## **Directed Scale-free Graphs**

#### B. Bollobás, C. Borgs, J. Chayes, O. Riordan

### April 11

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# Outline



2 Analysis of the model



## Other models

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# Outline







## Other models

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# Outline



#### Description of the model





## Other models

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# Outline



#### Description of the model







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# Description of the model

- Graph G(t<sub>0</sub>) = G<sub>0</sub> with n<sub>0</sub> vertices and t<sub>0</sub> edges and constants α, β, γ, δ<sub>in</sub> and δ<sub>out</sub> s.t. α + β + γ = 1
- At timestep t, perform one of the following
  - w.p. α add a new vertex v and an edge (v, w) to an existing vertex w

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- w.p.  $\beta$  add an edge (v, w) to the existing graph
- w.p. γ add a new vertex v and an edge (w, v) from an existing vertex w

# Analysis of the model

- G(t)- Graph at time t
- n(t) # of vertices at time t
- $x_i(t) \#$  of nodes with in-degree *i* at time *t*
- $y_i(t) \#$  of nodes with out-degree *i* at time *t*

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# Analysis of the model(contd.)

**Theorem**: Let  $i \ge 0$  be fixed. There are constants  $p_i$  and  $q_i$  such that  $x_i(t) = p_i t + o(t)$  and  $y_i(t) = q_i t + o(t)$  hold with probability 1. Furthermore, if  $\alpha \delta_{in} + \gamma > 0$  and  $\gamma < 1$ , then as  $i \to \infty$  we have

$$p_i \sim C_{IN} i^{-X_{IN}}$$
 .

If  $\gamma \delta_{out} + \alpha > 0$  and  $\alpha < 1$ , then as  $i \to \infty$  we have

$$q_i \sim \mathcal{C}_{OUT} i^{-X_{OUT}}$$
 .

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# Analysis of the model(contd.)

**Observation**: Number of vertices, n(t) is  $n_0$  plus a Binomial distribution with mean  $(\alpha + \gamma)(t - t_0)$ .

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# **Chernoff Bound**

Let  $X_1, X_2, ..., X_n$  be independent Bernoulli trials such that, for  $1 \le i \le n$ ,  $\Pr[X_i = 1] = p$ , where  $0 . Then, for <math>X = \sum_{i=1}^n X_i, \mu = E[X] = np$ , and any  $0 < \delta < 1$ ,

$$\mathsf{Pr}[X > (1+\delta)\mu] < \left[rac{e^{\delta}}{(1+\delta)^{1+\delta}}
ight]^{\mu}$$

$$\Pr[X < (1-\delta)\mu] < e^{-\mu\delta^2/2}$$

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# Analysis of the model(contd.)

#### By the Chernoff bound, it follows that

$$\Pr[|n(t) - (\alpha + \gamma)t| \ge t^{1/2}\log t] \le e^{-c(\log t)^2}$$

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In particular, the probability above is  $o(t^{-1})$  as  $t \to \infty$ .

# Analysis of the model(contd.)

- Now, we will estimate x<sub>i</sub>(t) # of vertices of in-degree i at time t.
- In going from G(t) to G(t + 1), what is the chance of destroying a vertex of degree i?

• The answer

$$(\alpha + \beta) x_i(t) \frac{i + \delta_{in}}{t + \delta_{in} n(t)}.$$

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# Analysis of the model(contd.)

#### Putting everything together, we get

# $\mathbf{E}[x_{i}(t+1)|G(t)] = x_{i}(t) + \frac{(\alpha+\beta)}{t+\delta_{in}n(t)} \Big( (i-1+\delta_{in})x_{i-1}(t) - (i+\delta_{in})x_{i}(t) \Big) + \alpha \mathbf{1}_{\{i=0\}} + \gamma \mathbf{1}_{\{i=1\}}.$

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 $x_{-1}(t)=0.$ 

# Analysis of the model(contd.)

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Analysis of the model(contd.)

By the Chernoff bound, it follows that

$$\Pr[|n(t) - (\alpha + \gamma)t| \ge t^{1/2}\log t] \le e^{-c(\log t)^2}.$$

In particular, the probability above is  $o(t^{-1})$  as  $t \to \infty$ .

We could assume that w.p.  $1 - o(t^{-1})$ ,

$$|n(t) - (\alpha + \gamma)t| = o(t^{3/5}).$$

$$(\alpha + \beta) \frac{i + \delta_{in}}{t + \delta_{in} n(t)} x_i(t) = O(1)$$
 true for  $n(t) \ge 0$ 

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We could assume that w.p.  $1 - o(t^{-1})$ ,

$$|n(t)-(\alpha+\gamma)t|=o(t^{3/5}).$$

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# Analysis of the model(contd.)

$$\mathbf{E}\left[\frac{\alpha+\beta}{t+\delta_{in}n(t)}(i+\delta_{in})x_i(t)\right]$$
  
=  $\frac{\alpha+\beta}{t+\delta_{in}(\alpha+\beta)t}(i+\delta_{in})\mathbf{E}x_i(t)\left(1+o(t^{-2/5})\right)$   
=  $\frac{\alpha+\beta}{t+\delta_{in}(\alpha+\beta)t}(i+\delta_{in})\mathbf{E}x_i(t)+o(t^{-2/5})$ 

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# Analysis of the model(contd.)

$$\mathbf{E}[x_i(t+1)|G(t)] = x_i(t) + \frac{(\alpha+\beta)}{t+\delta_{in}n(t)} \Big( (i-1+\delta_{in})x_{i-1}(t) - (i+\delta_{in})x_i(t) \Big) + \alpha \mathbf{1}_{\{i=0\}} + \gamma \mathbf{1}_{\{i=1\}}.$$

Taking expectation of both sides,

$$\begin{aligned} \mathbf{E} x_i(t+1) = \mathbf{E} x_i(t) &+ \\ &\frac{(\alpha+\beta)}{t+\delta_{in}(\alpha+\beta)t} \Big( (i-1+\delta_{in}) \mathbf{E} x_{i-1}(t) - (i+\delta_{in}) \mathbf{E} x_i(t) \Big) \\ &+ \alpha \mathbf{1}_{\{i=0\}} + \gamma \mathbf{1}_{\{i=1\}} + o(t^{-2/5}). \end{aligned}$$

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Analysis of the model(contd.)

#### As in the Preferential attachment model, we will assume

 $\mathbf{E} x_i(t) \sim p_i t$ 

Now we obtain a recurrence relation in  $p_i$ 's.

$$p_{i}(t+1) = p_{i}t + \frac{(\alpha+\beta)}{t+\delta_{in}(\alpha+\beta)t} \Big( (i-1+\delta_{in})p_{i-1}t - (i+\delta_{in})p_{i}t \Big) + \alpha \mathbf{1}_{\{i=0\}} + \gamma \mathbf{1}_{\{i=1\}}.$$

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# Analysis of the model(contd.)

Let

$$c_1 = \frac{(\alpha + \beta)}{1 + \delta_{in}(\alpha + \beta)}.$$

Let  $p_{-1} = 0$  and for  $i \ge 0$ ,

$$p_i = c_1 \left( (i - 1 + \delta_{in}) p_{i-1} - (i + \delta_{in}) p_i \right) \\ + \alpha \mathbf{1}_{\{i=0\}} + \gamma \mathbf{1}_{\{i=1\}}.$$

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Analysis of the model(contd.)

Solving the recurrence, we get  $p_0 = \alpha/(1 + c_1 \delta_m)$ ,

$$\boldsymbol{\rho}_1 = (1 + \delta_m + \boldsymbol{c}_1^{-1})^{-1} \left( \frac{\alpha \delta_i \boldsymbol{n}}{1 + \boldsymbol{c}_1 \delta_{in}} + \frac{\gamma}{\boldsymbol{c}_1} \right),$$

and for  $i \ge 1$ ,

$$p_{i} = \frac{(i-1+\delta_{in})_{i-1}}{(i+\delta_{in}+c_{1}^{-1})_{i-1}}p_{1}$$
  
=  $\frac{(i-1+\delta_{in})!}{(i+\delta_{in}+c_{1}^{-1})!}\frac{(1+\delta_{in}+c_{1}^{-1})!}{\delta_{in}!}p_{1}.$ 

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# Analysis of the model(contd.)

Wanted to show that

$$p_i \sim C_{IN} i^{-X_{IN}}$$

$$X_{IN} = (\delta_{in} + c_1^{-1}) - (\delta_{in} + c_1^{-1}) = 1 + c_1^{-1}$$

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$$X_{OUT} = (\delta_{out} + c_2^{-1}) - (\delta_{out} + c_2^{-1}) = 1 + c_2^{-1}$$
  
where  $c_1 = \frac{\alpha + \beta}{1 + \delta_{in}(\alpha + \gamma)}$  and  $c_2 = \frac{\beta + \gamma}{1 + \delta_{out}(\alpha + \gamma)}$ .

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Analysis of the model(contd.)

#### Theorem:

$$\mathbf{E}x_{i}(t) - p_{i}t| = O(t^{3/5}).$$
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Proof: Recurrence for expectation

$$\begin{aligned} \mathbf{E}x_{i}(t+1) = \mathbf{E}x_{i}(t) &+ \\ & \frac{(\alpha+\beta)}{t+\delta_{in}(\alpha+\beta)t} \Big( (i-1+\delta_{in})\mathbf{E}x_{i-1}(t) - (i+\delta_{in})\mathbf{E}x_{i}(t) \Big) \\ &+ \alpha \mathbf{1}_{\{i=0\}} + \gamma \mathbf{1}_{\{i=1\}} + o(t^{-2/5}). \end{aligned}$$

$$\begin{aligned} \text{Writing } \mathbf{E}x_{i}(t) = p_{i}t + \epsilon_{i}(t), \\ \epsilon_{i}(t+1) = \Big( \frac{c_{1}(i-1+\delta_{in})}{t} \Big) \epsilon_{i-1}(t) + \Big( 1 - \frac{c_{1}(i+\delta_{in})}{t} \Big) \epsilon_{i}(t) + o(t^{-2/5}). \end{aligned}$$

Analysis of the model(contd.)

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Writing  $\mathbf{E}x_{i}(t) = p_{i}t + \epsilon_{i}(t),$ 

$$\epsilon_{i}(t+1) = \Big(\frac{c_{1}(i-1+\delta_{in})}{t}\Big)\epsilon_{i-1}(t) + \Big(1 - \frac{c_{1}(i+\delta_{in})}{t}\Big)\epsilon_{i}(t) + o(t^{-2/5}). \end{aligned}$$

Analysis of the model(contd.)

**Proof contd**: We induct on *t*. Assume  $|\epsilon_i(t)| \le At^{3/5} \forall i \ge 0$ 

$$\begin{split} |\epsilon_{i}(t+1)| &\leq \\ \left(\frac{c_{1}(i-1+\delta_{in})}{t}\right) |\epsilon_{i-1}(t)| + \left(1 - \frac{c_{1}(i+\delta_{in})}{t}\right) |\epsilon_{i}(t)| + |o(t^{-2/5})| \\ \left(\frac{c_{1}(i-1+\delta_{in})}{t}\right) A t^{3/5} + \left(1 - \frac{c_{1}(i+\delta_{in})}{t}\right) A t^{3/5} + |o(t^{-2/5})| \\ \left(1 - \frac{c_{1}}{t}\right) A t^{3/5} + |o(t^{-2/5})| \end{split}$$

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Analysis of the model(contd.)

**Proof contd**: For large *t*, we approximate

$$(t+1)^{3/5} = t^{3/5}(1+1/t)^{3/5} \ge t^{3/5}(1+\frac{2}{5}t^{-1}).$$

$$\begin{aligned} |\epsilon_i(t+1)| &\leq \quad \left(1 - \frac{c_1}{t}\right) A t^{3/5} + |o(t^{-2/5})| \\ &\leq \quad A t^{3/5} - c_1 t^{-2/5} + |o(t^{-2/5})| \\ &\leq \quad A (t+1)^{3/5} - \frac{2}{5} t^{-2/5} - c_1 t^{-2/5} + |o(t^{-2/5})| \\ &\leq \quad A (t+1)^{3/5} \end{aligned}$$

# Azuma-Hoeffding inequality

Suppose that  $X_0, X_1, ..., X_n$  is a martingale w.r.t  $A_0, A_1, ..., A_n$ , and  $a_i \le X_{i+1} - X_i \le b_i$  i = 1, 2, ..., n - 1, then for any  $t \ge 0$ , we have

$$\Pr[|X_n - X_0| \ge t] \le 2e^{-t^2 / \sum_i (b_i - a_i)^2}$$

We use Azuma-Hoeffding inequality to show concentration of  $x_i(t)$ .

# Azuma-Hoeffding inequality

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We use Azuma-Hoeffding inequality to show concentration of  $x_i(t)$ .

# Restriction in the sequence of choices

- A-H inequality can be applied only to those sequences where the number of vertices introduced is roughly the mean.
- $\bullet\,$  Denote the set of such sequences by  ${\cal A}\,$
- **Pr**( a sequence  $C \in A$ )  $\geq 1 o(t^{-1})$

It is easy to see that

$$\mathbf{E}(x_i(t)|\mathcal{A}) = p_i t + o(t) \tag{2}$$

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# Concentration

Given  $\mathcal{A}$ ,

- # of old vertices involved in determining G(t) at most 2t
- Changing a choice from v to v' only affects the degrees of v and v'
- $x_i(t)$  changes by at most 2.

Applying A-H inequality,

$$\Pr[|x_i(t) - \mathsf{E}(x_i(t)|\mathcal{A})| \ge t^{3/4}|\mathcal{A}] \le 2e^{-2t^{3/2}/((2t)2^2)} = 2e^{-\sqrt{t}/4}.$$
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**Theorem:** Assume  $\alpha, \gamma < 1$ , and that  $\alpha \delta_{in} + \gamma \delta_{out} > 0$ . Let  $i, j \ge 0$  be fixed. Let  $n_{i,j} - \#$  of vertices with in-degree *i* and out-degree *j*. Then there is a constant  $f_{i,j}$  such that  $n_{ij}(t) = f_{ij}t + o(t)$  holds with probability 1. Furthermore, for  $j \ge 1$  fixed and  $i \to \infty$ ,

$$f_{ij}\sim C_j i^{-X_{IN}'}$$

while for  $i \ge 1$  fixed and  $j \to \infty$ ,

$$f_{ij} \sim D_i j^{-X'_{OUT}}.$$

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where the  $C_i$  and  $D_i$  are positive constants.

# Results

- In the world wide web,  $X_{IN} = 2.1$  and  $X_{OUT} = 2.7$ .
- For  $c_2 = 0.59, \alpha = 0.41, c_1 = 1/1.1$  and

$$\delta_{in} = \frac{1.1(\alpha + \beta) - 1}{1 - \beta},$$

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the model achieves the measured exponents.

# **Other Models**

#### Nodes with *fitness*:

- Every node has an in-fitness and out-fitness associated with it, denoted by λ<sub>ν</sub> and μ<sub>ν</sub>.
- $\lambda_v$  and  $\mu_v$  are drawn from distributions  $D_{IN}$  and  $D_{OUT}$  on the non-negative reals.

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