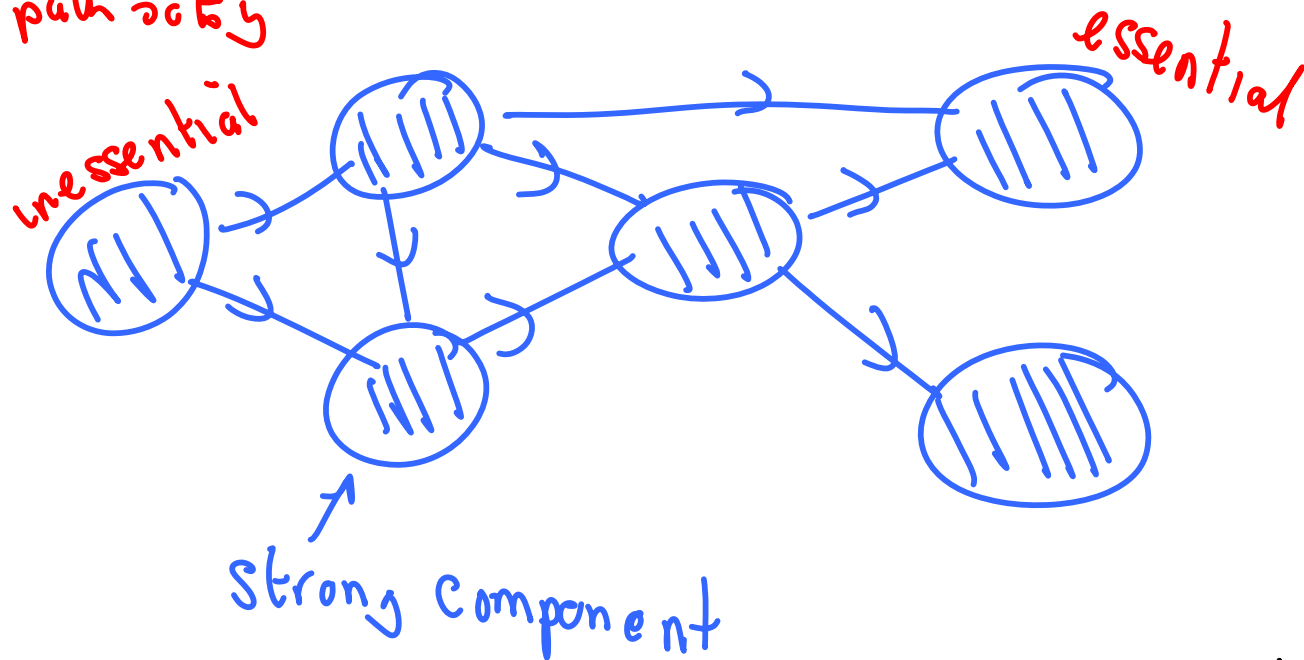


1/20/10

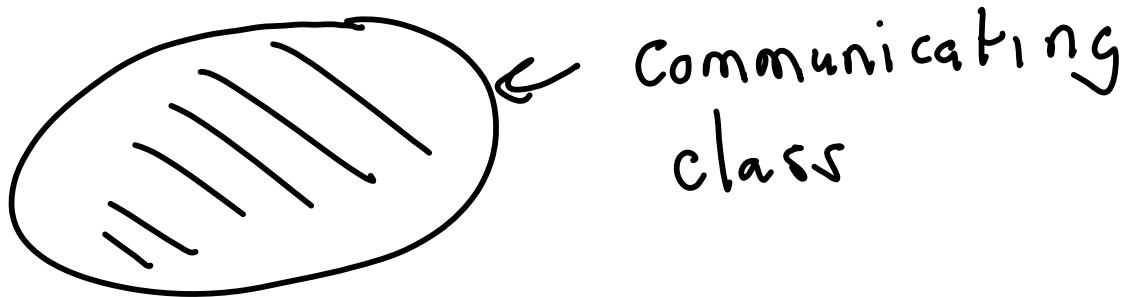
1.7

$x \in \Sigma$ is essential if

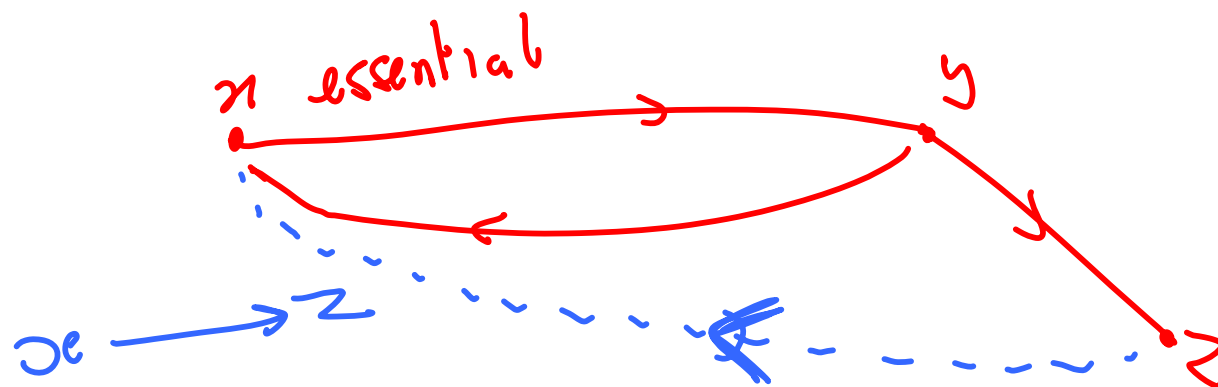
$x \rightarrow y$ implies $y \rightarrow x$
if path to y



Communicating Class $\equiv$ strong component
Defined by equivalence relation $x \sim y : x \rightarrow y \wedge y \rightarrow x$

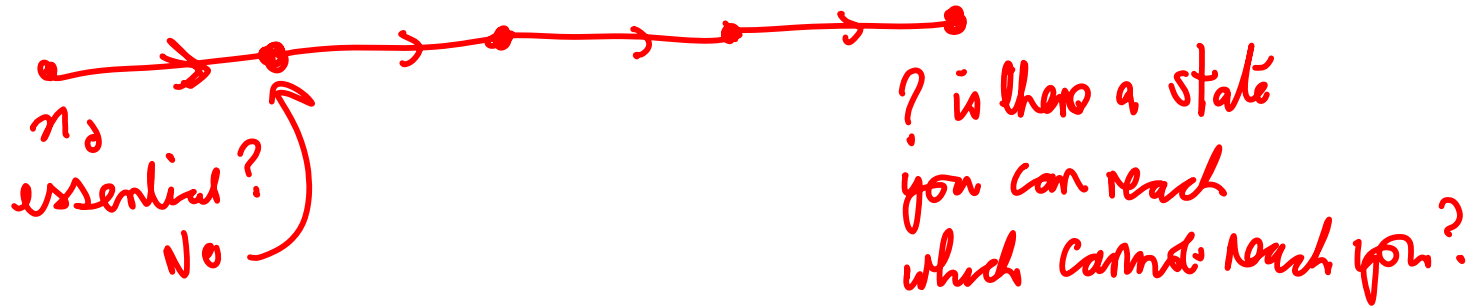


∴ All essential $\sim$ all inessential



L 1.24

Every chain has at least one essential
class. (state)



P1.25

If π is stationary then $\pi(x) = 0$
for all essential states.

Pf

$\mathcal{C}$ = essential class.

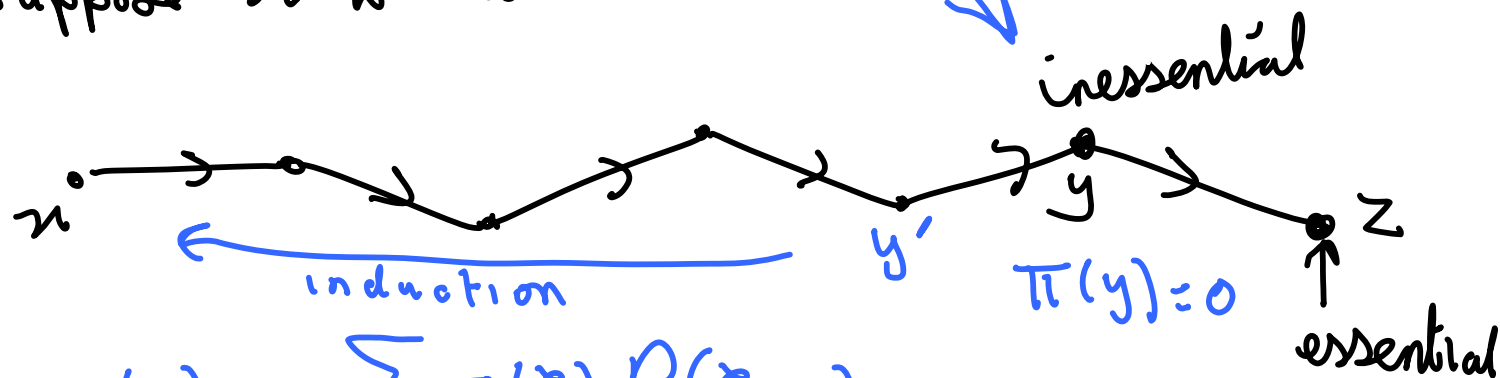
$$\begin{aligned}\pi P(\mathcal{C}) &= \sum_{z \in \mathcal{C}} (\pi P)(z) \\ &= \sum_{z \in \mathcal{C}} \left(\sum_{y \in \mathcal{C}} \pi(y) P(y, z) + \sum_{y \notin \mathcal{C}} \pi(y) P(y, z) \right) \\ &= \sum_{y \in \mathcal{C}} \pi(y) \overbrace{\sum_{z \in \mathcal{C}} P(y, z)}^{=1} + \sum_{z \in \mathcal{C}} \sum_{y \notin \mathcal{C}} \pi(y) P(y, z)\end{aligned}$$

$$\overbrace{\pi P(\mathcal{C})}^{\text{same}} = \pi(\mathcal{C}) + \underbrace{\sum_{z \in \mathcal{C}} \sum_{y \notin \mathcal{C}} \pi(y) P(y, z)}_{= 0}$$

$\pi P = \pi$

$$\pi(y) P(y, z) = 0 \quad z \in \mathcal{C}, y \notin \mathcal{C}$$

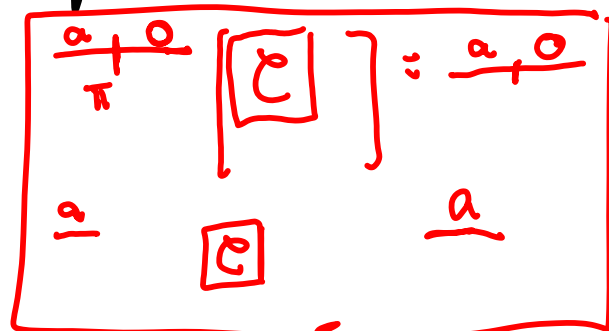
Suppose π is inessential.



$$0 = \pi(y) = \sum_{\xi \in \Omega} \pi(\xi) P(\xi, y) \geq \underbrace{\pi(y')}_{= 0} P(y', y)$$

P1.26

The stationary distribution π is unique iff $\exists$ a unique essential class.



Pf

Unique class $\mathcal{C}$.

$P|_{\mathcal{C}}$ is irreducible $\Rightarrow \pi|_{\mathcal{C}}$ is unique.

Markov Chain Restricted to $\mathcal{C}$

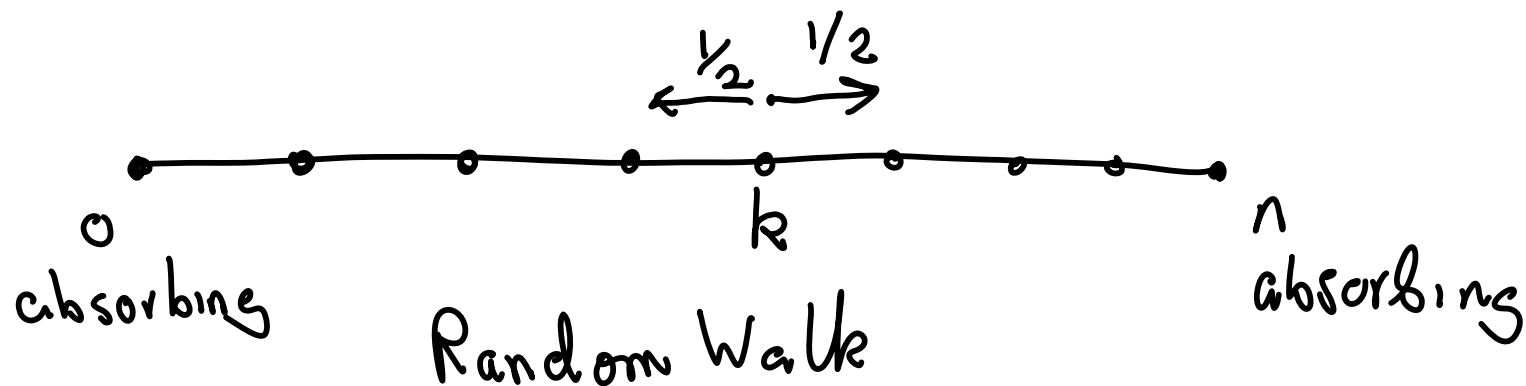
Suppose π is stationary, then $\pi(y) = 0$, $y \notin \mathcal{C}$ and π is stationary for $P|_{\mathcal{C}}$

must equal $\pi|_{\mathcal{C}}$

Two essential states $\mathcal{C}_1, \mathcal{C}_2$ give
rise to 2 stationary distributions

Chapter 2

1) Gambler's Ruin



X_t = fortune at time t

τ = absorption time.

$$P_k = P_k(X_\tau = n) = \frac{k}{n}$$
$$E_k(\tau) = k(n-k)$$

Pf

$$(i) \quad P_k = \frac{1}{2} P_{k-1} + \frac{1}{2} P_{k+1} \quad P_0 = 0; P_n = 1$$

$$\begin{aligned} P_k(X_T = n) &= P_k(X_T = n \mid X_1 = k-1) P_1(X_1 = k-1) \\ &\quad + P_k(X_T = n \mid X_1 = k+1) P_1(X_1 = k+1) \end{aligned}$$

$X_0 = k$ (pointing to the k in P_k)

Red annotations: A bracket under the first term is labeled $P_{k+1}(X_T = n)$. A bracket under the second term is labeled $\frac{1}{2}$. A red arrow points from the $\frac{1}{2}$ bracket to the $P_1(X_1 = k+1)$ term.

$$(ii) \quad f_k = E_k(T) = \frac{1}{2}(1 + f_{k-1}) + \frac{1}{2}(1 + f_{k+1})$$

$0 < k < n$

$$f_0 = 0$$

$$f_n = 0$$

$k(n-k)$ is the solution.

Coupon Collecting

As a Markov Chain

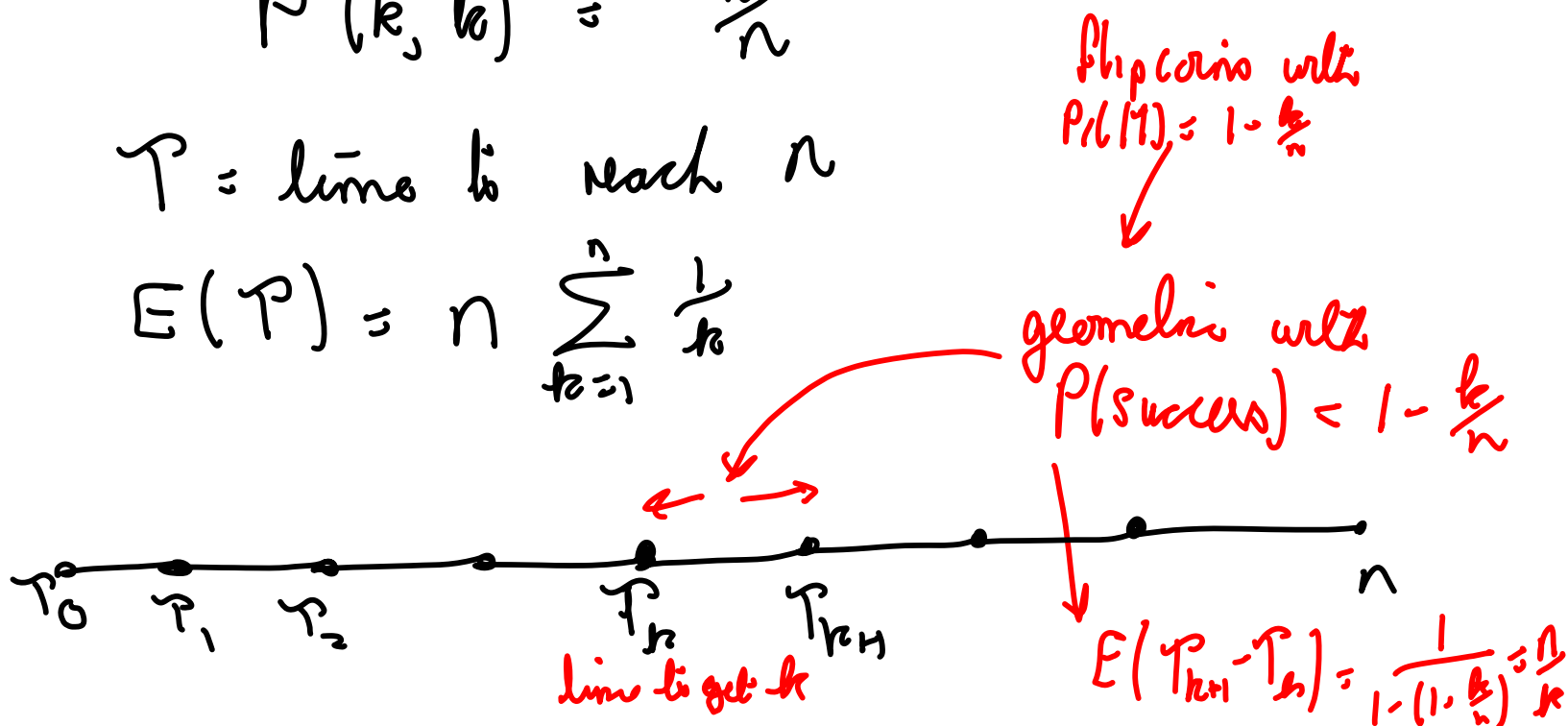
X_t = # of distinct coupons at time t

$$P(k, k+1) = \frac{n-k}{n}$$

$$P(k, k) = \frac{k}{n}$$

T = time to reach n

$$E(T) = n \sum_{k=1}^n \frac{1}{k}$$



2.3

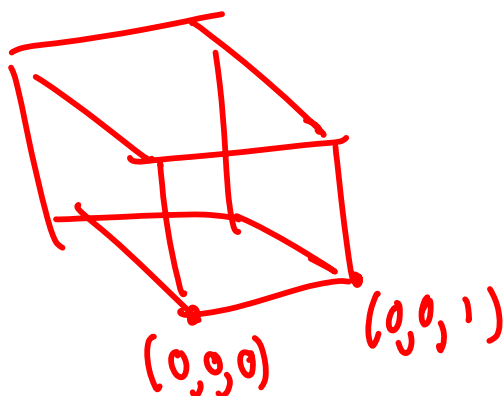
Hypercube and Ehrenfest Urn model.

$$Q_n = \{0,1\}^n; \quad x, y \in Q_n, \quad h(x, y) = \#\{i: x_i \neq y_i\}$$

↑
Hamming
Distance

Graph: Vertices Q_n

$$x, y \leftrightarrow h(x, y) = 1$$



Random Walk on Cube

Lazy Random Walk

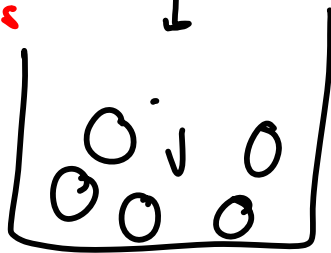
$$P[x, x] = \frac{1}{2}$$

$$P[x, y] = \frac{1}{2n} \quad h(x, y) = 1$$

$$\text{Walk} = Z_0, Z_1, \dots, Z_n, \dots$$

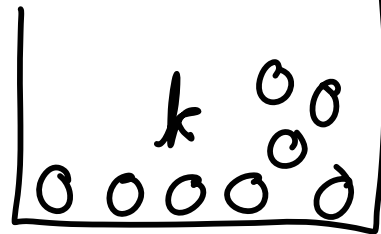
Ehrenfest Urn

$\# \text{ balls in } Z_t$



II

$\# \text{ balls in } Z_t$



$j+k=n$

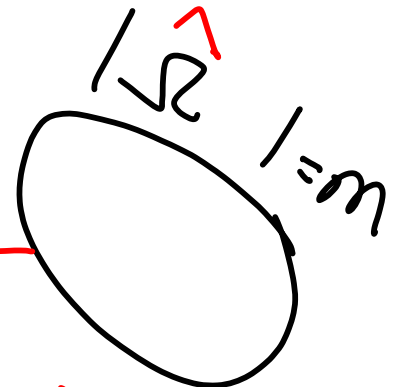
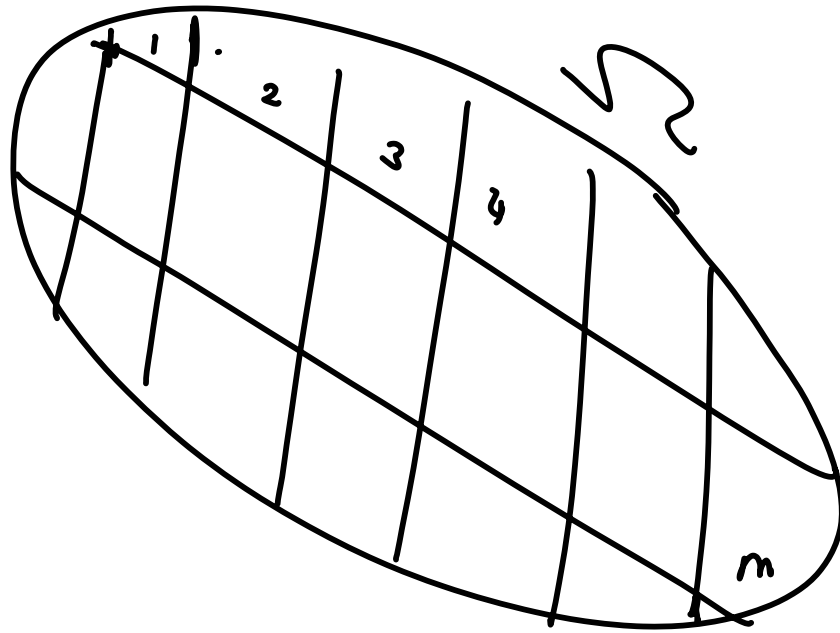
Pick random ball and move to opposite urn.

$X_t = \# \text{ balls in urn I.}$

$$P[j, k] = \begin{cases} 1 - \frac{j}{n} & k = j+1 \\ \frac{j}{n} & k = j-1 \\ 0 & \text{otherwise} \end{cases}$$

$X_t = \text{len}(Z_t, 0)$

$U_m \equiv$ **projection** of RW on Q_n .

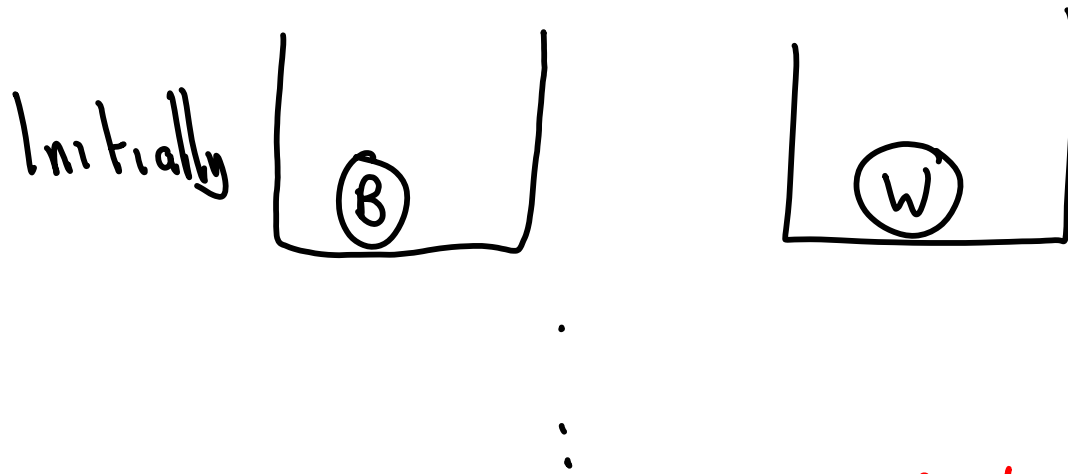


$$\hat{P}([x], [y]) = \sum_{z \in [y]} P(x, z)$$

$\square \leftarrow \{ \text{state with same number of } \downarrow \text{'s} \}$

must be independent of $\xi \in [x]$

Polya Urn



Choose a random ball
Add a new ball of that color.

$B_k = \#$ of black balls at time k

Thm: B_k is uniform over $\{1, 2, \dots, k+1\}$

Proof by induction

$$P_i(B_k = j) = \underbrace{P_i(B_k = j | B_{k-1} = j)}_{1 - j/k+1} \underbrace{P_i(B_{k-1} = j)}_{\text{induction: } 1/k} + \underbrace{P_i(B_k = j | B_{k-1} = j-1)}_{j-1/k+1} \underbrace{P_i(B_{k-1} = j-1)}_{1/k}$$

$$= \left(1 - \frac{j}{k+1}\right) \cdot \frac{1}{k} + \frac{j-1}{k+1} \cdot \frac{1}{k}$$

$$= \left(1 - \frac{1}{k+1}\right) \cdot \frac{1}{k}$$

$$= \frac{1}{k+1}$$