

A MULTIDIMENSIONAL ANALOGUE OF THE RADEMACHER-GAUSSIAN TAIL COMPARISON

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ABSTRACT. We prove a dimension-free tail comparison between the Euclidean norms of sums of independent random vectors uniformly distributed in centred Euclidean spheres and properly rescaled standard Gaussian random vectors.

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Introduction. Tail comparison bounds, such as Hoeffding's inequality, have always played a crucial role in probability theory. When specified to concrete examples, very precise estimates for tail probabilities are usually known. For instance, if $\epsilon_1, \epsilon_2, \dots$ are independent random variables each taking values ± 1 with probability $\frac{1}{2}$ and g_1, g_2, \dots are independent standard Gaussian random variables, then for every $m \geq 1$, real numbers a_1, \dots, a_m and positive t ,

$$(P) \quad \mathbb{P}(|a_1\epsilon_1 + \dots + a_m\epsilon_m| > t) \leq c \cdot \mathbb{P}(|a_1g_1 + \dots + a_mg_m| > t)$$

for some absolute constant c . This inequality was first proved by Pinelis in [5] with $c \approx 4.46$. Talagrand in [6] treated the case of independent (but not necessarily identically distributed) bounded random variables by means of the Laplace transform establishing similar Gaussian tail bounds. Bobkov, Götze and Houdré obtained a bigger constant $c \approx 12.01$ in (P), but their inductive argument was much simpler (see [1]). Only very recently the best constant (equal approximately to 3.18) has been found (see [2]).

Oleszkiewicz conjectured the following multidimensional generalisation of Pinelis' Rademacher-Gaussian tail comparison (P): fix $d \geq 1$, let ξ_1, ξ_2, \dots be independent random vectors uniformly distributed in the Euclidean unit sphere $S^{d-1} \subset \mathbb{R}^d$ and let G_1, G_2, \dots be independent standard Gaussian random vectors in \mathbb{R}^d with mean zero and identity covariance matrix; there exists a universal constant C such that

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for every $m \geq 1$, real numbers a_1, \dots, a_m and $t > 0$ we have

$$(KO) \quad \mathbb{P} \left(\left\| \sum_{i=1}^m a_i \xi_i \right\| > t \right) \leq C \cdot \mathbb{P} \left(\left\| \sum_{i=1}^m a_i \frac{G_i}{\sqrt{d}} \right\| > t \right).$$

Here and throughout, $\|\cdot\|$ denotes the standard Euclidean norm in \mathbb{R}^d .

Note that the normalisation is chosen so that the vectors ξ_1 and G_1/\sqrt{d} have the same covariance matrix. Plainly, when $d = 1$, (KO) reduces to (P). For general d , it is possible to deduce (KO) with $C = O(\sqrt{d})$ from Theorem 2 in [4].

The goal of this note is to positively resolve Oleszkiewicz's conjecture. We shall show the following two theorems which are our main results. The latter will easily follow from the former.

Theorem 1. *For each $d \geq 2$, let C_d be the best constant in inequality (KO). Then, $C_d \leq e^2 < 7.39$ for $d \geq 2$ and $\limsup_{d \rightarrow \infty} C_d \leq 2$.*

Theorem 2. *Let X_1, X_2, \dots be independent rotationally invariant random vectors having values in the unit Euclidean ball in \mathbb{R}^d . Let G_1, G_2, \dots be independent standard Gaussian random vectors in \mathbb{R}^d with mean zero and identity covariance matrix. Then for every $m \geq 1$, real numbers a_1, \dots, a_m and $t > 0$ we have*

$$(1) \quad \mathbb{P} \left(\left\| \sum_{i=1}^m a_i X_i \right\| > t \right) \leq e^2 \cdot \mathbb{P} \left(\left\| \sum_{i=1}^m a_i \frac{G_i}{\sqrt{d}} \right\| > t \right),$$

where $\|\cdot\|$ stands for the standard Euclidean norm in \mathbb{R}^d .

Remark. This will no longer hold if we only assume the boundedness of the X_i . For example, consider independent X_i taking only two values $(\pm 1, 0, \dots, 0)$ each with probability $\frac{1}{2}$. Then for, say $a_1 = \dots = a_m = 1/\sqrt{m}$, $t = 2$, the right-hand side of (1) goes to zero when d goes to infinity, whereas the left-hand side does not depend on d .

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Proofs. Our proof of Theorem 1 is inductive, inspired by the inductive approach to the one dimensional case from [1]. In the inductive step, using the spherical

symmetry of our problem, we arrive at an inequality comparing the Gaussian volume between centred and shifted balls (Lemma 3 below). This inequality can be viewed as a multidimensional generalisation of the two point-inequality derived in the inductive step in [1]. Its proof leads us to somewhat subtle estimates for the Laplace transform of the first coordinate of ξ_1 (Lemma 2 below).

We shall need four lemmas. We start with a result which will be used to provide numerical values of our constants (see also Lemma 2 in [3]).

Lemma 1. *Let $d \geq 2$ and G be a standard Gaussian random vector in \mathbb{R}^d . Then*

$$\mathbb{P}\left(\|G\| > \sqrt{d+2}\right) \geq 1/e^2.$$

Proof (by I. Pinelis). Set $p_d = \mathbb{P}\left(\|G\| > \sqrt{d+2}\right)$. By inequality (2.6) in [5]

$$\mathbb{P}\left(\|G\| > \sqrt{d+2}\right) \geq \mathbb{P}\left(g > (\sqrt{d+2} - \sqrt{d-1})\sqrt{2}\right),$$

where g is a standard Gaussian random variable (with mean zero and variance one).

Since the right-hand side is increasing in d , we have for $d \geq 4$

$$p_d \geq \mathbb{P}\left(g > \sqrt{12} - \sqrt{6}\right) > 1/e^2.$$

Moreover, we directly check that $p_2 = 1/e^2 < p_3$. Thus, $p_d \geq 1/e^2$ for $d \geq 2$. \square

The next lemma gives tight estimates for the Laplace transform of the first coordinate of a random vector uniformly distributed in the unit sphere. We hope these estimates are of independent interest, in addition to playing a major role in our proof.

Lemma 2. *For $d \geq -1$ and $b \geq 0$ let us denote*

$$J_d = J_d(b) = \int_{-1}^1 (1-x^2)^{d/2} e^{bx} dx.$$

Then for every $d \geq 2$ we have

- (a) $b^2 J_{d+1} = -d(d+1)J_{d-1} + (d+1)(d-1)J_{d-3}$,
- (b) $\frac{J_{d-3}}{J_{d-1}} \geq \frac{d}{d-1} \left(\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}} \right)$,
- (c) $\frac{J_{d-1}}{J_{d+1}} \leq \frac{d+2}{d+1} \left(\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}} \right)$,
- (d) $J_{d+1}J_{d-3} \geq J_{d-1}^2 \frac{d(d+1)}{(d-1)(d+2)}$.

Proof. (a) Integrating by parts twice we get $b^2 J_{d+1} = \int_{-1}^1 \left(\sqrt{1-x^2}^{d+1}\right)'' e^{bx} dx$, which, after computing the second derivative in the above expression, easily leads to the desired relation.

For the proof of (b) and (c) let us first observe that due to (a) these two assertions hold true for $b = 0$. We then show that for $b > 0$, (b) and (c) are equivalent. Indeed, part (a) yields

$$(2) \quad \frac{J_{d-3}}{J_{d-1}} = \frac{d}{d-1} + \frac{b^2}{(d+1)(d-1)} \frac{J_{d+1}}{J_{d-1}}, \quad d \geq 2.$$

Thus, (b) is equivalent to

$$\frac{b^2}{(d+1)(d-1)} \frac{J_{d+1}}{J_{d-1}} \geq \frac{d}{d-1} \left(-\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}} \right) = \frac{d}{d-1} \frac{\frac{b^2}{d(d+2)}}{\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}}}.$$

After cancelling common factors on both sides this becomes (c).

Let us fix $b > 0$. We shall show (b) by backwards induction on d . We can use (2) for $d+2$, that is the equality

$$\frac{J_{d-1}}{J_{d+1}} = \frac{d+2}{d+1} + \frac{b^2}{(d+3)(d+1)} \frac{J_{d+3}}{J_{d+1}},$$

to rewrite (c) in the form

$$\frac{b^2}{(d+3)(d+1)} \frac{J_{d+3}}{J_{d+1}} \leq \frac{d+2}{d+1} \left(-\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}} \right) = \frac{d+2}{d+1} \frac{\frac{b^2}{d(d+2)}}{\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}}},$$

which becomes

$$(3) \quad \frac{J_{d+1}}{J_{d+3}} \geq \frac{d}{d+3} \left(\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}} \right).$$

First notice that (b) and equivalently (3) hold for all $d \geq d_0(b)$ for some large enough $d_0(b)$ which depends only on b . To see this observe that the left-hand side of (3) is strictly greater than 1, whereas the right-hand side for large d is of order $1 - 3/d + o(1/d)$. Now suppose (b) holds for $d+4$, that is

$$\frac{J_{d+1}}{J_{d+3}} \geq \frac{d+4}{d+3} \left(\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{(d+4)(d+6)}} \right)$$

and we want to show (b) (induction step). By the above and the fact that (3) and (b) are equivalent, it is enough to show that

$$\frac{d+4}{d+3} \left(\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{(d+4)(d+6)}} \right) \geq \frac{d}{d+3} \left(\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}} \right).$$

This follows from $\frac{d+4}{d+6} \geq \frac{d}{d+2}$ and the estimate

$$\frac{d+4}{2} + \sqrt{\frac{(d+4)^2}{4} + b^2 \frac{d+4}{d+6}} \geq \frac{d}{2} + \sqrt{\frac{d^2}{4} + b^2 \frac{d}{d+2}}.$$

Clearly (d) immediately follows from (b) and (c). \square

Remark. Part (d) improves on Hölder's inequality which gives $J_{d-1}^2 \leq J_{d+1}J_{d-3}$.

Remark. Let us define for $d \geq -1$ and $b \geq 0$ the normalised integrals $\bar{J}_d(b) = J_d(b)/J_d(0)$ so that they are the Laplace transforms of the probability densities: if $d \geq 2$ and ξ is a random vector in \mathbb{R}^d uniformly distributed in the Euclidean unit sphere S^{d-1} , we check that (by rotational invariance)

$$\bar{J}_{d-3}(b) = \mathbb{E}e^{\langle v, \xi \rangle},$$

for any vector $v \in \mathbb{R}^d$ of length b . Part (a) for $b = 0$ gives $J_{d-3}(0)/J_{d-1}(0) = d/(d-1)$. This allows to simplify (b),(c),(d) rewritten in terms of \bar{J}_d to get for $d \geq 2$

$$\begin{aligned} \text{(a')} \quad & \frac{b^2}{d(d+2)}\bar{J}_{d+1} = -\bar{J}_{d-1} + \bar{J}_{d-3}, \\ \text{(b')} \quad & \frac{\bar{J}_{d-3}}{\bar{J}_{d-1}} \geq \frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}}, \\ \text{(c')} \quad & \frac{\bar{J}_{d-1}}{\bar{J}_{d+1}} \leq \frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{d(d+2)}}, \\ \text{(d')} \quad & \bar{J}_{d+1}\bar{J}_{d-3} \geq \bar{J}_{d-1}^2. \end{aligned}$$

The following lemma lies at the heart of our inductive argument. It compares the standard Gaussian measure of centred and shifted Euclidean balls.

Lemma 3. *Let $d \geq 2$ and G be a standard Gaussian random vector in \mathbb{R}^d . For every $a \geq 0$, $R \geq \sqrt{d+2}$ and a vector $x \in \mathbb{R}^d$ of length $a\sqrt{d}$ we have*

$$\mathbb{P}(\|G\| \leq R) \leq \mathbb{P}(\|G - x\| \leq R\sqrt{1+a^2}).$$

Proof. Since for $a = 0$ we have equality, it is enough to show that the right-hand side,

$$h(a, R) = \mathbb{P}(\|G - a\sqrt{d}e_1\| \leq R\sqrt{1+a^2})$$

is nondecreasing with respect to a (by rotational invariance, for concreteness we can choose $x = ae_1$, where $e_1 = (1, 0, \dots, 0)$). Using Fubini's theorem we can write

$$h(a, R) = \frac{|S^{d-2}|}{\sqrt{2\pi}^{d-1}} \int_0^{R\sqrt{1+a^2}} \left\{ r^{d-2} e^{-r^2/2} \int_{a\sqrt{d}-\sqrt{R^2(1+a^2)-r^2}}^{a\sqrt{d}+\sqrt{R^2(1+a^2)-r^2}} \phi(t) dt \right\} dr,$$

where $\phi(t) = \frac{1}{\sqrt{2\pi}}e^{-t^2/2}$. The derivative with respect to a equals

$$\begin{aligned} \frac{\partial}{\partial a} h(a, R) &= \frac{|S^{d-2}|}{\sqrt{2\pi}^{d-1}} \int_0^{R\sqrt{1+a^2}} \left\{ r^{d-2} e^{-r^2/2} \right. \\ &\quad \cdot \left[\phi(a\sqrt{d} - \sqrt{R^2(1+a^2) - r^2}) \left(\sqrt{d} + \frac{aR^2}{\sqrt{R^2(1+a^2) - r^2}} \right) \right. \\ &\quad \left. \left. - \phi(a\sqrt{d} + \sqrt{R^2(1+a^2) - r^2}) \left(\sqrt{d} - \frac{aR^2}{\sqrt{R^2(1+a^2) - r^2}} \right) \right] \right\} dr. \end{aligned}$$

After changing the variables $r = R\sqrt{1+a^2}\sqrt{1-x^2}$ we see that this is nonnegative if and only if

$$\begin{aligned} \int_0^1 (1-x^2)^{\frac{d-3}{2}} \left[e^{-xa\sqrt{d}R\sqrt{1+a^2}} \left(x\sqrt{d} + \frac{aR}{\sqrt{1+a^2}} \right) \right. \\ \left. - e^{xa\sqrt{d}R\sqrt{1+a^2}} \left(x\sqrt{d} - \frac{aR}{\sqrt{1+a^2}} \right) \right] dx \geq 0. \end{aligned}$$

This condition can be further simplified by integration by parts using $\left((1-x^2)^{\frac{d-1}{2}} \right)' = -(d-1)x(1-x^2)^{\frac{d-3}{2}}$. We obtain an equivalent inequality

$$\int_0^1 (1-x^2)^{\frac{d-3}{2}} (d-1-d(1+a^2)(1-x^2)) \cosh\left(aR\sqrt{d}\sqrt{1+a^2}x\right) dx \geq 0.$$

Let $b = a\sqrt{d}R\sqrt{1+a^2}$. Then

$$1+a^2 = \frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{R^2d}}.$$

Observe that

$$\int_0^1 (1-x^2)^{\frac{d}{2}} \cosh(bx) dx = \frac{1}{2} \int_{-1}^1 (1-x^2)^{\frac{d}{2}} e^{bx} dx = \frac{1}{2} J_d(b).$$

Thus, the inequality we want to show becomes

$$\frac{J_{d-3}(b)}{J_{d-1}(b)} \geq \frac{d}{d-1} \left(\frac{1}{2} + \sqrt{\frac{1}{4} + \frac{b^2}{R^2d}} \right).$$

For a fixed b , the right-hand side as a function of R is clearly decreasing, so given our assumption $R \geq \sqrt{d+2}$ it is enough to consider $R = \sqrt{d+2}$, which follows from Lemma 2(b). \square

Remark. The statement for $d = 1$ remains true and was proved in [1], where it played a key role in the inductive proof of Pinelis' inequality (P).

The last lemma will help us use the spherical symmetry of our problem.

Lemma 4. *Let X be a rotationally invariant random vector in \mathbb{R}^d . Let $x \in \mathbb{R}^d$ and $t > 0$ be such that $t > \|x\|$. Then*

$$\mathbb{P}(\|X + x\| > t) = \mathbb{P}\left(\|X\| > -\theta\|x\| + \sqrt{t^2 + \theta^2\|x\|^2 - \|x\|^2}\right),$$

where θ is the first coordinate of an independent of X random vector uniformly distributed in the unit sphere S^{d-1} in \mathbb{R}^d .

Proof. Let ξ be an independent of X random vector uniformly distributed in the unit sphere S^{d-1} in \mathbb{R}^d . By rotational invariance X has the same distribution as $R\xi$, where $R = \|X\|$. We have

$$\mathbb{P}(\|X + x\| > t) = \mathbb{P}(R^2 + 2R\langle\xi, x\rangle + \|x\|^2 > t^2)$$

and by the rotational invariance of ξ , $\langle\xi, x\rangle$ has the same distribution as $\theta\|x\|$ with θ being the first coordinate of ξ . The inequality $R^2 + 2R\theta\|x\| + \|x\|^2 > t^2$ is equivalent to $R > -\theta\|x\| + \sqrt{t^2 + \theta^2\|x\|^2 - \|x\|^2}$ or $R < -\theta\|x\| - \sqrt{t^2 + \theta^2\|x\|^2 - \|x\|^2}$, but the second case does not hold as the right-hand side is negative, for we assume that $t > \|x\|$. \square

Proof of Theorem 1. For $d \geq 2$ we define $M_d = 1/\mathbb{P}(\|G_1\| > \sqrt{d+2})$. We fix $d \geq 2$ and show by induction on m that inequality (KO) holds with $C = M_d$. Then $C_d \leq M_d$, by Lemma 1, $M_d \leq e^2$ and by the central limit theorem, $\limsup_{d \rightarrow \infty} M_d = 2$.

For $m = 1$ we have to check that for $0 < t < |a_1|$ we have

$$1 \leq M_d \cdot \mathbb{P}\left(\left\|a_1 \frac{G_1}{\sqrt{d}}\right\| > t\right).$$

This follows because $\mathbb{P}(\|G_1\| > \sqrt{d}) \geq \mathbb{P}(\|G_1\| > \sqrt{d+2})$.

Suppose the assertion is true for $m \geq 1$. We shall show it for $m+1$. We can assume that the a_i are nonzero. By homogeneity we can also assume that $\sum_{i=2}^{m+1} a_i^2 = d$. If $t \leq \sqrt{d+2}\sqrt{\frac{a_1^2+d}{d}}$ we trivially bound the right-hand side as follows:

$$\mathbb{P}\left(\left\|\sum_{i=1}^{m+1} a_i \frac{G_i}{\sqrt{d}}\right\| > t\right) = \mathbb{P}\left(\left\|\sqrt{\frac{a_1^2+d}{d}} G_1\right\| > t\right) \geq \mathbb{P}(\|G_1\| > \sqrt{d+2}) = 1/M_d.$$

Now suppose $t > \sqrt{d+2}\sqrt{\frac{a_1^2+d}{d}}$. Notice that in particular $t > |a_1|$. Consider $v = \sum_{i=2}^{m+1} a_i \xi_i$. By independence and rotational invariance,

$$\mathbb{P}\left(\left\|\sum_{i=1}^{m+1} a_i \xi_i\right\| > t\right) = \mathbb{P}(\|a_1 e_1 + v\| > t).$$

Lemma 4 applied to $X = v$ and $x = a_1 e_1$ yields

$$\mathbb{P} \left(\left\| \sum_{i=1}^{m+1} a_i \xi_i \right\| > t \right) = \mathbb{E}_\theta \mathbb{P}_v \left(\|v\| > -\theta|a_1| + \sqrt{t^2 + \theta^2 a_1^2 - a_1^2} \right).$$

As a consequence, by the independence of θ and v , and the inductive hypothesis,

$$\mathbb{P} \left(\left\| \sum_{i=1}^{m+1} a_i \xi_i \right\| > t \right) \leq M_d \cdot \mathbb{E}_\theta \mathbb{P}_{(G_i)_{i=2}^{m+1}} \left(\left\| \sum_{i=2}^{m+1} a_i \frac{G_i}{\sqrt{d}} \right\| > -\theta|a_1| + \sqrt{t^2 + \theta^2 a_1^2 - a_1^2} \right).$$

The vector $\sum_{i=2}^{m+1} a_i \frac{G_i}{\sqrt{d}}$ has the same distribution as $\sqrt{\frac{\sum_{i=2}^{m+1} a_i^2}{d}} G_1 = G_1$. Therefore, applying again Lemma 4 yields

$$\mathbb{P} \left(\left\| \sum_{i=1}^{m+1} a_i \xi_i \right\| > t \right) \leq M_d \cdot \mathbb{P} (\|G_1 + a_1 e_1\| > t).$$

To finish the inductive step it suffices to show that

$$\mathbb{P} (\|G_1 + a_1 e_1\| > t) \leq \mathbb{P} \left(\left\| \sum_{i=1}^{m+1} a_i \frac{G_i}{\sqrt{d}} \right\| > t \right) = \mathbb{P} \left(\left\| \sqrt{\frac{a_1^2 + d}{d}} G_1 \right\| > t \right).$$

This follows from Lemma 3 applied to $a = \frac{|a_1|}{\sqrt{d}}$ and $R = t \sqrt{\frac{d}{a_1^2 + d}} > \sqrt{d+2}$, which completes the proof. \square

Proof of Theorem 2. Let ξ_1, ξ_2, \dots be independent random vectors uniformly distributed in the unit Euclidean sphere $S^{d-1} \subset \mathbb{R}^d$, independent of the sequence X_1, X_2, \dots . Since X_i is rotational invariant, it has the same distribution as $R_i \xi_i$, where $R_i = \|X_i\|$. Note that almost surely $0 \leq R_i \leq 1$. Applying (KO) (conditionally on the R_i) with $C = e^2$ we get

$$\begin{aligned} \mathbb{P} \left(\left\| \sum_{i=1}^m a_i X_i \right\| > t \right) &= \mathbb{E}_{(R_i)_{i=1}^m} \mathbb{P}_{(\xi_i)_{i=1}^m} \left(\left\| \sum_{i=1}^m a_i R_i \xi_i \right\| > t \right) \\ &\leq e^2 \cdot \mathbb{E}_{(R_i)_{i=1}^m} \mathbb{P}_{(G_i)_{i=1}^m} \left(\left\| \sum_{i=1}^m a_i R_i \frac{G_i}{\sqrt{d}} \right\| > t \right). \end{aligned}$$

To finish the proof notice that for any fixed numbers $R_i \in [0, 1]$ we have

$$\mathbb{P} \left(\left\| \sqrt{\frac{\sum_{i=1}^m a_i^2 R_i^2}{d}} G_1 \right\| > t \right) \leq \mathbb{P} \left(\left\| \sqrt{\frac{\sum_{i=1}^m a_i^2}{d}} G_1 \right\| > t \right). \quad \square$$

REFERENCES

- [1] Bobkov, S., Götze, F., Houdré, Ch., On Gaussian and Bernoulli covariance representations. *Bernoulli* 7 (2001), no. 3, 439–451.
- [2] Bentkus, V. K., Dzindzalieta, D., A tight Gaussian bound for weighted sums of Rademacher random variables. *Bernoulli* 21 (2015), no. 2, 1231–1237.
- [3] Latała, R., Oleszkiewicz, K., Small ball probability estimates in terms of width. *Studia Math.* 169 (2005), no. 3, 305–314.
- [4] Ledoux, M., Oleszkiewicz, K., On measure concentration of vector-valued maps. *Bull. Pol. Acad. Sci. Math.* 55 (2007), no. 3, 261–278.
- [5] Pinelis, I., Extremal probabilistic problems and Hotelling’s T^2 test under a symmetry condition. *Ann. Statist.* 22 (1994), no. 1, 357–368.
- [6] Talagrand, M., The missing factor in Hoeffding’s inequalities. *Ann. Inst. H. Poincaré Probab. Statist.* 31 (1995), no. 4, 689–702.

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