# On Perfect Matchings and Hamilton Cycles in Sums of Random Trees 

Alan Frieze, ${ }^{*}$ Michał Karoński† Luboš Thoma ${ }^{\ddagger}$


#### Abstract

We prove that the sum of two random trees possesses with high probability a perfect matching and the sum of five random trees possesses with high probability a Hamilton cycle.


AMS Subject Classification.(1991) 05C80

Keywords. Sums of random trees, perfect matching, hamilton cycle.

[^0]
## 1 Introduction

In this paper we prove that an appropriately defined sum of two random trees possesses with high probability a perfect matching. Secondly, we show that the sum of five random trees possesses with high probability a Hamilton cycle.

We say that a sequence of events $\mathcal{E}_{n}$ (defined on a sequence of probabilistic spaces) holds with high probability (whp in short) if the probabilities of these events converge to 1 as $n \rightarrow \infty$.

For an integer $n$, we use $[n]$ to denote the set $\{1, \ldots, n\}$. A random tree on the set $[n]$ is a tree on this set chosen uniformly at random from the family of all trees on the set $[n]$.

Definition 1 (Sums) Let $k$ be a positive integer. For trees $T_{1}, \ldots, T_{k}$, all of them on the set $[n]$, we define their sum $\mathcal{S T}\left(T_{1}, \ldots, T_{k}\right)$ as the graph on the vertex set $[n]$ and edge set being the union of edge sets of the trees $T_{1}, \ldots, T_{k}$, where the parallel edges coalesce.

Let $f$ be a mapping from $[n] \rightarrow[n]$. Let $D(f)$ be its associated functional digraph i.e. the graph with vertex set $[n]$ and edges $(i, f(i)), i \in[n]$. For a set $f_{1}, \ldots, f_{k}$ of such mappings we define their sum $\mathcal{S M}\left(f_{1}, \ldots, f_{k}\right)$ as the union of the digraphs $D\left(f_{i}\right), 1 \leq i \leq k$.

Let $k$ be a positive integer. Consider $k$ random trees $T_{1}, \ldots, T_{k}$ on $[n]$ chosen independently. We use the notation $\mathbf{S T}_{n}(k)$ for $\mathcal{S T}\left(T_{1}, \ldots, T_{k}\right)$.

A random mapping $f:[n] \rightarrow[n]$ is a mapping from the set $[n]$ to itself chosen uniformly at random from the family of all mappings $[n] \rightarrow[n]$. Similarly, as in the case for trees, we use $\mathbf{S M}_{n}(k)$ to denote the sum of $k$ random mappings.
$\mathbf{S M}_{n}(k)$ is a well studied model of random graph. Frieze [4] showed that whp $\mathbf{S M}_{n}(2)$ has a perfect matching (see also Shamir and Upfal [11] who showed that whp $\mathbf{S M}_{n}(6)$ has a perfect matching). Cooper and Frieze [2] have shown that whp $\mathbf{S M}_{n}(4)$ has a Hamilton cycle but the problem of whether
or not $\mathbf{S M}_{n}(3)$ has whp a Hamilton cycle is one of the most important open questions in the theory of Random graphs.

There is also a well known bipartite mapping model $\mathbf{S M}_{n, n}(k)$. Walkup [12] had earlier shown that $\mathbf{S M}_{n, n}(2)$ has whp a perfect matching.
$\mathbf{S T}_{n}(k)$ is less well studied. Schmutz [9] computed the expected number of perfect matchings in $\operatorname{ST}_{n}(2)$ and showed that asymptotically it is $(4 / e)^{n}$. He also studied a bipartite model $\mathbf{S T}_{n, n}(k)$ where the trees involved are random subtrees of the complete bipartite graph $K_{n, n}$ and showed that $\mathbf{S T}_{n, n}(2)$ has whp a perfect matching.

In Section 2, we prove the following theorem

## Theorem 2

(a) $\quad \lim _{\substack{n \rightarrow \infty \\ n \text { even }}} \operatorname{Prob}\left(\mathbf{S T}_{n}(1)\right.$ has a perfect matching $)=0$.
(b) $\quad \lim _{\substack{n \rightarrow \infty \\ n \text { even }}} \operatorname{Prob}\left(\mathbf{S T}_{n}(2)\right.$ has a perfect matching $)=1$.

Using the proof methodology of Frieze and Łuczak [5] who showed that $\mathrm{SM}_{n}(5)$ has whp a Hamilton cycle, we prove a result on the existence of Hamiltonian cycles in $\mathbf{S T}_{n}(5)$ in Section 3.

## Theorem 3

$$
\lim _{n \rightarrow \infty} \operatorname{Prob}\left(\mathbf{S T}_{n}(5) \text { has a Hamilton cycle }\right)=1
$$

## 2 Perfect Matchings - Proof of Theorem 2

(a) This follows immediately from Meir and Moon's result [8] that the size of the largest matching in a random tree is whp asymptotic to $(1-\rho) n \approx .432 n$ where $\rho e^{\rho}=1$.
(b) The proof of the second limit in Theorem 2 consists of several lemmas. Our starting point is a lemma by Gallai and Edmonds (Lemma 4) which
gives a sufficient condition for the existence of a perfect matching. In the view of this lemma, it is enough to show that whp there is no bad set in $\mathbf{S T}_{n}(2)$. To show that we are going to distinguish different sizes of a bad set. Lemma 5 implies that for any fixed positive integer $k_{0} \mathbf{S T}_{n}(2)$ has whp no bad set of a size at most $k_{0}$. The next range of bad sets we eliminate are bad sets of size at most $u_{0} n$ for some positive constant $u_{0}$. Using Lemma 7 we conclude that whp $\mathbf{S T}_{n}(2)$ has no such bad sets. Finally, a correspondence between labelled trees on $n$ vertices and mappings from the set $[n]$ into itself and Lemma 10 imply that whp $\mathbf{S T}_{n}(2)$ does not contain a bad set of size larger than $u_{0} n$.

Before giving the lemmas we need some notation. Let $G=(V, E)$ be a graph. For $U \subseteq V$, let $G[U]=\left(U, E_{U}\right)$ be a subgraph of $G$ induced on $U$, i.e., $E_{U}=\{e \in E$; both vertices of $e$ belong to $U\}$. Further, let $N_{G}(U)=\{v \in V \backslash U$; there is $u \in U$ such that $\{u, v\} \in E\}$ denote the neighborhood of the set $U$ and set $N(U)=N_{\mathbf{S T}_{n}(2)}(U)$. A subset $U \subseteq V$ is said to be stable if $E_{U}=\emptyset$.

The following lemma is due to Gallai [6] and Edmonds [3] (cf. [4]).

Lemma 4 If a graph $G$ does not have a perfect matching then there exists $K \subseteq V(G),|K|=k \geq 0$ such that if $H=G[V(G) \backslash K]$ then
$H$ has at least $k+1$ components with an odd number of vertices ;

No odd component of $H$, which is not an isolated vertex, is a tree .

The set $K$ guaranteed by Lemma 4 will be called a bad set.
In the following sequence of lemmas, we are going to show that for $n$ even $\mathbf{S T}_{n}(2)$ has whp no bad set.

Before starting with the lemmas, we recall the following two formulas: the number of forests on n vertices with $k$ fixed roots is equal to $k n^{n-k-1}$, and the number of forests on n vertices with $k$ roots (there roots can be any $k$ of the n vertices) is equal to $\binom{n-1}{k-1} n^{n-k}$.

Lemma 5 For sets $K, L \subseteq V_{n}$, let $\mathcal{A}_{1}(K, L)$ be the event that $N(L) \subseteq K$. For positive integers $k, l$ define the event $\mathcal{A}_{1}(k, l)$ by
there exist $K, L \subseteq V_{n}, K \cap L=\emptyset,|K|=k,|L|=l$ such that $\mathcal{A}_{1}(K, L)$ occurs.
For $0<\epsilon<1$ let $u(\epsilon)=\left[\frac{1-\epsilon}{3 e^{4}(1+\epsilon)^{1+\epsilon}}\right]^{1 / \epsilon}$ and suppose that $u=u(\epsilon)$ satisfies $\left(5 e^{4}\right)^{u} / u^{u} \leq 2^{1 / 8 e^{2}}$.

Then setting $n_{1}=\lfloor u n\rfloor$ and $l_{1}=\lceil(1+\epsilon) k\rceil$ and

$$
\mathcal{A}_{1}(\epsilon)=\bigcup_{k=1}^{n_{1}} \bigcup_{l=l_{1}}^{\lfloor n / 2\rfloor} \mathcal{A}_{1}(k, l)
$$

we have

$$
\lim _{n \rightarrow \infty} \operatorname{Prob}\left(\mathcal{A}_{1}(\epsilon)\right)=0
$$

## Proof

To bound $\operatorname{Prob}\left(\mathcal{A}_{1}(k, l)\right)$ we are going to divide the ranges of $k$ and $l$ into the following two cases:
(A) $l \leq\left\lfloor n /\left(2 e^{2}\right)\right\rfloor$ and any $k$, and
(B) $l>\left\lfloor n /\left(2 e^{2}\right)\right\rfloor$ and any $k$.

Fix $K, L$ and the lowest numbered vertex $v \in K$. Now, each tree $T$ with $N_{T}(L) \subseteq K$ is considered to be oriented towards $v$.

Case A. Let $T$ be a tree oriented as described above. Delete edges oriented out of vertices in $L$. This leaves a forest $F^{\prime}$ with $l+1$ roots and $n$ vertices. There are at most $(l+1) n^{n-l-2}$ such forests, (not all forests with $l+1$ roots and $n$ vertices respect $\left.N_{T}(L) \subseteq K\right)$. To obtain $T$ we construct a forest $F^{\prime \prime}$ with vertex set $K \cup L$ and roots $K$ and take $T=F^{\prime} \cup F^{\prime \prime}$. We can construct
$F^{\prime \prime}$ in $k(k+l)^{l-1}$ ways. Hence,

$$
\begin{align*}
\operatorname{Prob}\left(\mathcal{A}_{1}(k, l)\right) & \leq\binom{ n}{k}\binom{n}{l}\left(\frac{(l+1) n^{n-l-2} k(k+l)^{l-1}}{n^{n-2}}\right)^{2}  \tag{3}\\
& \leq \frac{(n e)^{k+l}}{k^{k} l^{l}} k^{2} \frac{l^{2 l}\left(1+\frac{k}{l}\right)^{2 l}}{n^{2 l}} \\
& \leq \frac{n^{k} k^{2} e^{3 k}}{k^{k}}\left(\frac{e l}{n}\right)^{l}
\end{align*}
$$

Putting $\mu_{l}=(e l / n)^{l}$, we get $\mu_{l} / \mu_{l-1}<1 / 2$ for $l<n / 2 e^{2}$.
Thus,

$$
\sum_{l=l_{1}}^{\left\lfloor n / 2 e^{2}\right\rfloor}\left(\frac{e l}{n}\right)^{l} \leq 2\left(\frac{e l_{1}}{n}\right)^{l_{1}}
$$

Hence,

$$
\begin{aligned}
\sum_{k=1}^{n_{1}} \sum_{l=l_{1}}^{\left\lfloor n / 2 e^{2}\right\rfloor} \operatorname{Prob}\left(\mathcal{A}_{1}(k, l)\right) & \leq \sum_{k=1}^{\lfloor u n\rfloor} \frac{n^{k} k^{2} e^{3 k}}{k^{k}} 2\left(\frac{e l_{1}}{n}\right)^{l_{1}} \\
& \leq 2 \sum_{k=1}^{\lfloor u n\rfloor}\left(\frac{3 e^{4}(1+\epsilon)^{1+\epsilon} k^{\epsilon}}{n^{\epsilon}}\right)^{k} \\
& =o(1) .
\end{aligned}
$$

Case B. Let $T$ be a tree oriented as described above. Let $F^{\prime}$ be the forest obtained by deleting edges oriented out of $K$ and deleting vertices in $L$. This forms a forest with $n-l$ vertices and $k$ roots $K$. There are $k(n-l)^{n-l-k-1}$ such forests and each forest can be extended in at most $k(k+l)^{l-1} n^{k-1}$ ways to form the oriented tree $T$. Indeed, we attach the vertices from $L$ by constructing a forest on $K \cup L$ with roots $K$ in at most $k(k+l)^{l-1}$ ways. The remaining $k-1$ edges oriented out of $K$ can be chosen in at most $n^{k-1}$ ways.

Hence,

$$
\begin{aligned}
\operatorname{Prob}\left(\mathcal{A}_{1}(k, l)\right) & \leq\binom{ n}{k}\binom{n}{l}\left[\frac{k(n-l)^{n-l-k-1} k(k+l)^{l-1} n^{k}}{n^{n-2}}\right]^{2} \\
& \leq \frac{(n e)^{k+l}}{k^{k} l^{l}} \frac{k^{4} e^{-2 l+\frac{2 l(l+k+1)}{n} l^{2(l-1)} e^{2 k}}}{n^{2(l-1)}} \\
& \leq \frac{n^{k} e^{3 k-l} k^{4} e^{l+k+1} l^{l-2}}{k^{k} n^{l-2}} \\
& =e\left(\frac{n k^{4 / k} e^{4}}{k}\right)^{k}\left(\frac{l}{n}\right)^{l-2}
\end{aligned}
$$

For $n$ large enough,

$$
\sum_{l=\left\lfloor n / 2 e^{2}\right\rfloor+1}^{\lfloor n / 2\rfloor}\left(\frac{l}{n}\right)^{l-2} \leq n\left(\frac{1}{2}\right)^{\frac{n}{2 e^{2}}-2} \leq\left(\frac{1}{2}\right)^{\frac{n}{4 e^{2}}}
$$

Hence,

$$
\begin{aligned}
\sum_{k=1}^{n_{1}} \sum_{l=\left\lfloor n / 2 e^{2}\right\rfloor+1}^{\lfloor n / 2\rfloor} \operatorname{Prob}\left(\mathcal{A}_{1}(k, l)\right) & \leq e\left(\frac{1}{2}\right)^{\frac{n}{4 e^{2}}} \sum_{k=1}^{\lfloor u n\rfloor}\left(\frac{n k^{4 / k} e^{4}}{k}\right)^{k} \\
& \leq e\left(\frac{1}{2}\right)^{\frac{n}{4 e^{2}}} \sum_{k=1}^{\lfloor u n\rfloor}\left(\frac{5 e^{4} n}{k}\right)^{k} \\
& \leq e n\left(\frac{\left(5 e^{4}\right)^{u}}{2^{1 / 4 e^{2}} u^{u}}\right)^{n} \\
& =o(1)
\end{aligned}
$$

Lemma 6 Let $\epsilon$ be as in Lemma 5. Suppose a graph $G$ contains a bad set $K, 1 \leq k=|K| \leq u(\epsilon) n$, and no subset of $K$ is bad. Let $H=G\left[V_{n} \backslash K\right]$ have $s \geq k+1$ odd components $C_{1}, C_{2}, \ldots, C_{s}$ with $n_{1}=n_{2}=\cdots=n_{p}=1<3 \leq$ $n_{p+1} \leq \cdots \leq n_{s}$ vertices, respectively.

Assume that $\mathcal{A}_{1}(\epsilon)$ does not occur. Then there exists a partition $K, P, Q, R$ of $V_{n}$ with $p=|P|, q=|Q|$ satisfying

$$
\begin{equation*}
N(R) \subseteq K, N(P) \subseteq K, N(Q) \subseteq K \tag{4}
\end{equation*}
$$

$$
\begin{equation*}
P \text { is a stable set ; } \tag{5}
\end{equation*}
$$

Each vertex of $K$ is adjacent to at least one member of $P \cup Q$;
$1 \leq k \leq u(\epsilon) n, 0 \leq p+q<(1+\epsilon) k, p+\lfloor q / 3\rfloor \geq k$ and $q=0$ implies $p \geq k+1$.

Proof See [4].
Let $\mathcal{A}_{2}(\epsilon)$ be the event that there is a partition satisfying (4) - (7) described in Lemma 6.

We can immediately show that for any fixed integer $k_{0}$

$$
\lim _{\substack{n \rightarrow \infty \\ n \text { even }}} \operatorname{Prob}\left(\mathbf{S T}_{n}(2) \text { has a bad set } K \text {, with } 1 \leq|K| \leq k_{0}\right)=0
$$

Let us take $\epsilon=1 / 2 k_{0}$ and assume that $\mathcal{A}_{1}(\epsilon)$ does not occur. If there is a bad set $K$ with $1 \leq|K| \leq k_{0}$ then the conditions of the Lemma 6 are satisfied for some $k \leq k_{0}$. But (7) implies $q<3 \epsilon k / 2$ which in this case forces $q<1$ or $q=0$. But then $p \geq k+1$ contradicts $p<(1+\epsilon) k$.

In the proof of the following lemma we assume that $k \geq k_{0}$ for some suitably large $k_{0}$.

Lemma 7 For small $\epsilon$

$$
\lim _{n \rightarrow \infty} \operatorname{Prob}\left(\mathcal{A}_{2}(\epsilon)\right)=0
$$

## Proof

Fix $K, P, Q$ and $v \in K$. Each tree satisfying (4) - (6) can be chosen in at most $k(n-p-q)^{n-p-q-k-1} n^{k-1} k^{p}(k+q)^{q}$ ways. We first build a forest on $V \backslash(P \cup Q)$ with roots in $K\left(k(n-p-q)^{n-p-q-k-1}\right.$ ways). Then each $x \in P$ is allowed to choose in $K$, each $y \in Q$ is allowed to choose in $K \cup Q$ and each $z \in K \backslash\{v\}$ is allowed to choose in $V_{n}$.

Let $K_{i}$ be the set of vertices in $K$ which have a neighbour in $P \cup Q$ in the tree $T_{i}, i=1,2$. There are two possibilities:

A: $\left|K_{1}\right| \geq .9 k$.
Of the $k^{p}(k+q)^{q}$ choices ascribed to vertices in $P \cup Q$, at most a proportion $.9^{k}$ will make $\left|K_{1}\right| \geq .9 k$. Indeed, for each $x \in K$ the probability it is included in such a choice is at most

$$
1-\left(1-\frac{1}{k}\right)^{p}\left(1-\frac{1}{k+q}\right)^{q} \leq .64
$$

for large enough $k$. The corresponding events for each $x$ are clearly negatively correlated - note that we do not claim this for the choice of tree $T_{1}$, but only for the choices defined by the upper estimate. Thus,

$$
\begin{aligned}
\operatorname{Prob}\left(\left|K_{1}\right| \geq .9 k\right) & \leq\binom{ k}{.9 k}(.64)^{k} \\
& \leq(.9)^{k}
\end{aligned}
$$

B: $\left|K_{1}\right|<.9 k$. By a similar argument,

$$
\text { Prob }\left(K_{2} \supseteq K \backslash K_{1}\left|K_{1},\left|K_{1}\right|<.9 k\right) \leq(.64)^{.1 k}\right.
$$

Combining the two cases we see that for $\delta=(.64)^{1}$ we have

$$
\begin{aligned}
\operatorname{Prob}\left(\mathcal{A}_{2}(k, p, q)\right) & \leq 2\binom{n}{k, p, q}\left[\frac{k(n-p-q)^{n-p-q-k-1} n^{k-1} k^{p}(k+q)^{q}}{n^{n-2}}\right]^{2} \delta^{k} \\
& \leq 2 \frac{(n e)^{k+p+q}}{k^{k} p^{p} q^{q}} \frac{k^{2 p+2 q+2} e^{2 q^{2} / k} e^{-\frac{p+q}{n}(n-p-q-k-1)}}{n^{2 p+2 q}} \delta^{k}
\end{aligned}
$$

where $\mathcal{A}_{2}(k, p, q)$ is the event that there is a partition satisfying (4) - (7) in Lemma 6 for given $k, p, q$. We obtain for the probability of the event $\mathcal{A}_{2}(k, p, q)$, under the condition $q \geq 2$

$$
\begin{aligned}
\operatorname{Prob}\left(\mathcal{A}_{2}(k, p, q)\right) & \leq 2 \frac{e^{k-p-q} k^{2 p+2 q+2}}{k^{k} q^{q} p^{p} n^{p+q-k}} e^{2 q^{2} / k} e^{\frac{2(1+\epsilon)(2+\epsilon) k^{2}+2(1+\epsilon) k}{n}} \delta^{k} \\
& \leq 2\left(\frac{k}{e n}\right)^{p+q-k}\left(\frac{k}{p}\right)^{p}\left(\frac{k}{q}\right)^{q} k^{2} e^{2 q^{2} / k} e^{\frac{2(1+\epsilon)(2+\epsilon) k^{2}+2(1+\epsilon) k}{n}} \delta^{k}
\end{aligned}
$$

We continue with the bound on $\operatorname{Prob}\left(\mathcal{A}_{2}(k, p, q)\right)$ using $\frac{k}{p} \leq 1+\frac{q}{p}, q \leq \frac{3}{2} \epsilon k$, and $p+q-k \geq 1$. Further, we use that the function $x^{-x}$ on the interval $(0, \infty)$ has its maxima at $x=1 / e$. Thus, for every $k \geq k_{0}=k_{0}(\epsilon)$

$$
\begin{aligned}
\operatorname{Prob}\left(\mathcal{A}_{2}(k, p, q)\right) & \leq 2\left(\frac{k}{e n}\right)^{p+q-k}\left(\frac{k}{p}\right)^{p}\left(\frac{k}{q}\right)^{q} k^{2} e^{9 \epsilon^{2} k / 2} e^{12 u(\epsilon) k} e^{4 u(\epsilon)} \delta^{k} \\
& \leq \frac{2 k^{3}}{e n} e^{4 u(\epsilon)}\left(\frac{2}{3 \epsilon}\right)^{3 \epsilon k / 2} e^{5 \epsilon k+12 u(\epsilon) k} \delta^{k}
\end{aligned}
$$

Choose $\epsilon$ small enough such that $(2 / 3 \epsilon)^{3 \epsilon / 2} e^{5 \epsilon+12 u(\epsilon)} \delta \leq \mu<1$ for some $0<\mu<1$. For $k \geq k_{0}$ let $\mathcal{S}(k)=\{(p, q) ; k, p, q$ satisfy conditions (7) $\}$. Note that $|\mathcal{S}(k)| \leq 2 k^{2}$ for $\epsilon$ small. We sum up over $k, p$, and $q$. Thus,

$$
\begin{aligned}
\operatorname{Prob}\left(\mathcal{A}_{2}(\epsilon)\right) & =\sum_{k=k_{0}}^{\lfloor u(\epsilon) n\rfloor} \sum_{(p, q) \in \mathcal{S}(k)} \operatorname{Prob}\left(\mathcal{A}_{2}(k, p, q)\right) \\
& \leq \frac{4 e^{3}}{n} \sum_{k=k_{0}}^{\lfloor u(\epsilon) n\rfloor} k^{5} \mu^{k} \\
& =o(1) .
\end{aligned}
$$

Summing up: choosing $\epsilon$ small enough and $k_{0}$ sufficiently large, so far, we have proved that there is a constant $u_{0}>0$ such that

$$
\lim _{\substack{n \rightarrow \infty \\ n \text { even }}} \operatorname{Prob}\left(\mathbf{S T}_{n}(2) \text { has a bad set } K \text {, with } 1 \leq|K| \leq u_{0} n\right)=0
$$

To complete the proof of Theorem 2, we need to take care about large bad sets.

Lemma 8 Let $\mathcal{A}_{3}$ denote the following event:
$\mathbf{S T}_{n}(2)$ contains at least $(\log n)^{3}$ sets $S \subseteq V_{n}$ satisfying

$$
\begin{equation*}
|S| \leq \log \log n \tag{8}
\end{equation*}
$$

$$
\begin{equation*}
\left|E_{S}\right| \geq|S| \tag{9}
\end{equation*}
$$

Then $\lim _{n \rightarrow \infty} \operatorname{Prob}\left(\mathcal{A}_{3}\right)=0$.

Proof $\quad$ Fix $k \geq 2$. Let $X_{k}$ be a random variable counting sets $S$ with $|S|=k$ and $\left|E_{S}\right| \geq k$. Then

$$
\begin{aligned}
E X_{k} & \leq\binom{ n}{k} \frac{\sum_{t=2}^{k} \sum_{\substack{i, j \geq 1 \\
i+j=t}}(n-k)^{t}\left(\binom{k-1}{i-1} k^{k-i}\right)\left(\binom{k-1}{j-1} k^{k-j}\right)}{\left[(n-k) n^{n-(n-k)-1}\right]^{2}} \\
& =\binom{n}{k} \frac{k^{2 k} \sum_{t=2}^{k}\binom{2(k-1)}{t-2}\left(\frac{n-k}{k}\right)^{t}}{\left[(n-k) n^{k-1}\right]^{2}} \\
& \leq e^{k}\left(\frac{n}{n-k}\right)^{2}\left(\frac{k}{n}\right)^{k} \sum_{t=2}^{k}\binom{2(k-1)}{t-2}\left(\frac{n-k}{k}\right)^{t} .
\end{aligned}
$$

As we have $k \leq \log \log n$, we get

$$
\begin{aligned}
E X_{k} & \leq e^{k}\left(\frac{n}{n-k}\right)^{2} k\binom{2 k}{k}\left(\frac{n-k}{k}\right)^{k}\left(\frac{k}{n}\right)^{k} \\
& \leq(4 e)^{k} k\left(\frac{n-k}{n}\right)^{k-2} \\
& \leq(4 e)^{k} \log \log n
\end{aligned}
$$

By the Markov inequality

$$
\begin{aligned}
\operatorname{Prob}\left(\mathcal{A}_{3}\right) & =\operatorname{Prob}\left(\sum_{k=3}^{\lfloor\log \log n\rfloor} X_{k} \geq(\log n)^{3}\right) \\
& \leq \frac{2 \log \log n(4 e)^{\log \log n}}{(\log n)^{3}} \\
& =o(1)
\end{aligned}
$$

Now we shall make a use of a known one-to-one correspondence between the family of labelled trees on $n$ vertices with two marked vertices and the
family of functional digraphs $D(f)$ of mappings $f:[n] \rightarrow[n]$. Each such digraph $D$ consists of vertices $S(f)$ which form cycles and the remaining vertices form a set of trees $\mathcal{T}$ which are attached to the cycles. To obtain a tree $T$, with two appropriately marked vertices from $D$ we shall consider vertices lying on the cycles as a permutation drawn in cyclic form. Next we write such a permutation in a line form, which in turn we treat as a directed path $P$. As a final step, we re-attach the trees in $\mathcal{T}$ to their vertices on $P$ to obtain a tree with two marked vertices (these two vertices are simply the beginning and the end of $P$ ). One can easily reproduce the correspondence from trees to mappings reversing the procedure described above. We believe that the one-to-one correspondence stated above is due to Joyal. A complete description of this correspondence can be found for example in Bender and Williamson [1].

This defines a natural measure preserving mapping $\phi$ from the space of random mappings to the space o random trees ( $\phi$ just "forgets" the random choice of pair of marked vertices). To finish the proof of Theorem 2 we will use $\phi$ to construct $\mathbf{S T}_{n}(2)$ in the following way: we first generate $\mathbf{S M}_{n}(2)$ from random functions $f_{1}, f_{2}$ and then apply $\phi$ to both of them.

Definition 9 Let a pair of sets $K, P \subseteq V_{n}$ be matched if
(८) $P$ is stable in $\mathbf{S T}_{n}(2)$;
(u) $N(P)=K$;
(८८) $|P| \geq|K|-\delta(n)$,
where $\delta(n)=\left\lceil\frac{n}{\log \log n}+(\log n)^{3}\right\rceil$.

Lemma 10 Suppose $\mathbf{S T}_{n}(2)$ has no bad sets of size $u_{0} n$ or less but $\mathbf{S T}_{n}(2)$ contains a bad set $K_{0}, k=\left|K_{0}\right|>u_{0} n$. Suppose $K_{0}$ does not strictly contain another bad set and $\mathcal{A}_{3}$ does not occur in $\mathbf{S T}_{n}(2)$. Let $S=S\left(f_{1}\right) \cup S\left(f_{2}\right)$. If $s=|S|$ then either $\mathbf{S M}_{n}(2)$ contains a matched pair $K, P$ with

$$
|P|+\delta(n)+s \geq k \geq|K| \geq|P|
$$

or
$K_{0}$ contains a bad set of $\mathbf{S M}_{n}(2)$.

Proof Arguing as in Lemma 2.7 of [4] we see that $\mathbf{S T}_{n}(2)$ contains a matched pair $K_{1}, P_{1}$ with

$$
\left|P_{1}\right|+\delta(n) \geq k \geq\left|K_{1}\right| \geq\left|P_{1}\right| .
$$

Let $P=P_{1} \backslash S$. Then $P$ is stable and $|P| \geq\left|P_{1}\right|-s$. Also, $N_{\text {SM }_{n}(2)}(P) \subseteq K_{1}$. Now take $K=N_{\mathbf{S M}_{n}(2)}(P)$. Either $|K|<|P|$ and $K$ is a bad set of $\mathbf{S M}_{n}(2)$ or $|K| \geq|P|$ and $K, P$ is the required matched pair.

Both possibilities in Lemma 10 are shown not to happen whp in [4], completing the proof of Theorem 2. (We observe first that whp $s=O(\sqrt{n})$ cf. Kolchin [7]. The definition of matched pair in [4] has to be amended to $\delta(n)+O(\sqrt{n})$, but this does not affect the proof there given in any significant way.)

## 3 Hamilton Cycles - Proof of Theorem 3

Frieze and Łuczak [5] proved that whp there is a Hamilton cycle in $\mathbf{S M}_{n}(5)$. We will use the same proof technique here, giving only a sketch as the main ideas are very similar.

We consider $\mathbf{S T}_{n}(5)$ to be the union of $\mathbf{S T}_{n}(4)$ and a random tree $T_{5}$. We observe first that Theorem 2 shows that whp $\mathbf{S T}_{n}(4)$ contains the union of two perfect matchings $M_{1}, M_{2}$. We can argue (see Lemma 2 of [5]) that $M_{1}$ and $M_{2}$ are an independent pair of matchings, chosen uniformly from the set of all possible perfect matchings. Furthermore, (see Lemma 3 of [5]) $M_{1} \cup M_{2}$ is whp the union of at most $3 \log n$ vertex disjoint cycles - some cycles may possibly just be double edges.

We show next that whp $\mathbf{S T}_{n}(4)$ has good expansion properties. For sets $K, L \subseteq V_{n}$, let $\tilde{\mathcal{A}}_{1}(K, L)$ be the event that $N_{\mathbf{S T}_{n}(4)}(L) \subseteq K$ and let

$$
\mathcal{A}_{4}=\bigcup_{\substack{|K| \leq 10-3_{n} \\|L|=2|K|}} \tilde{\mathcal{A}}_{1}(K, L) .
$$

Lemma $11 \operatorname{Prob}\left(\mathcal{A}_{4}\right)=o(1)$.

Proof It follows from (3) that

$$
\begin{aligned}
\operatorname{Prob}\left(\mathcal{A}_{4}\right) & \leq \sum_{k=1}^{10^{-3} n}\binom{n}{k}\binom{n}{2 k}\left(\frac{(k+1) n^{n-k-2} 2 k(3 k)^{k-1}}{n^{n-2}}\right)^{4} \\
& \leq \sum_{k=1}^{10^{-3} n}(k+1)^{4}\left(\frac{81 e^{3} k}{4 n}\right)^{k} \\
& =o(1) .
\end{aligned}
$$

The idea now is to use the extension-rotation procedure (as described in [5]). The main idea that we get from [5] is to reserve the edges of $T_{5}$ for closing paths. More precisely, at some points of our extension-rotation procedure we will have a set $A,|A| \geq 10^{-3} n$ and for each $a \in A$ there is a collection of paths with endpoints $B(a),|B(a)| \geq 10^{-3} n$ and we succeed if we always find a $T_{5}$-edge of the form $(a, b)$ where $b \in B(a)$. With high probability we only need to attempt this at most $3 \log n$ times (from Lemma 11). Let us suppose that the edges of $T_{5}$ come from a random mapping $f_{5}$ where an adversary has altered the edges coming out of a set $S$ of $O(\sqrt{n})$ nodes. When given $A,\{B(a): a \in A\}$ we choose the lowest numbered $a \in A \backslash S$ whose $f_{5}$ value has not been examined. So, whp we examine a further $O(\log n) a$ 's before finding one with $f_{5}(a) \in B(a)$. Thus, whp the number of edges examined and altered throughtout the procedure is $O(\sqrt{n})$ and we succeed in finding a Hamilton cycle.

## References

[1] E.A. Bender, S.G. Williamson, Foundations of Applied Combinatorics, Addison - Wesley Pub. Co., 1991.
[2] C. Cooper and A. M. Frieze, Hamilton cycles in random graphs and directed graphs, to appear.
[3] J. Edmonds, Paths, Trees and Flowers, Canad. J. Math. 17 (1965), pp. 449-467.
[4] A. M. Frieze, Maximum Matchings in a Class of Random Graphs, J. Comb. Theory B , Vol. 40, No. 2 (1986), pp. 196-212.
[5] A. M. Frieze, T. Łuczak, Hamiltonian Cycles in a Class of Random Graphs: One Step Further, in Random Graphs '87, eds. M. Karonski, J. Jaworski, A. Rucinski, 1990, pp. 53-59.
[6] T. Gallai, Über extreme Punkt- und Kantenmengen, Ann. Univ. Sci. Budapest, Eötvös Sect. Math. 2 (1959), pp. 133-138.
[7] V. F. Kolchin, Random Mappings, Optimization Software Inc., New York 1986.
[8] A. Meir and J. W. Moon, The expected node-independence number of random trees, Nederl. Akad. Wetensch. Proc. Ser. Indag. Math. 35 (1974), pp. 335-341.
[9] E. Schmutz, Private Communication.
[10] E. Schmutz, Matchings in superpositions of $(n, n)$-bipartite trees, Random Structures and Algorithms 5 (1994), pp. 235-241.
[11] E. Shamir and E. Upfal, One-factor in random graphs based on vertex choice, Discrete Mathematics 41 (1982), pp. 281-286.
[12] D. W. Walkup, Matchings in random regular bipartite digraphs, Discrete Mathematics 31 (1980), pp. 59-64.


[^0]:    *Department of Mathematical Sciences, Carnegie Mellon University, Pittsburgh PA15213, af1p@andrew.cmu.edu; Partial support from NSF grant 953074.
    ${ }^{\dagger}$ Department of Mathematics and Computer Science, Emory University, Atlanta, Georgia 30322, and Adam Mickiewicz University, Poznań, Poland, michal@mathcs.emory.edu, karonski@math.amu.edu.pl; Partial support from NSF grant INT9406971 and KBN grant No. 2 P03A 02309.
    ${ }^{\ddagger}$ DIMACS, Rutgers University, P.O. Box 1179, Piscataway, NJ 08855; thoma@dimacs.rutgers.edu. The third author gratefully acknowledges support as a DIMACS Postdoctoral Fellow. DIMACS is a cooperative project of Rutgers University, Princeton University, AT\&T Research, Bellcore, and Bell Laboratories. DIMACS is an NSF Science and Technology Center, funded under contract STC-91-19999; and also receives support from the New Jersey Commission on Science and Technology.
    The third author was also partially supported by the NSF grant INT-9406971.

