\right) - W\left(\frac{iT}{N}\right) \right| = \sum_{i=0}^{N-1} C \left(\frac{T}{N}\right)^{1/2}.$$

$$= \sqrt{NT} \cdot C \xrightarrow{N \rightarrow \infty} +\infty.$$

Quadratic Variation:

Def: Let M be any process.

$[M, M](T) =$ Quadratic variation of M at time T .

$$= \lim_{\|P\| \rightarrow 0} \sum (M(t_{i+1}) - M(t_i))^2$$

Remark: (1) If M has finite first variation, then the Quadratic variation of $M = 0$

(2) If M has finite Quad variation then the first variation of $M = +\infty$.

Claim: $[W, W](T) = T$ (for a standard BM W).

Check: $t_i = \frac{Ti}{n}$, $n \in \mathbb{N}$.

$$\Delta_i W = W(t_{i+1}) - W(t_i) = W\left(\frac{T(i+1)}{n}\right) - W\left(\frac{Ti}{n}\right).$$

Q.v. $[W, W](T) = \lim_{n \rightarrow \infty} \sum_{i=0}^{n-1} (\Delta_i W)^2$

Note $\sum_{i=0}^{n-1} (\Delta_i W)^2 - T = \sum_{i=0}^{n-1} \left((\Delta_i W)^2 - \frac{T}{n} \right)$.

Let $\xi_i = (\Delta_i W)^2 - \frac{T}{n}$

Know ξ_i 's are iid, dist $N(0, \frac{T}{n})^2 - \frac{T}{n}$

$$\Rightarrow E \xi_i = E N(0, \frac{T}{n})^2 - \frac{T}{n} = 0$$

$$\& E \xi_i^2 = E \left(N(0, \frac{T}{n})^2 - \frac{T}{n} \right)^2.$$

(You check) $\underline{\underline{=}} \frac{T^2}{n} \left(E(N(0,1)^4 - 1) \right).$

$$\Rightarrow \sum \xi_i^{\#} = \sum (\Delta_i W)^2 - T$$

(Note $E \left(\sum_{i=0}^{n-1} \xi_i^{\#} \right)^2 = E \sum_{i=0}^{n-1} \xi_i^2 = \sum_{i=0}^{n-1} \frac{T^2}{n} \left(E N(0,1)^4 - 1 \right)$

$$= \frac{T^2}{n} \left(E N(0,1)^4 - 1 \right) \longrightarrow 0$$

$$\Rightarrow \lim_{n \rightarrow \infty} E \left(\sum \xi_i \right)^2 \longrightarrow 0$$

$$\Rightarrow \lim_{n \rightarrow \infty} \sum \xi_i \longrightarrow 0$$

$$\Rightarrow \lim_{n \rightarrow \infty} \left(\sum (\Delta_i W)^2 - t \right) = 0$$

QED

Claim: The Process $M(t) = W(t)^2 - [W, W](t)$ is a mg.

Pf: Knows $[W, W](t) = t$.

$$\Rightarrow M(t) = W(t)^2 - t.$$

Compute $E(M(t) | \mathcal{F}_s)$ (check $= M(s)$).

$$E(M(t) | \mathcal{F}_s) = E(W(t)^2 - t | \mathcal{F}_s).$$

$$= E\left((W(t) - W(s) + W(s))^2 | \mathcal{F}_s\right) - t$$

$$= E\left((W(t) - W(s))^2 + W(s)^2 + 2W(s)(W(t) - W(s)) | \mathcal{F}_s\right) - t$$

$$= t - s + W(s)^2 + 2W(s)E(W(t) - W(s) | \mathcal{F}_s) - t$$

$$= -s + W(s)^2 + 0 = M(s)$$

QED.

Theorem 6.6 (1) If M is a mg w.r.t $\{\mathcal{F}_t\}$.

then $E(M(t)^2) < \infty \iff E[M, M](t) < \infty$.

in this case $M(t)^2 - [M, M](t)$ is also a mg.

(consequently $E M(t)^2 = E M(0)^2 + E [M, M](t)$)

Theorem 6.7 (2) If $A(t)$ is any cts. increasing adapted process.

such that $A(0) = 0$ & $M(t)^2 - A(t)$ is a mg.

then $A(t) = [M, M](t)$.

Construct the Itô integral.

Main Idea: $\lim_{\|P\| \rightarrow 0} \sum \Delta(t_i) (S(t_{i+1}) - S(t_i))$.

DNC in the usual sense.

However, claim $\sum_{i=0}^{n-1} \Delta(t_i) (S(t_{i+1}^n) - S(t_i^n))$.

become mg.