

Conditional Probability

Suppose $A \subseteq \Omega$. We define an *induced* probability \mathbf{P}_A by

$$\mathbf{P}_A(\omega) = \frac{\mathbf{P}(\omega)}{\mathbf{P}(A)} \quad \text{for } \omega \in A.$$

Two Coins

$A = \{\text{At least one coin is a Head}\}$. $\mathbf{P}(A) = 3/4$.

$B = \{\text{Both are heads}\}$. $\mathbf{P}(B) = 1/4$.
So $\mathbf{P}_A(B) = 1/3$.

Usually write $\mathbf{P}(B | A)$ for $\mathbf{P}_A(B)$.

If B is an arbitrary subset of Ω we write

$$\mathbf{P}(B | A) = \mathbf{P}_A(A \cap B) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(A)}.$$

Two Dice

$A = \{x_1 + x_2 \text{ is odd}\}$. $\mathbf{P}(A) = 1/2$

$B = \{x_1 > x_2\}$. $\mathbf{P}(B) = 5/12$.

$\mathbf{P}(A \cap B) = 1/4$.

$$\mathbf{P}(B | A) = 1/2 > \mathbf{P}(B).$$

Binomial n coin tosses. $p = \mathbf{P}(\text{Heads})$ for each toss.

$\Omega = \{H, T\}^n$.

$$\mathbf{P}(\omega) = p^k(1-p)^{n-k}$$

where k is the number of H 's in ω .

E.g. $\mathbf{P}(HHTTHTHHTHHTHT) = p^8(1-p)^6$.

Fix k . $A = \{\omega : H \text{ appears } k \text{ times}\}$

$$\mathbf{P}(A) = \binom{n}{k} p^k (1-p)^{n-k}.$$

If $\omega \in A$ then

$$\mathbf{P}_A(\omega) = \frac{p^k(1-p)^{n-k}}{\binom{n}{k} p^k (1-p)^{n-k}} = \frac{1}{\binom{n}{k}}$$

i.e. *conditional* on there being k heads, each sequence with k heads is equally likely.

Law of Total Probability

Let B_1, B_2, \dots, B_n be pairwise disjoint events which partition Ω . For any other event A ,

$$\mathbf{P}(A) = \sum_{i=1}^n \mathbf{P}(A | B_i) \mathbf{P}(B_i).$$

Proof

$$\begin{aligned} \sum_{i=1}^n \mathbf{P}(A | B_i) \mathbf{P}(B_i) &= \sum_{i=1}^n \mathbf{P}(B_i \cap A) \\ &= \mathbf{P}\left(\bigcup_{i=1}^n (B_i \cap A)\right) \quad (1) \\ &= \mathbf{P}(A). \end{aligned}$$

There is equality in (1) because the events $B_i \cap A$ are pairwise disjoint.

Suppose we choose 2 numbers X_1, X_2 randomly from $[n]$. What is the probability of $A = \{|X_1 - X_2| \leq 1\}$?

Let $B_i = \{X_1 = i\}$ for $i \in [n]$. Then

$$\begin{aligned} \mathbf{P}(A) &= \sum_{i=1}^n \mathbf{P}(A \mid B_i) \mathbf{P}(B_i) \\ &= \frac{1}{n} \sum_{i=1}^n \mathbf{P}(A \mid B_i) \\ &= \frac{1}{n} \left(\mathbf{P}(A \mid B_1) + \mathbf{P}(A \mid B_n) + \right. \\ &\quad \left. \sum_{i=2}^{n-1} \mathbf{P}(A \mid B_i) \right) \\ &= \frac{1}{n} \left(\frac{2}{n} + \frac{2}{n} + \sum_{i=2}^{n-1} \frac{3}{n} \right) \\ &= \frac{3n-2}{n^2}. \end{aligned}$$

2 sets $S, T \subseteq [n]$ are chosen (i) independently and (ii) uniformly at random from all possible sets. ($\Omega = \{0, 1\}^{2n}$). Let

$$A = \{|S| = |T| \text{ and } S \cap T = \emptyset\}.$$

For each $X \subseteq [n]$ we let $B_X = \{S = X\} = \{(X, T) : T \subseteq [n]\}$. Thus for each X , $\mathbf{P}(B_X) = 2^{-n}$. So,

$$\begin{aligned} \mathbf{P}(A) &= \sum_X \mathbf{P}(A \mid B_X) \mathbf{P}(B_X) \\ &= 2^{-n} \sum_X \binom{n-|X|}{|X|} 2^{-n} \quad (2) \\ &= 4^{-n} \sum_{k=0}^n \binom{n}{k} \binom{n-k}{k}. \end{aligned}$$

(2) follows from the fact that there are $\binom{n-|X|}{|X|}$ subsets of the same size as X which are disjoint from X .

Independence

Two events A, B are said to be *independent* if

$$\mathbf{P}(A \cap B) = \mathbf{P}(A)\mathbf{P}(B),$$

or equivalently

$$\mathbf{P}(A \mid B) = \mathbf{P}(A).$$

(i) **Two Dice**

$$A = \{\omega : x_1 \text{ is odd}\}, B = \{\omega : x_1 = x_2\}.$$
$$|A|=18, \quad |B|=6, \quad |A \cap B|=3.$$
$$\mathbf{P}(A) = 1/2, \mathbf{P}(B) = 1/6, \mathbf{P}(A \cap B) = 1/12.$$

A, B are independent.

(ii) $A = \{x_1 \geq 3\}$, $B = \{x_1 \geq x_2\}$.

 $|A|=24, |B|=21, |A \cap B|=18.$
$$\mathbf{P}(A) = 2/3, \mathbf{P}(B) = 7/12, \mathbf{P}(A \cap B) = 1/2.$$

A, B are not independent.

Random Bits

Suppose $\Omega = \{0, 1\}^n = \{(x_1, x_2, \dots, x_n) : x_j = 0/1\}$ with uniform distribution.

Suppose event A is determined by the values of $x_i, i \in \Delta_A$

e.g. if $A = \{x_1 = x_2 = \dots = x_{10} = 0\}$ then $\Delta_A = \{1, 2, \dots, 10\}$.

More Precisely: for $S \subseteq [n]$ and $x \in \Omega$ let $x_S \in \{0, 1\}^S$ be defined by $(x_S)_i = x_i, i \in S$.

Ex. $n = 10$, $S = \{2, 5, 8\}$ and

$$x = (0, 0, 1, 0, 0, 1, 1, 1, 1, 0). \quad x_S = \{0, 0, 1\}.$$

A is *determined* by Δ_A if $\exists S_A \subseteq \{0, 1\}^{\Delta_A}$ such that $x \in A$ iff $x_{\Delta_A} \in S_A$. Furthermore, no subset of Δ_A has this property.

In our example above,

$S_A = \{(0, 0, 0, 0, 0, 0, 0, 0, 0, 0)\}$ – ($|S_A| = 1$ here.)

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Claim: if events A, B are such that $\Delta_A \cap \Delta_B = \emptyset$ then A and B are independent.

$$\mathbf{P}(A) = \frac{|S_A|}{2^{|\Delta_A|}} \text{ and } \mathbf{P}(B) = \frac{|S_B|}{2^{|\Delta_B|}}.$$

$$\begin{aligned} \mathbf{P}(A \cap B) &= \frac{1}{2^n} \sum_{x \in \{0,1\}^n} 1_{\{x_{\Delta_A} \in S_A, x_{\Delta_B} \in S_B\}} \\ &= \frac{1}{2^n} |S_A| |S_B| 2^{n-|I_A|-|I_B|} \\ &= \mathbf{P}(A) \mathbf{P}(B). \end{aligned}$$

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