

Spanners in randomly weighted graphs: Euclidean case

Alan Frieze* and Wesley Pegden†
Department of Mathematical Sciences
Carnegie Mellon University
Pittsburgh PA 15213

Abstract

Given a connected graph $G = (V, E)$ and a length function $\ell : E \rightarrow \mathbb{R}$ we let $d_{v,w}$ denote the shortest distance between vertex v and vertex w . A t -spanner is a subset $E' \subseteq E$ such that if $d'_{v,w}$ denotes shortest distances in the subgraph $G' = (V, E')$ then $d'_{v,w} \leq td_{v,w}$ for all $v, w \in V$. We study the size of spanners in the following scenario: we consider a random embedding of $G_{n,p}$ into the unit square with Euclidean edge lengths.

1 Introduction

Given a connected graph $G = (V, E)$ and a length function $\ell : E \rightarrow \mathbb{R}$ we let $d_{v,w}$ denote the shortest distance between vertex v and vertex w . A t -spanner is a subset $E' \subseteq E$ such that if $d'_{v,w}$ denotes shortest distances in the subgraph $G' = (V, E')$ then $d'_{v,w} \leq td_{v,w}$ for all $v, w \in V$. In general, the closer t is to one, the larger we need E' to be relative to E . Spanners have theoretical and practical applications in various network design problems. For a recent survey on this topic see Ahmed et al [1]. Work in this area has in the main been restricted to the analysis of the worst-case properties of spanners. In this note, we assume that edge lengths are random variables and do a probabilistic analysis.

We consider the case where $\ell_{i,j} = |X_i - X_j|$, where $\mathcal{X} = \{X_1, X_2, \dots, X_n\}$ are n randomly chosen points from $[0, 1]^2$. In addition we assume that edges exist between the points in $\mathcal{X}$, independently with probability p . This model, denoted by the graph $\mathcal{X}_p$ of a random embedding of $G_{n,p}$ into $[0, 1]^2$ has been considered in Frieze and Pegden [7], Mehrabian and Wormald [11]. The case of $p = 1$ (i.e. randomly chosen points) is discussed in Section 15.1.2 of Narasimham and Smid [13].

Now $d_{i,j} = |X_i - X_j|$ when $\{i, j\} \in \mathcal{X}_p$ implies that with probability one, a 1-spanner contains $\approx \binom{n}{2}p$ edges. We prove the following:

Theorem 1. *Suppose that the edges of $\mathcal{X}_p$ are given their Euclidean length. Let $\varepsilon > 0$ be fixed. We describe the construction of a $(1 + \varepsilon)$ -spanner E_ε .*

- (a) *If $np / \log n \rightarrow \infty$ then $\mathbb{E}(|E_\varepsilon|) = O(\varepsilon^{-3}(\log 1/\varepsilon)p^{-1}n)$.*
- (b) *If $\frac{1}{p \log 1/p} = o(\log^{1/2} n)$ then $|E_\varepsilon| \leq \mathbb{E}(|E_\varepsilon|) + O(n)$ w.h.p.*
- (c) *If $np \rightarrow \infty$ then w.h.p. any $(1 + \varepsilon)$ -spanner requires $\Omega(n\varepsilon^{-1}p^{-1})$ edges.*

*Research supported in part by NSF grant DMS1952285

†Research supported in part by NSF grant DMS1363136

Assuming $\Omega(n \log n)$ edges we get a much simpler proof. Also, if we only want an $O(n)$ bound on the expected number of edges in our spanner then we can relax the constraint on p to $np/\log n \rightarrow \infty$. I.e. the extra restrictions are needed to prove the high probability result.

Theorem 2. *Suppose that $np^2/\log n \rightarrow \infty$. Then w.h.p. there is a $(1 + \varepsilon)$ -spanner with $O(\varepsilon^{-2}p^{-1}n \log n)$ edges.*

The argument we present for Theorem 1 can be easily adapted to deal with random geometric graphs $G_{\mathcal{X},r}$ for sufficiently large radius r . Here we generate $\mathcal{X}$ as in Theorem 1 and now we join two vertices/points X, Y by an edge if $|X - Y| \leq r$. See Penrose [14] for an early book on this model.

Theorem 3. *If $r^2 \gg \frac{\log n}{n}$ then w.h.p. there is a $(1 + \varepsilon)$ -spanner using $O(n\varepsilon^{-2})$ edges.*

We note finally that Frieze and Pegden [8] have also considered the case where edge lengths are independently exponential mean one. The results there are much tighter.

2 Proof of Theorem 1

Suppose that $0 < \varepsilon \ll 1$. Let

$$r_\varepsilon = \frac{1000 \log^{1/2} 1/\varepsilon}{n^{1/2} \varepsilon^{3/2} p} \text{ and } R_\varepsilon = \left(\frac{10^6 \log n}{\pi n p \varepsilon^3} \right)^{1/2}. \quad (1)$$

The constraint on p means that

$$nR_\varepsilon^2 = n^{o(1)}.$$

Let

$$E_1 = \{\{A, B\} \in \mathcal{X}_p : |A - B| \leq r_\varepsilon\}.$$

We have

$$\mathbb{E}(|E_1|) \leq \binom{n}{2} \pi r_\varepsilon^2 p \leq 10^6 \pi n \log(1/\varepsilon) / (2\varepsilon^3 p) \quad (2)$$

and then we can assert that

$$|E_1| \leq \frac{(10^6 \pi \log 1/\varepsilon) n}{\varepsilon^3 p} \text{ w.h.p.} \quad (3)$$

using the Chebyshev inequality. Here we can use the fact that the events $\{|A - B| \leq R\}$ are pair-wise independent.

For each $A \in \mathcal{X}$ we define τ cones $K(i, A), 0 \leq i < \tau$ with apex A and whose boundary rays make angles $i\varepsilon$ and $(i+1)\varepsilon$ with the horizontal. We then let $Y(i, A)$ denote the closest point in Euclidean distance to A in $K(i, A)$ that is adjacent to A in $\mathcal{X}_p$. We put $Y(i, A) = \perp$ if there is no such Y and let $d_{A,\perp} = \infty$. Also, define $i = i_{A,Y}$ by $Y \in K(i, A)$. Let

$$E_2 = \{(A, Y_{i,A}) : A \in \mathcal{X}, i \in \{0, 1, \dots, \tau-1\}\} \text{ so that } |E_2| = O(n/\varepsilon). \quad (4)$$

For $A, B \in \mathcal{A}$ we let $P_{A,B}$ denote the shortest path between A, B in $\mathcal{X}_p$ and we let $d_{A,B}$ denote the length of $P_{A,B}$. The next two lemmas will discuss the case where A, B are sufficiently distant.

Lemma 4. *If $|A - B| \geq R_\varepsilon$ then with probability $1 - o(n^{-100})$, $|A - Y| \leq \varepsilon|A - B|$, where $Y = Y(i_{A,B}, A)$.*

Proof. We have

$$\mathbb{P}(|A - Y| > \varepsilon |A - B|) \leq (1 - \varepsilon \pi (\varepsilon R_\varepsilon)^2 p / 2)^{n-1} \leq e^{-10^6 \log n / (\pi \varepsilon)} = o(n^{-100}).$$

The 2 in the middle expression allows half the cone to be outside $[0, 1]^2$. $\square$

Lemma 5. *If $r_\varepsilon \leq r = |A - B| \geq R_\varepsilon$ then with probability $1 - o(n^{-2})$, $d_{A,B} \leq (1 + 4\varepsilon)|A - B|$.*

Proof. Let X_1, X_2 be points on the line segment AB at distance $|A - B|/3, 2|A - B|/3$ from A respectively. Let $B_i, i = 1, 2$ be the ball of radius εr centred at X_i . Let A_1 be the set of $\mathcal{X}_p$ neighbors of A in X_1 and let A_2 be the set of $\mathcal{X}_p$ neighbors of B in X_2 . $\mathcal{E}_i, i = 1, 2$ be the event that $|A_i| \geq r^2 np / 10$. Then the Chernoff bounds imply that

$$\mathbb{P}(\mathcal{E}_1 \wedge \mathcal{E}_2) \geq 1 - 2e^{-\pi r^2 np / 1000} = 1 - o(n^{-100}).$$

Let $\mathcal{E}_3$ be the event that there is an $\mathcal{X}_p$ edge between A_1 and A_2 . Then

$$\mathbb{P}(\mathcal{E}_3 \mid \mathcal{E}_1 \wedge \mathcal{E}_2) \geq 1 - (1 - p)^{r^4 n^2 p^2 / 100} = 1 - o(n^{-2}).$$

Finally note that if $\mathcal{E}_i, i = 1, 2, 3$ all occur then $d_{A,B} \leq (1 + 4\varepsilon)|A - B|$. (4 is trivial and avoids any computation.) $\square$

Let

$$\mathcal{B}_\varepsilon = \{(A, B) : d_{A,B} \geq (1 + \varepsilon)|B - A| \text{ and } r = |A - B| \geq r_\varepsilon\} \quad (5)$$

and

$$E_3 = \bigcup_{(A,B) \in \mathcal{B}_\varepsilon} E(P_{A,B}).$$

Let

$$\mathcal{C}_\varepsilon = \{(A, B) : d_{A,B} \leq (1 + \varepsilon)|B - A| \text{ and } r = |A - B| \geq r_\varepsilon \text{ and } |A - Y| \geq \varepsilon |A - B|\},$$

where $Y = Y(i_{A,B}, A)$. Let

$$E_4 = \bigcup_{(A,B) \in \mathcal{C}_\varepsilon} E(P_{A,B}).$$

We show in Lemmas 9 and 12 that the sets E_3, E_4 have linear size w.h.p. Let $E_\varepsilon = \bigcup_{i=0}^4 E_i$.

For $X, Y \in \mathcal{X}$ we let $\hat{d}_{X,Y}$ denote the length of the path from X to Y constructed by the following procedure: Given $A, B \in \mathcal{X}$ where $\{A, B\} \notin E$ we construct a path $A = Z_0, Z_1, \dots, Z_k = B$ as follows: in the following, $Y_j = Y(i, Z_j)$ for $B \in K(i, Z_j), j \geq 0$.

CONSTRUCT:

D1 If $\{Z_j, B\} \in E_1$, use the edge $\{Z_j, B\}$, otherwise

D2 If $|Z_j - Y_j| > |Z_j - B|$ then use $P_{Z_j,B}$ to complete the path, otherwise

D3 If $d_{Y_j,B} \geq (1 + 5\varepsilon)|Y_j - B|$ then use $P_{Z_j,B}$ to complete the path, otherwise

D4 $Z_{j+1} \leftarrow Y_j$.

Remark 1. *We observe that Lemmas 4 and 5 imply that with probability $1 - o(n^{-100})$ we do not use $P_{Z_j,B}$ for $|Z_j - B| \geq 2R_\varepsilon$. Denote the corresponding event by $\mathcal{U}$.*

The next lemma is used to estimate the quality of the path built by CONSTRUCT. (We can obviously replace 5ε by ε in order to get a $(1 + \varepsilon)$ -spanner.)

Lemma 6. Let $A = Z_0, Z_1, \dots, Z_k, Z_{k+1} = B$ be a sequence of points where

- (i) $Z_{j+1} \in K(i_{Z_j, B}, Z_j)$ for all $0 \leq j < k$,
- (ii) $|Z_j - Z_{j+1}| \leq |Z_j - B|$ for $0 \leq j < k$,
- (iii) $d_{Z_k, B} \leq (1 + 5\varepsilon)|Z_k - B|$.

Then

$$d_{Z_k, B} + \sum_{j=0}^{k-1} |Z_{j+1} - Z_j| \leq (1 + 5\varepsilon)|A - B|. \quad (6)$$

Proof. Let $d_j = |Z_j - B|$ for $0 \leq j \leq k$. To prove the lemma, it suffices to show for all j that

$$|Z_{j+1} - Z_j| \leq (1 + 5\varepsilon)(d_j - d_{j+1}). \quad (7)$$

This implies that

$$d_{Z_k, B} + \sum_{j=0}^{k-1} |Z_{j+1} - Z_j| \leq (1 + 5\varepsilon)d_k + (1 + 5\varepsilon) \sum_{j=0}^{k-1} (d_j - d_{j+1}) = (1 + 5\varepsilon)d_0.$$

To this end, define $\bar{Z}_{j+1}$ to the point on the segment $Z_j Z_k$ such that $|\bar{Z}_{j+1} - Z_k| = |Z_{j+1} - Z_k|$. By the assumption that $|Z_j - Z_{j+1}| \leq |Z_j - Z_k|$, we have that $\angle Z_{j+1} Z_k \bar{Z}_{j+1} < \pi/2$, and thus that the ratio

$$\frac{|Z_{j+1} - Z_j|}{d_j - d_{j+1}}$$

can be bounded by considering the case where $\angle Z_{j+1} Z_k \bar{Z}_{j+1} = \pi/2$, as it is drawn in Figure 1.

We have in that case that $\sin \varepsilon = \frac{d_{j+1}}{|Z_j - Z_{j+1}|}$ and $\cos \varepsilon = \frac{d_j}{|Z_j - Z_{j+1}|}$, giving $d_j - d_{j+1} = (\cos \varepsilon - \sin \varepsilon)|Z_j - Z_{j+1}|$ which implies (7). $\square$

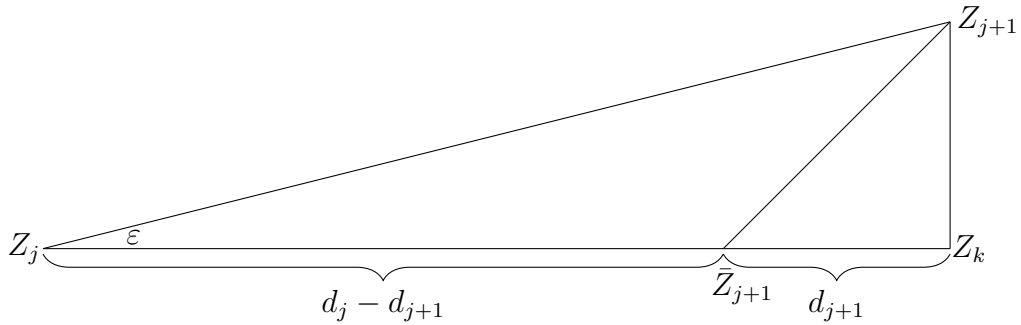


Figure 1:

Assuming (6) we see immediately that the path built by CONSTRUCT has a length within a $1 + 5\varepsilon$ factor of the minimum. We argue next that

Lemma 7. The edges of the paths $P_{Z_j, B}$ used in CONSTRUCT are contained in $E_3 \cup E_4$.

Proof. First consider the path $P = P_{Z_j, B}$ used in D2. If $d_{Z_j, B} \geq (1 + \varepsilon)|Z_j - B|$ then $E(P) \subseteq E_3$. Otherwise, $E(P) \subseteq E_4$.

Now consider the path $P = P_{Z_j, B}$ used in D3. If $d_{Z_j, B} \geq (1 + \varepsilon)|Z_j - B|$ then $E(P) \subseteq E_3$. So assume that $d_{Z_j, B} \leq (1 + \varepsilon)|Z_j - B|$.

If $|Z_j - Y_j| \geq \varepsilon|Z_j - B|$ then $E(P) \subseteq E_4$. So assume that $|Z_j - Y_j| \leq \varepsilon|Z_j - B|$. At this point we have

$$(1 + 5\varepsilon)|Y_j - B| \leq d_{Y_j, B} \leq |Z_j - Y_j| + d_{Z_j, B} \leq (1 + 2\varepsilon|Z_j - B|) \leq (1 + 2\varepsilon)(|Z_j - Y_j| + |Y_j - B|).$$

This implies that $|Z_j - Y_j| \geq 3\varepsilon|Y_j - B|/(1 + 2\varepsilon)$. If $|Y_j - B| \geq |Z_j - B|/2$ then we have $E(P) \subseteq E_4$. So assume that $|Y_j - B| \leq |Z_j - B|/2$. But then $|Z_j - Y_j| \geq |Z_j - B| - |Y_j - B| \geq |Z_j - B|/2$, a contradiction. $\square$

Lemma 8. CONSTRUCT produces a path of length at most $(1 + 6\varepsilon)d_{A, B}$ and only edges of length at most $\varepsilon R_\varepsilon$ contribute to E_3, E_4 .

Proof. We can assume that $|A - B| \geq R_\varepsilon$. Let $A = Z_0, Z_1, \dots, Z_k, Z_{k+1} = B$ be the path constructed and let Z_ℓ be the first point within $\varepsilon R_\varepsilon$ of B . Note that this means $Z_\ell, Z_{\ell+1}, \dots, Z_k$ are all within $\varepsilon R_\varepsilon$ of B . This yields the second claim of the lemma.

Let $\widehat{d}_\ell = \sum_{i=0}^{\ell-1} |Z_{i+1} - Z_i|$ be the length of the subpath from A to Z_ℓ . It follows from Lemma 6 that $\widehat{d}_\ell \leq (1 + 5\varepsilon)|A - Z_\ell|$. Because $|Z_{\ell-1} - B| \geq R_\varepsilon$ and $|Z_{\ell-1} - Z_\ell| \leq \varepsilon|Z_{\ell-1} - B|$ we see that $|Z_\ell - B| \geq (1 - \varepsilon)R_\varepsilon$. It follows from Lemma 5 that $d_{Z_\ell, B} \leq (1 + 4\varepsilon)|Z_\ell - B|$. The path constructed by SECURE therefore has length at most

$$(1 + 5\varepsilon)|A - Z_\ell| + (1 + 5\varepsilon)(1 + 4\varepsilon)\varepsilon R_\varepsilon \leq (1 + 6\varepsilon)|A - B|.$$

$\square$

The next two lemmas bound the expected number of edges in the sets E_3, E_4 .

2.1 $\mathbb{E}(|E_3|)$

Lemma 9. $\mathbb{E}(|E_3|) = O(n/\varepsilon^2)$.

Proof. Fix a pair of points $A, B \in \mathcal{X}$ and let $r = |A - B|$ where $r_\varepsilon \leq r \leq R_\varepsilon$ ($r_\varepsilon, R_\varepsilon$ defined in (5)). Note next that shortest paths are always induced paths. We let $\mathcal{L}_{K, k, A, B}$ denote the set of induced paths from A to B with $k + 1 \geq 2$ edges in $\mathcal{X}_p$, of total length at most $(1 + (K + 1)\varepsilon)r$.

We let $L_{K, k, A, B} = |\mathcal{L}_{K, k, A, B}|$. Then we have

$$|E_3| \leq \sum_{A, B \in \mathcal{X}} \sum_{k, K=1}^{\infty} \sum_{P \in \mathcal{L}_{K, k, A, B}} |P| \cdot 1_{d_{A, B} \geq (1 + K\varepsilon)r}. \quad (8)$$

This is because if $d_{A, B} \geq (1 + \varepsilon)|A - B|$ then the shortest path from A to B has its length in $J_{K, r} = [(1 + K\varepsilon)r, (1 + (K + 1)\varepsilon)r]$, for some $K \geq 1$. Next define, for $L \geq 1$,

$$F(L, \varepsilon) := \sqrt{2L\varepsilon + L^2\varepsilon^2}.$$

Claim 1. *There exists an absolute constant Λ such that for $K \geq 1$,*

$$\mathbb{E}(L_{K, k, A, B} \cdot 1_{d_{A, B} \geq (1 + \varepsilon K)r} \mid |A - B| = r) \leq \left(\frac{\Lambda F(K, \varepsilon)(1 + K\varepsilon)r^2 np(1 - p)^{(k-1)/2}}{k^2} \right)^k e^{-cF(K/4, \varepsilon)(1 + K\varepsilon/4)r^2 np}. \quad (9)$$

Proof of Claim 1: Let $E_{A,B}(L)$ denote the ellipse with centre the midpoint of AB , foci at A, B so that one axis is along the line through AB and the other is orthogonal to it. The axis lengths a, b being given by $a = (1 + L\varepsilon)r$ and $b = r\sqrt{(1 + L\varepsilon)^2 - 1} = rF(L, \varepsilon)$. Thus $E_{A,B}(L)$ is the set of points whose sum of distances to A, B is at most $(1 + L\varepsilon)r$.

Given k points $P_1, \dots, P_k$, the path $P = (A = P_0, P_1, \dots, P_k, P_{k+1} = B)$ is of length at most $(1 + (K + 1)\varepsilon)r$ only if all these points lie in $E_{A,B}(K + 1)$. Thus the length of P is at most the sum $Z_1 + \dots + Z_k$ of independent random variables where Z_i is the distance to the origin of a random point in an ellipse with axes $2a, 2b$ centred at the origin. Here we are using the fact that if a point x lies in an ellipse E then E is contained in a copy of $2E$ centered at x . Indeed, suppose that $(x_i, y_i), i = 1, 2$ are two points in the ellipse $E = \left\{ \frac{x^2}{\xi^2} + \frac{y^2}{\eta^2} \leq 1 \right\}$. Then

$$\frac{(x_1 - x_2)^2}{\xi^2} + \frac{(y_1 - y_2)^2}{\eta^2} \leq \frac{2(x_1^2 + x_2^2)}{\xi^2} + \frac{2(y_1^2 + y_2^2)}{\eta^2} = 2 \sum_{i=1}^2 \left(\frac{x_i^2}{\xi^2} + \frac{y_i^2}{\eta^2} \right) \leq 4. \quad (10)$$

It follows that (x_1, y_1) is contained in a copy of $2E$ centered at (x_2, y_2) .

The following is well-known:

Lemma 10. Z_1 is distributed as $U^{1/2}(a^2 \cos^2(2\pi V) + b^2 \sin^2(2\pi V))^{1/2}$ where U, V are independent uniform $[0, 1]$ random variables.

Proof. □

Now $(2\xi + \xi^2)^{1/2} \leq (1 + \xi)$ and so $F(K + 1, \varepsilon) \leq 1 + (K + 1)\varepsilon$. So, Z_1 is dominated by $U^{1/2}(1 + (K + 1)\varepsilon)r$. It follows from Lemma 10 that $\ell(P)$ is dominated by $(1 + (K + 1)\varepsilon)r$ times the sum of k independent copies of $U^{1/2}$. Lemma 9 of Frieze and Tkocz [9] shows that

$$\mathbb{P}(Z_1 + Z_2 + \dots + Z_k \leq (1 + (K + 1)\varepsilon)r) \leq \frac{2^k}{(2k)!} \leq \frac{e^{2k}}{k^{2k}2^k}. \quad (11)$$

Thus, given k random points $P_1, \dots, P_k$, the probability that $A, P_1, \dots, P_k$ is an induced path of length $\leq (1 + (K + 1)\varepsilon)r$ is at most

$$\left(\frac{\Lambda F(K + 1, \varepsilon)(1 + (K + 1)\varepsilon)r^2 np(1 - p)^{(k-1)/2}}{k^2} \right)^k.$$

(In equation (9), the ratio $\frac{F(K+1, \varepsilon)(1 + (K+1)\varepsilon)}{F(K, \varepsilon)(1 + K\varepsilon)} \leq 4$ has been absorbed into Λ .)

To get (9), we need to also make use of the term $1_{d_{A,B} \geq (1 + \varepsilon K)r}$ in (9).

Case 1: $K\varepsilon \leq 1$: We define two rhombi, R_A, R_B . Let M denote the middle of the segment AB . Then R_A has one diagonal AM and another diagonal $W_A W'_A$ of length $(K\varepsilon)^{1/2}r/10$ that is orthogonal to AM and bisects it. The rhombus R_A is defined similarly. Finally let $\hat{R}_A = R_A \cap [0, 1]^2$ and $\hat{R}_B = R_B \cap [0, 1]^2$. Note that at $\hat{R}_A$ has area at least $1/2$ of the area of $\hat{R}_A$ and similarly for $\hat{R}_B$. Thus if $K \geq 1$ then

$$\text{area}(\hat{R}_A) \geq \frac{(K\varepsilon)^{1/2}r^2}{20} \geq \frac{F(K + 1, \varepsilon)(1 + (K + 1)\varepsilon)r^2}{100}. \quad (12)$$

Here we have used $K\varepsilon \leq 1$ to justify the second inequality.

For a pair of points A, B let $d_{A,B}^*(X)$ denote the minimum length of a path $Q = (A, S, T, B)$ in $\mathcal{X}_p$ where $S \in \hat{R}_A \setminus X$ and $T \in \hat{R}_B \setminus X$. We wish to show that

$$\ell(Q) \leq (1 + K\varepsilon)r \text{ for all choices of } S, T. \quad (13)$$

Now fix S and consider the function $f(T) = |ST| + |TB|$. This is a convex function and so it is maximised at an extreme point of $\widehat{R}_B \setminus X$. Then for a fixed T we find that maximising over S , we have that S must be an extreme point of $\widehat{R}_A \setminus X$. To verify (13), it is enough to check the length

$$\begin{aligned} \ell(AW_A W'_B B) &\leq \left(2 \left(\left(\frac{1}{4} \right)^2 + \left(\frac{(K\varepsilon)^{1/2}}{20} \right)^2 \right)^{1/2} + \left(\left(\frac{1}{2} \right)^2 + \left(\frac{(K\varepsilon)^{1/2}}{10} \right)^2 \right)^{1/2} \right) r \\ &\leq \left(2 \left(\frac{1}{4} \left(1 + \frac{K\varepsilon}{400} \right) + \left(\frac{1}{2} \left(1 + \frac{K\varepsilon}{100} \right) \right) \right) \right) r < \left(1 + \frac{K\varepsilon}{40} \right) r. \end{aligned}$$

This dominates the other 15 possibilities for a pair S, T .

We have that for any X , we have

$$1_{d_{A,B} \geq (1+\varepsilon K)r} \leq 1_{d_{A,B}^*(X) \geq (1+\varepsilon K)r}.$$

The number of points ν_A in $\widehat{R}_A \setminus \{P_i\}$ that are adjacent to A in $\mathcal{X}_p$ is distributed as a binomial with mean at least $c_A F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2(n-k)p$ for some absolute constant $c_A > 0$. This follows from (12). Because $k = o(n)$, w.h.p., as will be discussed below, see (23), we have that $\mathbb{P}(\nu_A \leq c_A F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/2) \leq e^{-c_A F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/2}$. Similarly, $\mathbb{P}(\nu_B \leq c_B F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/2) \leq e^{-c_B F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/2}$ where ν_B is the number of points in $\widehat{R}_B$ that are adjacent to B in $\mathcal{X}_p$. An edge between $\widehat{R}_A$ and $\widehat{R}_B$ will show that $d_{A,B}^* < (1 + K\varepsilon)r$, see (13). We see that there will be at least $\nu_A \nu_B / 2$ possible edges. It follows that if $K \geq 1$ then

$$\begin{aligned} \rho_{k,K,\varepsilon} &= \mathbb{P}(d_{A,B}^* \geq (1 + K\varepsilon)r \mid |A - B| = r, P_1, \dots, P_k) \leq \\ &\quad e^{-c_A F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/20} + e^{-c_B F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/20} + (1-p)^{c_A c_B F((K+1), \varepsilon)^2 (1 + (K+1)\varepsilon)^2 r^4 n^2 p^3 / 4} \\ &\leq e^{-c_A F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/20} + e^{-c_B F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np/20} + (1-p)^{c_A c_B F(K+1, \varepsilon)^2 (1 + (K+1)\varepsilon)^2 r^2 n^2 p^3 / 4} \\ &\quad e^{-c F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np}, \quad (14) \end{aligned}$$

for some absolute constant $c > 0$, since $r^2 np^2 \geq 10^4 \varepsilon^{-2}$.

Case 2: $K\varepsilon \geq 1$: We replace R_A, R_B by two halves S_A, S_B of the rectangle with center M and one side of length $(1 + (K+1)\varepsilon/10)r$ parallel to AB and the other of side $K\varepsilon/10$ orthogonal to AB . Putting $\widehat{S}_A = S_A \cap [0, 1]^2$ and $\widehat{S}_B = S_B \cap [0, 1]^2$ we see that all we need do now is to prove the equivalent of (12) and (13). Then,

$$\text{area}(\widehat{S}_A) \geq \left(1 + \frac{(K+1)\varepsilon}{10} \right) \frac{K\varepsilon}{20} r^2 \geq \frac{F(K+1, \varepsilon)(1 + (K+1)\varepsilon)}{1000} r^2.$$

We have used $K\varepsilon \geq 1$ to justify the second inequality.

We further have that for all $S \in \widehat{S}_A, T \in \widehat{S}_B$ that, using the triangle inequality,

$$\ell(ASTB) \leq \left(1 + \frac{(K+1)\varepsilon}{10} \right) r + 4 \left(\frac{K\varepsilon}{10} + \frac{(K+1)\varepsilon}{10} \right) r < (1 + (K+1)\varepsilon)r.$$

Thus, the probability $\rho_{k,K,\varepsilon}$ defined above satisfies

$$\rho_{k,K,\varepsilon} \leq \left(\frac{\Lambda F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np(1-p)^{(k-1)/2}}{k^2} \right)^k e^{-c F(K+1, \varepsilon)(1 + (K+1)\varepsilon)r^2 np},$$

and the claim follows by linearity of expectation.

End of proof of Claim 1

It will be convenient to replace r by $\frac{\rho}{(np)^{1/2}}$ and write $J_\rho = [\frac{\rho}{n^{1/2}}, \frac{\rho+1}{n^{1/2}}]$ and let $\rho_{\min} = r_\varepsilon(np)^{1/2}$. Then,

$$\begin{aligned}
& \mathbb{E}(|E_3|) \\
& \leq \binom{n}{2} \sum_{\rho=\rho_{\min}}^{\infty} \sum_{K=1}^{\infty} \sum_{k=1}^{n-2} k \left(\frac{\Lambda F(K, \varepsilon)(1 + K\varepsilon)r^2 np(1-p)^{(k-1)/2}}{k^2} \right)^k e^{-cF(K+1, \varepsilon)(1+(K+1)\varepsilon)r^2 np} \mathbb{P}(|A - B| \in J_\rho) \\
& \leq \binom{n}{2} \pi \sum_{\rho=\rho_{\min}}^{\infty} \sum_{K=1}^{\infty} \sum_{k=1}^{n-2} k \left(\frac{\Lambda F(K, \varepsilon)(1 + K\varepsilon)r^2 np(1-p)^{(k-1)/2}}{k^2} \right)^k e^{-cF(K+1, \varepsilon)(1+(K+1)\varepsilon)r^2 np} \left(\frac{2\rho + 1}{n} \right) \\
& \leq 2\pi n \sum_{k=1}^{n-2} k \sum_{K=1}^{\infty} \left(\frac{\Lambda F(K, \varepsilon)(1 + K\varepsilon)(1-p)^{(k-1)/2}}{k^2} \right)^k \sum_{\rho=\rho_{\min}}^{\infty} e^{-cF(K+1, \varepsilon)(1+(K+1)\varepsilon)\rho^2} \rho^{2k+1} \\
& \leq 2\pi n \sum_{k=1}^{n-2} k \sum_{K=1}^{\infty} \left(\frac{\Lambda F(K, \varepsilon)(1 + K\varepsilon)(1-p)^{(k-1)/2}}{k^2} \right)^k \int_{s=0}^{\infty} e^{-cF(K+1, \varepsilon)(1+(K+1)\varepsilon)s} s^k ds \\
& = 2\pi n \sum_{k=1}^{n-2} k \sum_{K=1}^{\infty} \left(\frac{\Lambda F(K, \varepsilon)(1 + K\varepsilon)(1-p)^{(k-1)/2}}{k^2} \right)^k \left(\frac{1}{cF(K+1, \varepsilon)(1+(K+1)\varepsilon)} \right)^{k+1} k! \\
& \leq 2\pi n \sum_{k=1}^{n-2} k \left(\frac{64\Lambda(1-p)^{(k-1)/2}}{k} \right)^k \sum_{K=1}^{\infty} \left(\frac{1}{cF(K+1, \varepsilon)(1+(K+1)\varepsilon)} \right)^{k+1} \\
& = O\left(\frac{n}{\varepsilon^2}\right). \tag{15}
\end{aligned}$$

□

2.2 $\mathbb{E}(|E_4|)$

Lemma 11. *The expected number of $(k+1)$ -edge induced paths of length at most $(1+\varepsilon)r$ from A to B in $\mathcal{X}_p$ can be bounded by*

$$\left(n\pi r^2 p(1-p)^{(k-1)/2} \frac{\varepsilon(1+\varepsilon)^3 e^2}{16k^2} \right)^k (1 - \pi \varepsilon^3 r^2 p)^{n-2}. \tag{16}$$

Proof. Let ρ_k denote the probability that k fixed points $X_1, \dots, X_k$ satisfy that:

- $A = X_0, X_1, \dots, X_k$ is an induced path
- For all $i = 1, \dots, k$, X_i lies in a copy of the ellipse $2 \cdot E_{A,B}$, translated to be centered at X_{i-1} , and
- The total length of the path has total length at most $(1+\varepsilon)r$.

By (10), this is the same as the probability that the the path has total length at most $(1+\varepsilon)r$. Applying (11), we have that

$$\rho_k \leq (2\pi \varepsilon(1+\varepsilon)r^2 p)^k (1-p)^{k(k-1)/2} \left(\frac{e^2(1+\varepsilon)^2}{16k^2} \right)^k.$$

Thus, by linearity of expectation, the number of induced paths $A = X_0, \dots, X_k$ such that

- the total length of the path is at most $(1+\varepsilon)r$, and
- no point off the path lies within distance εr of A in the cone $K(i, A)$

is at most

$$\begin{aligned}
n^k (2\pi\varepsilon(1+\varepsilon)r^2p)^k (1-p)^{k(k-1)/2} \left(\frac{e^2(1+\varepsilon)^2}{16k^2} \right)^k (1-\pi\varepsilon^3r^2p)^{n-k-2} = \\
\left(\frac{n\pi r^2p(1-p)^{(k-1)/2} \varepsilon(1+\varepsilon)^3e^2}{1-\pi\varepsilon^3r^2p} \frac{1}{16k^2} \right)^k (1-\pi\varepsilon^3r^2p)^{n-2} \leq \\
\left(n\pi r^2p(1-p)^{(k-1)/2} \frac{\varepsilon(1+\varepsilon)^3e^2}{16k^2} \right)^k (1-\pi\varepsilon^3r^2p)^{n-2}.
\end{aligned}$$

□

Lemma 12. $\mathbb{E}(|E_4|) = O\left(\frac{n}{\varepsilon^3p}\right)$.

Proof. We have

$$\mathbb{E}(|E_4|) \leq 2\pi \int_{r=r_\varepsilon}^{\infty} \binom{n}{2} \sum_{k=1}^{\infty} k \left(n\pi r^2p(1-p)^{(k-1)/2} \frac{\varepsilon(1+\varepsilon)^3e^2}{16k^2} \right)^k (1-\pi\varepsilon^3r^2p)^{n-2} r dr \quad (17)$$

$$\begin{aligned}
&\leq 2\pi \binom{n}{2} \sum_{k=1}^{\infty} k \int_{r=r_\varepsilon}^{\infty} \left(\frac{\pi\varepsilon r^2np(1-p)^{(k-1)/2}}{k^2} \right)^k e^{-\pi\varepsilon^3r^2np} r dr \\
&\leq \frac{n}{\varepsilon^3p} \sum_{k=1}^{\infty} k \int_{s=A}^{\infty} \left(\frac{\varepsilon(1-p)^{(k-1)/2}s}{\varepsilon^3k^2} \right)^k e^{-s} ds,
\end{aligned} \quad (18)$$

where $A = \pi\varepsilon^2r_\varepsilon^2np = 10^6\varepsilon^{-1}\pi\log(1/\varepsilon)/p$. Now,

$$I_k = \int_{s=A}^{\infty} s^k e^{-s} = k! \sum_{\ell=0}^k \frac{e^{-A} A^\ell}{\ell!} \leq 2e^{-A} A^k, \quad \text{if } k \leq A/2. \quad (19)$$

(Use $I_k = kA^{k-1}e^{-A} + kI_{k-1}$ to obtain the equation.)

Using (19) in (18) we get, for small ε and $k_0 = 10\log_b 1/\varepsilon$ where $b = 1/(1-p)$,

$$\begin{aligned}
\sum_{k=1}^{k_0} k \int_{s=A}^{\infty} \left(\frac{(1-p)^{(k-1)/2}s}{\varepsilon^2k^2} \right)^k e^{-s} ds &\leq Ae^{-A} \sum_{k=1}^{k_0} \left(\frac{A}{\varepsilon^2k^2} \right)^k \leq \\
&Ak_0 \exp \left\{ -\frac{10^6(\log 1/\varepsilon)\pi}{\varepsilon p} + \left(\frac{10^6(\log 1/\varepsilon)\pi}{\varepsilon^2p} \right)^{1/2} \right\} \leq \exp \left\{ -\frac{10^5(\log 1/\varepsilon)\pi}{\varepsilon p} \right\},
\end{aligned} \quad (20)$$

where we have used $(C/x^2)^x \leq e^{2C^{1/2}/e}$ for $C > 0$.

Finally,

$$\sum_{k=k_0+1}^{\infty} k \int_{s=A}^{\infty} \left(\frac{(1-p)^{(k-1)/2}s}{\varepsilon^2k^2} \right)^k e^{-s} ds \leq \int_{s=A}^{\infty} e^{-s} \sum_{k=k_0+1}^{\infty} \left(\frac{2\varepsilon^3s}{k^2} \right)^k ds \leq \int_{s=A}^{\infty} e^{-(1-\varepsilon)s} ds \leq e^{-A/2}. \quad (21)$$

Substituting (20), (21) into (18) we see that $\mathbb{E}(|E_4|) = O\left(\frac{n}{\varepsilon^3p}\right)$. □

We have argued that CONSTRUCT builds a $(1+\varepsilon)$ -spanner w.h.p. The set of edges in this spanner is that of $\bigcup_{i=0}^4 E_i$. Part (a) of Theorem 1 now follows from (2), (4), Lemma 9 and Lemma 12.

2.3 Concentration of measure

Theorem 1 claims a high probability result. We apply McDiarmid's inequality [12] to prove that $|E_3|, |E_4|$ are within range w.h.p. We do not seem to be able to apply the inequality directly and so a little preparation is necessary. We first let $m = \lfloor 1/R_\varepsilon \rfloor$ and divide $[0, 1]^2$ into a grid of m^2 subsquares $\mathcal{C} = (C_1, C_2, \dots, C_{m^2})$ of size $1/m \geq R_\varepsilon$. The Chernoff bounds imply that with probability $1 - o(n^{-101})$ each $C \in \mathcal{C}$ contains at most $\rho_0 = 2nR_\varepsilon^2$ randomly chosen points of $\mathcal{X}$. Suppose that we generate the points one by one and color a point blue if it is one of the first ρ_0 points in its subsquare. Otherwise, color it red. Let $\mathcal{B}$ be the event that all points of $\mathcal{X}$ are blue and we note that

$$\mathbb{P}(\mathcal{B}) = 1 - o(n^{-100}). \quad (22)$$

Let

$$\kappa_1 = \frac{100 \log^{1/2} n}{p}. \quad (23)$$

The significance of κ_1 is that the factors $(1 - p)^{k(k-1)/2}$ in equations (15) and (17) imply that

$$\text{with probability } 1 - o(n^{-100}), \text{ no path contributing to } E_3 \text{ or } E_4 \text{ has more than } \kappa_1 \text{ edges.} \quad (24)$$

We let Z_3 denote the number of edges $e = \{A, B\}$ that satisfy

- (i) A, B are blue.
- (ii) $r_\varepsilon \leq |A - B| \leq 2R_\varepsilon$ and $|Y(i_{A,B}, A) - A| \geq \varepsilon|A - B|$.
- (iii) e is on an induced path in $\mathcal{X}_p$ that has length at least $(1 + \varepsilon)|A - B|$ and at most κ_1 edges, each of length at most R_ε .

Similarly, let Z_4 denote the number of edges $e = \{A, B\}$ that satisfy

- (i) A, B are blue.
- (ii) $r_\varepsilon \leq |A - B| \leq 2R_\varepsilon$.
- (iii) e is on an induced path in $\mathcal{X}_p$ that has length at most $(1 + \varepsilon)|A - B|$ and at most κ_1 edges, each of length at most R_ε .

Let $Z'_i, i = 3, 4$ be defined as for Z_i , without (i). Note that Lemma's 9 and 12 estimate $|E_i|$ through $|E_i| \leq Z'_i$ and showing $\mathbb{E}(Z'_i) = O(n)$. Furthermore, $Z_i = Z'_i, i = 3, 4$ if $\mathcal{U}, \mathcal{B}$ (see Remark 1) occur and these two events occur with probability $1 - o(n^{-100})$. Thus we have for $i = 3, 4$,

$$|E_i| \leq Z_i, \text{ w.h.p.}$$

and

$$E(Z_i) \leq \mathbb{E}(Z'_i \mid \mathcal{B} \cap \mathcal{U}) \mathbb{P}(\mathcal{B} \cap \mathcal{U}) + n^2 \mathbb{P}(\neg \mathcal{B} \vee \neg \mathcal{U}) \leq \mathbb{E}(Z'_i) + n^2 \mathbb{P}(\neg \mathcal{B} \vee \neg \mathcal{U}) = O(n).$$

We will therefore bound the probability that either Z_3 or Z_4 exceeds its mean by n . We let $W = Z_3 + Z_4$. To apply McDiarmid's Inequality we have to establish a Lipschitz bound for W . Our probability space consists of $\times_{i=1}^{m^2} \Omega_i \times \times_{C_j \sim C_k} \Omega_{j,k}$ where Ω_i is a set of at most ρ_0 random points in subsquare C_i together with a list of all of the edges inside C_i . We say that $C_j \sim C_k$ if their boundaries share a common point. Thus for a fixed C_j there are usually 8 subsquares C_k such that $C_j \sim C_k$. The set $\Omega_{j,k}$ determines the edges between points in C_j and C_k . It can be represented by a $\rho_0 \times \rho_0$ $\{0, 1\}$ -matrix in which each entry appears independently with probability p . All in all there are $n^{1-o(1)}$ components of this probability space.

A point $X \in \mathcal{X}$ is in at most $\nu_0 = (9\rho_0)^{\kappa_1} = n^{o(1)}$ of the paths counted by W . So, changing an Ω_i or an $\Omega_{i,j}$ can only change W by at most $\nu_1 = 2\rho_0\nu_0\kappa_1 = n^{o(1)}$ and so the random variable W is ν_1 -Lipschitz.

It then follows from McDiarmid's inequality that

$$\mathbb{P}(W \geq \mathbb{E}(W) + n) \leq \exp \left\{ -\frac{n^2}{2n^{1-o(1)}\nu_1^2} \right\} = e^{-n^{1-o(1)}}.$$

This verifies the existence of the claimed $(1 + \varepsilon)$ -spanner and now we argue that w.h.p. we need $\Omega(n(\varepsilon p)^{-1})$ edges. We say that an edge $\{A, B\}$ is *lonely* if its length is r and there are no $\mathcal{X}_p$ -adjacent points in the ellipse C with foci A, B defined by $|X - A| + |X - B| \leq (1 + \varepsilon)r$. Any $(1 + \varepsilon)$ -spanner must contain all lonely edges. Now the volume of C is $\pi\varepsilon(1 + \varepsilon)r^2/2$. By concentrating on points that are at least 0.1 from the boundary ∂D of $D = [0, 1]^2$, we see that the expected number of lonely edges is at least

$$(0.64 - o(1)) \binom{n}{2} p \int_{r=0}^{0.8\sqrt{2}} \left(1 - \frac{\pi\varepsilon(1 + \varepsilon)pr^2}{2} \right)^n 2\pi r dr \geq \pi \binom{n}{2} \cdot \frac{1}{\pi\varepsilon(1 + \varepsilon)p} \int_{s=0}^{\varepsilon p} (1 - s)^n ds \approx \frac{n}{2\varepsilon(1 + \varepsilon)p}.$$

This completes the proof of Theorem 1.

3 Proof of Theorem 2

Let $K(i, X), Y(i, X)$ be as in Section 2 and let $M = 4/\varepsilon^2$ and

$$E_5 = \left\{ X, Y \in E : |X - Y| \leq \left(\frac{M \log n}{np} \right)^{1/2} \right\}.$$

The Chebyshev inequality implies that w.h.p. $|E_5| = O((\varepsilon^2 p)^{-1} n \log n)$. Let E_1, E_2 be as in Section 2.

Suppose first that X is at least $\left(\frac{M \log n}{np} \right)^{1/2}$ from the boundary.

$$\mathbb{P} \left(\nexists Y(i, X) \text{ or } |X - Y(i, X)| \geq \left(\frac{M \log n}{np} \right)^{1/2} \right) \leq \left(1 - \frac{\pi\varepsilon M \log n}{2np} \right)^{n-1} = o(n^{-2}). \quad (25)$$

If X is closer to the boundary, then (25) may not be true for some i . This could be the case where X, Y are both close to ∂D and $\{X, Y\} \notin E_1$. In these cases we merge $K(i, X)$ with $K(i + 1, X)$ or $K(i - 1, X)$, which ever is appropriate. The net effect is to replace ε by 2ε for this particular cone. This has negligible effect.

Assume that $|A - B| \geq \left(\frac{M \log n}{np} \right)^{1/2}$. It follows that w.h.p. we can go $A \rightarrow Y(i, B) \rightarrow Y(i, Y(i, B)) \rightarrow \dots$, at least until we are at Z within $\left(\frac{M \log n}{np} \right)^{1/2}$ of B . Then with probability at least $1 - \left(1 - \frac{\pi M \log n}{4n} \right)^{n-2}$ we can find a path ZXB from Z to B using only edges from E_1 , of total length at most $2 \left(\frac{M \log n}{np} \right)^{1/2}$. The length of the path from A to B is then at most

$$\left(1 + \frac{\varepsilon}{2} \right) |A - B| + 2 \left(\frac{M \log n}{np} \right)^{1/2} = |A - B| \left(1 + \frac{\varepsilon}{2} + \frac{2}{M} \right) \leq (1 + \varepsilon) d_{A,B}.$$

The validity of the first term $\left(1 + \frac{\varepsilon}{2} \right) |A - B|$ follows from Lemma 6.

4 Proof of Theorem 3

For this we only have to observe that w.h.p. $K(X, i)$ exists for all X, i . This follows from the Chernoff bounds and the fact that the expected number of vertices in $K(X, i)$ grows faster than $\log n$. We can therefore use Lemma 6 to prove the existence of the required spanner.

5 Summary and open questions

We have considered a Euclidean version, asking for a $(1 + \varepsilon)$ -spanner and random geometric graphs. We could perhaps extend the results of Theorems 1, 2, 3 to $[0, 1]^d, d \geq 3$. This does not seem impossible. There is a slight problem in that the cones $K(i, X)$ intersect in sets of positive volume. The intersection volumes are relatively small and so the problems should be minor. We do not claim to have done this.

There are a number of related questions one can tackle:

1. We could replace edge lengths by E_2^s where $s < 1$. This would allow us to generalise edge lengths to distributions with a density f for which $f(x) \approx x^{1/s}$ as $x \rightarrow 0$. This is a more difficult case than $s = 1$ and it was considered by Bahmidi and van der Hofstadt [3]. They prove that w.h.p. $d_{1,2}$ grows like $\frac{n^s}{\Gamma(1+1/s)^s}$ where Γ denotes Euler's Gamma function. The analysis is more complex than that of [10] and it is not clear that our proof ideas can be generalised to handle this situation.
2. Can we reduce the reliance on ε in the upper bound to ε^{-1} ?

References

- [1] R. Ahmed, G. Bodwin, F. Sahneh, K. Hamm, M. Javad, S. Kobourov and R. Spence, Graph Spanners: A Tutorial Review.
- [2] N. Alon and J. Spencer, The Probabilistic Method, Third Edition, Wiley and Sons, 2008
- [3] S. Bahmidi and R. van der Hofstadt, Weak disorder asymptotics in the stochastic mean-field model of distance, *Annals of Applied Probability* 22 (2012) 29-69.
- [4] A note on two problems in connexion with graphs, *Numerische Mathematik* 1 (1959) 269-271.
- [5] A.M. Frieze, The effect of adding randomly weighted edges.
- [6] A.M. Frieze and M. Karoński, Introduction to Random Graphs, Cambridge University Press, 2015.
- [7] A.M. Frieze and W. Pegden, Travelling in randomly embedded random graphs, *Random Structures and Algorithms* 55 (2019) 649-676
- [8] A.M. Frieze and W. Pegden, Spanners in randomly weighted graphs: independent edge lengths.
- [9] A.M. Frieze and T. Tkocz, Shortest paths with a cost constraint: a probabilistic analysis.
- [10] S. Janson, One, two and three times $\log n/n$ for paths in a complete graph with random weights, *Combinatorics, Probability and Computing* 8 (1999) 347-361.
- [11] A. Mehrabian and N. Wormald, On the Stretch Factor of Randomly Embedded Random Graphs, *Discrete & Computational Geometry* 49 (2013) 647-658.

- [12] C. McDiarmid, On the method of bounded differences, *Surveys in combinatorics* 141 (1989) 148-188.
- [13] G. Narasimhan and Smid, Geometric Spanner Networks, Cambridge University Press, 2007.
- [14] M. Penrose, Random Geometric Graphs, Oxford University Press, 2003.