

Almost universal graphs

Alan Frieze* Michael Krivelevich†

December 30, 2005

Abstract

We study the question as to when a random graph with n vertices and m edges contains a copy of almost all graphs with n vertices and $cn/2$ edges, c constant. We identify a "phase transition" at $c = 1$. For $c < 1$, m needs to grow slightly faster than n , and we prove that $m = O(n \log \log n / \log \log \log n)$ is sufficient. When $c > 1$, m needs to grow at a rate $m = n^{1+a}$, where $a = a(c) > 0$ for every $c > 1$, and $a(c)$ is between $1 - \frac{2}{(1+o(1))c}$ and $1 - \frac{1}{c}$ for large enough c .

1 Introduction

1.1 Problem statement and results

A graph G is *universal* for a class of graphs \mathcal{H} if for every $H \in \mathcal{H}$, there is a subgraph of G which is isomorphic to H . The problem of constructing small G which are universal for interesting classes \mathcal{H} has attracted much attention as it arises in the study of VLSI circuit design. See for example [1] and the references there-in. This paper shows for example that if $\mathcal{H} = \mathcal{H}(c, n)$ is the class of graphs with vertex set $[n]$ and maximum degree c , then any \mathcal{H} -universal graph must contain $\Omega(n^{2-2/c})$ edges. On the other hand, it is shown in [1] that almost every graph with $(1+\epsilon)n$ vertices and $An^{2-1/c}(\log n)^{1/c}$ edges is $\mathcal{H}(c, n)$ -universal. Here A depends only on ϵ . Furthermore, the results of [2] prove the existence of an $\mathcal{H}(c, n)$ universal graph with $O(n)$ vertices and $O(n^{2-2/c}(\log n)^{1+8/c})$ edges.

*Department of Mathematical Sciences, Carnegie Mellon University, Pittsburgh PA15213, U.S.A. Supported in part by NSF grant CCR-0200945.

†Department of Mathematics, Raymond and Beverly Sackler Faculty of Exact Sciences, Tel Aviv University, Tel Aviv 69978, Israel. E-mail: krivelev@post.tau.ac.il. Research supported in part by USA-Israel BSF Grant 2002-133, and by grant 64/01 from the Israel Science Foundation.

Suppose now we consider $\mathcal{H} = \mathcal{H}^*(c, n)$ where $\mathcal{H}^*(c, n)$ is the set of all labeled graphs with vertex set $[n]$ and *average degree* at most c . Clearly $\mathcal{H}^*(c, n) \supseteq \mathcal{H}(c, n)$ but in fact an $\mathcal{H}^*(c, n)$ -universal graph requires $\Omega(n^{2-o(1)})$ edges since it must contain all graphs with $(1 - \epsilon)n$ isolated vertices and a $\lfloor c/\epsilon \rfloor$ -regular graph on the remaining ϵn vertices.

In this paper we relax the strict notion of universality. We say that G is *almost universal* for the class of graphs \mathcal{H} if it contains subgraphs isomorphic to all but $o(|\mathcal{H}|)$ graphs in \mathcal{H} . In particular we consider $\mathcal{H} = \mathcal{H}^*(c, n)$. As customary we denote by $\mathcal{G}_{n,m}$ the probability space of all labeled graphs on n vertices with m edges, where all such graphs are equiprobable, i.e., $\Pr[G] = \binom{\binom{n}{2}}{m}^{-1}$. Similarly, $\mathcal{G}_{n,p}$ stands for the probability space of all labeled graphs with vertex set $\{1, \dots, n\}$, where each pair $1 \leq i \neq j \leq n$ is an edge of a sample graph $G_{n,p}$ with probability $p = p(n)$, independently of all other edges. We study

$$\Pr(G_{n,m} \text{ is almost universal for } \mathcal{H}^*(c, n)). \quad (1)$$

More formally, we should perhaps consider

$$\Pi^*(c, m, \epsilon) = \Pr(|\{H \in \mathcal{H}^*(c, n) : G_{n,m} \supseteq H\}| \geq (1 - \epsilon)|\mathcal{H}^*(c, n)|) \quad (2)$$

where \supseteq denotes “contains a subgraph isomorphic to”.

We can reduce this to estimating the probability

$$\Pi(c, m) = \Pr(G_{n,m} \supseteq G_{n, cn/2}) \quad (3)$$

where $G_{n,m}$ and $G_{n, cn/2}$ are generated independently.

We can link definitions (2) and (3) via

$$1 - \epsilon^{-1}(1 - \Pi(c, m)) \leq \Pi^*(c, m, \epsilon) \leq \frac{\Pi(c, m)}{1 - \epsilon}. \quad (4)$$

To verify this, let $N = \binom{n}{2}$ and $M_1 = \binom{N}{m}$, $M_2 = \binom{N}{cn/2}$. Consider the $M_1 \times M_2$ matrix A where $A(i, j) = 1_{G_i \supseteq H_j}$, assuming that G_i , $i \in [M_1]$ (resp. H_j , $j \in [M_2]$) is an enumeration of all graphs with vertex set $[n]$ and m (resp. $cn/2$) edges. $\Pi^*(c, m, \epsilon)$ is the proportion of rows of A with at least $(1 - \epsilon)M_2$ 1's and $\Pi(c, m)$ is the proportion of entries of A which are 1. Equation (4) is now easy to verify.

We show that there is a sharp difference in $\Pi(c, m)$ for the cases $c < 1$ and $c > 1$ respectively. We prove the following:

Theorem 1

(a) *Suppose that $c < 1$ is constant. Then if A is constant,*

$$\Pi(c, m) \begin{cases} \leq 1 - (1 - e^{-c^3/6})e^{-A^3/6} + o(1), & m = An \\ = 1 - o(1), & m \geq \frac{C_0 \log \log n}{\log \log \log n} \end{cases} .$$

for some sufficiently large $C_0 = C_0(c)$.

(b) Suppose that $c > 1$ is constant. Then for some constants C_1, C_2 ,

$$\Pi(c, m) = \begin{cases} o(1), & m \leq C_1 n^{2-2/(c+x_c)} \\ 1 - o(1), & m \geq C_2 n^{2-1/(c-y_c)} \end{cases},$$

for some $x_c, y_c \rightarrow 0$ as $c \rightarrow \infty$.

□

We doubt that the upper bound in (a) is tight:

Conjecture If $c < 1$ and $m/n \rightarrow \infty$ then $\Pi(c, m) = 1 - o(1)$.

It is more difficult to guess whether the upper or lower bound is correct in (b).

1.2 m or p

In work on random graphs, it is usually more convenient to work in the independent model $G_{n,p}$ rather than in $G_{n,m}$. We therefore estimate

$$\Pi^\#(p_1, p_2) = \Pr(G_{n,p_2} \supseteq G_{n,p_1})$$

where G_{n,p_1} and G_{n,p_2} are generated independently.

Putting $p_1 = c/n$ and $p_2 = 2m/n$ we relate Π and $\Pi^\#$ through

$$\Pi(c, m) \geq \Pi^\# \left(p_1 + \frac{\log n}{n^{3/2}}, p_2 - \frac{m^{1/2} \log n}{n^2} \right) - o(1) \quad (5)$$

$$\Pi(c, m) \leq \Pi^\# \left(p_1 - \frac{\log n}{n^{3/2}}, p_2 + \frac{m^{1/2} \log n}{n^2} \right) + o(1). \quad (6)$$

The inequalities come from the fact that **whp** $G_{n,p_1 - \log n/n^{3/2}}$ contains fewer than $cn/2$ edges etc.

We break the proof of Theorem 1 into 4 pieces:

2 $c < 1, m = An$.

This is straightforward. It is well known (see, e.g., Theorem 3.19 of [6]) that the probability of $G_{n,(c+o(1))/n}$ not to contain a triangle is asymptotically equal to $e^{-c^3/6}$.

We thus get:

$$\Pr(G_{n,p_1-\frac{\log n}{n^{3/2}}} \text{ contains a triangle and } G_{n,p_2+\frac{\log n}{n^{3/2}}} \text{ is triangle free}) \\ = (1 - e^{-c^3/6})e^{-A^3/6} + o(1).$$

The upper bound in part (a) of the theorem now follows from this and (3)–(6). Having made this connection once, we will just focus on $\Pi^\#(p_1, p_2)$ from now on.

3 $c < 1$, $m = o(n \log \log n)$.

From now on let us shorten G_{n,p_i} to G_i for $i = 1, 2$.

We first state some almost sure properties of the random graph G_1 required in our proof.

Lemma 1 *Let $0 < c < 1$ be a constant. The random graph $G = G_{n,c/n}$ has **whp** the following properties:*

1. All connected components of G have size at most $L = (c - 1 - \log c)^{-1} \log n$;
2. All connected components of G are trees or unicyclic;
3. The number of vertices in unicyclic components of G does not exceed $\frac{\log \log n}{\log \log \log n}$;
4. The number of isolated vertices in G is at least n/e .

All of the above properties are quite standard and can be found, e.g., in [6].

Let $p_1 = c/n$, $m = \omega n/2$ and $p_2 = \omega/n$ where

$$\omega = \frac{120 \log \log n}{\log \log \log n}.$$

The following properties of G_1 are more technical and pertain closely to our arguments. They aim to estimate from above the expected number of vertices of G_2 needed to embed all large tree components of G_1 .

Let a vertex of G_1 be *large* if its degree is at least $\geq \omega/20$.

For a tree T of G_1 , let

$$\sigma(T) = \prod_{i=1}^{a(T)} (d_i - 1)!$$

where $d_1, d_2, \dots, d_{a(T)}$ are the degrees of the large vertices of T .

Lemma 2 Let $\omega/20 \leq k \leq L$, $a > 0$, σ, D be integers. Denote by $Y(k, a, \sigma, D)$ the number of isolated tree components T of G with $|V(T)| = k$, $a(T) = a$, $\sigma(T) = \sigma$ and $d_1 + \dots + d_a = D$, where d_1, \dots, d_a are the degrees of large vertices in T . Let $0 < c < 1$ be a constant. Then the random graph $G = G(n, c/n)$ has **whp** the following property:

$$\sum_{k \leq L, a, \sigma, D} Y(k, a, \sigma, D) k (1 - e^{-\omega/10})^{-k} e^{2\omega a} 3^{D-a} \sigma \omega^{-D+a} < \frac{n}{\log^2 n}.$$

Proof Let $k^* = k - 2 - D + a$. The number of trees T with parameters k, a, σ, D in the complete graph K_n can be estimated from above by:

$$\begin{aligned} & \binom{n}{k} \sum_{d_1 + \dots + d_a = D} \binom{k-2}{d_1-1, d_2-1, \dots, d_a-1, k^*} (k)_a (k-a)^{k^*} \\ & \leq \frac{(1+o(1))n^k}{k!} \frac{(k-2)! k^a (k-a)^{k^*}}{k^*!} \sum_{d_1 + \dots + d_a = D} \prod_{i=1}^a \frac{1}{(d_i-1)!} \\ & < (1+o(1)) n^k k^a \binom{D-1}{a-1} \sigma^{-1} e^{k-a}. \end{aligned}$$

(In the first line above, we choose k vertices of the tree in $\binom{n}{k}$ ways, then choose names of a large vertices in $(k)_a$ ways, then decide which positions in the Prüfer code of the tree are occupied by each of the large vertices, then fill the rest of the positions by remaining vertices in $(k-a)^{k^*}$ ways.) For each such tree, the probability that it is an isolated component in $G(n, c/n)$ is $(c/n)^{k-1} (1-c/n)^{k(n-k)} \leq (1+o(1))(c/n)^{k-1} e^{-ck}$. Using the linearity of expectation we derive that the expected value of the random variable in the lemma's formulation is at most:

$$\begin{aligned} & (1+o(1)) \sum_{k, a, \sigma, D} n^k k^a \binom{D-1}{a-1} \sigma^{-1} e^{k-a} (c/n)^{k-1} e^{-ck} k (1 - e^{-\omega/10})^{-k} e^{2\omega a} 3^{D-a} \sigma \omega^{-D+a} \\ & \leq (1+o(1)) \frac{n}{c} \sum_{k, a, \sigma, D} (ce^{1-c})^k (k/e)^a 2^D k (1 + e^{-\omega/9})^k e^{2\omega a} 3^{D-a} \omega^{-D+a} \\ & < 2n \sum_{k, a, \sigma, D} k ((1+o(1))ce^{1-c})^k (ke^{2\omega} w/3e)^a (6/\omega)^D. \end{aligned}$$

Next observe that $D \geq a\omega/20$ and that there are at most $\binom{D-1}{a-1} \leq 2^D$ possible values for σ , given a, D . Thus the last expression can be bounded by

$$\begin{aligned} & 2n \sum_{k, a, D} k ((1+o(1))ce^{1-c})^k ((ke^{2\omega} w/3e)^{20/\omega})^D (12/\omega)^D \\ & \leq 2n \sum_{k, a, D} k ((1+o(1))ce^{1-c})^k (k^{20/\omega} e^{50})^D (12/\omega)^D. \end{aligned}$$

Observe that $k^{20/\omega} \ll \omega^{1/2}$. Indeed, $k \leq L = O(\log n)$, while

$$w^{\omega/40} = \left(\frac{120 \log \log n}{\log \log \log n} \right)^{\frac{120 \log \log n}{40 \log \log \log n}} = (\log n)^{3+o(1)}.$$

Therefore, $(k^{20/\omega} e^{50})^D (12/\omega)^D \ll (12\omega^{1/2}/\omega)^D \leq \omega^{-(1+o(1))\omega/20} = (\log n)^{-6+o(1)}$. Also, for $c < 1$ one has $ce^{1-c} < 1$, and thus $k(ce^{1-c})^k = O(1)$. Since $k \leq L = O(\log n)$, $D \leq 2k$, $a \leq k$, we have altogether $O(\log^3 n)$ summands, each at most $(\log n)^{-6+o(1)}$, hence the sum is at most $(\log n)^{-3+o(1)}$. It follows by the Markov Inequality that the random variable of the lemma **whp** has value at most $n/\log^2 n$.

□

Lemma 3 *For an integer $k \geq 2$, denote $\pi_k = (1 - e^{-\omega/10})^{k-1}$. Let τ_k be the number of isolated tree components of size k in G . Then, for every $0 < c < 1$, **whp** in $G(n, c/n)$*

$$\sum_{k=(\log \log n)^2}^L \frac{k\tau_k}{\pi_k} \leq \frac{n}{\log^2 n}.$$

Proof First we estimate the expectation of the expression in question:

$$\begin{aligned} & \mathbf{E} \left(\sum_{k=(\log \log n)^2}^L \frac{k\tau_k}{\pi_k} \right) \\ &= \sum_{k=(\log \log n)^2}^L \binom{n}{k} k^{k-1} p_1^{k-1} (1-p_1)^{k(n-k)+\binom{k}{2}-k+1} \pi_k^{-1} \\ &= \sum_{k=(\log \log n)^2}^L \binom{n}{k} k^{k-1} \bar{p}_1^{k-1} (1-\bar{p}_1)^{k(n-k)+\binom{k}{2}-k+1} (1+\epsilon_1)^{k(n-k)+\binom{k}{2}-k+1} \end{aligned}$$

where

$$\bar{p}_1 = p_1(1 - e^{-\omega/10})^{-1} \text{ and } \epsilon_1 = \frac{p_1 e^{-\omega/10}}{1 - e^{-\omega/10} - p_1} = \frac{c e^{-\omega/10}}{n - c - n e^{-\omega/10}} \leq \frac{2c e^{-\omega/10}}{n}.$$

Thus, writing $\bar{p}_1 = \bar{c}_1/n$ we see that

$$\mathbf{E} \left(\sum_{k=(\log \log n)^2}^L \frac{k\tau_k}{\pi_k} \right) \leq \frac{n}{\bar{c}_1} \sum_{k=(\log \log n)^2}^L (\bar{c}_1 e^{1-\bar{c}_1+o(1)})^k,$$

and since $1 - \bar{c}_1 e^{1-\bar{c}_1+o(1)}$ is bounded below by a positive function of c ,

$$\mathbf{E} \left(\sum_{k=(\log \log n)^2}^L \frac{k\tau_k}{\pi_k} \right) = o(n/\log^2 n).$$

Applying the Markov inequality, the lemma follows. \square

Suppose now that G_1 consists of isolated trees T_1, T_2, \dots, T_s and unicyclic components K_1, K_2, \dots, K_t . We will assume that G_1 satisfies the conditions stated in Lemmas 1–3. We will assume that $a(T_i) \neq 0$ iff $i \leq r$. Note that $r \leq 21cn/\omega = o(n)$.

We try to embed the trees one by one in G_2 . Our strategy for embedding a tree T_i into G_2 is as follows: Choose a vertex v_1 of degree one of T_i and then let the vertices of T_i be v_1, v_2, \dots, v_k where the order comes from some breadth first search of T_i . Let d_j be the degree of v_j in T_i . Suppose that the embedding of trees T_1, T_2, \dots, T_{i-1} has involved *examining* vertices w_1, w_2, \dots, w_ℓ . Being a subset of $[n]$, $U = [n] \setminus \{w_1, w_2, \dots, w_\ell\}$, the set of unexamined vertices, has a natural order. Let $w_{\ell+1}$ be the first member of U in this order. We use the following algorithm to try to embed T_i into G_2 . Basically, at each stage we try to embed v_i as $w_{\ell+i}$ by finding $d_i^* = d_i - 1 + 1_{i=1}$ new neighbours for $w_{\ell+i}$ from which to continue the embedding.

Begin

$\lambda = \ell + 1$.

For $i = 1, 2, \dots, k$ **do**

Begin

If $w_{\ell+i}$ has $\geq d_i^*$ neighbours in U **then**

Begin

Let $x_j, j = 1, 2, \dots, d_i^*$, be the first d_i^* neighbours of $w_{\ell+i}$ in U , in the natural order.

$U \leftarrow U \setminus \{x_j : j = 1, 2, \dots, d_i^*\}$

$w_\kappa \leftarrow x_i, \kappa = \lambda + 1, \lambda + 2, \dots, \lambda + d_i^*$.

$\lambda \leftarrow \lambda + d_i^*$.

End

Else FAIL

End

End

If we fail then we repeat this procedure until either we succeed or ℓ reaches $n_F = n(1 - e^{-1})$, in which case we abandon the attempt to embed G_1 into G_2 .

3.1 Phase 1: Embedding trees with large vertices

Consider first the embedding of the first r trees, i.e. those with $a(T) > 0$. At the end of this phase of the embedding we expect ℓ to be at most $O(n/(\log n)^2)$ and consider the attempt to be a failure if ℓ reaches $n/(\log n)^{1/2}$ before T_r has been embedded.

Suppose the tree $T_i, i \leq r$, has the same parameters as in Lemma 2 and the degrees of the large vertices are d_1, d_2, \dots, d_a . Let us condition on having consumed less than $n/2$ vertices before starting to embed T_i . The number of G_2 neighbours of a vertex in

U dominates $\text{Bin}(n/2, p_2)$ and so the probability that this number is at least $\omega/20$ is at least $1 - e^{-\omega/10}$. Thus the probability of success in any attempt is at least

$$(1 - e^{-\omega/10})^k \prod_{i=1}^a \psi(d_i)$$

where $\psi(d_i)$ is the probability that a vertex has $\geq d_i - 1$ G_2 neighbours in U . Now

$$\psi(d_i) \geq \binom{n/2}{d_i - 1} p_2^{d_i - 1} (1 - p_2)^n \geq \frac{\omega^{d_i - 1}}{3^{d_i - 1} (d_i - 1)!} e^{-2\omega}$$

and then

$$\prod_{i=1}^a \psi(d_i) \geq \frac{\omega^{D-a}}{3^{D-a} \sigma(T)} e^{-2\omega a}.$$

So the expected time (increase in ℓ) while embedding this tree is at most

$$k(1 - e^{-\omega/10})^{-k} e^{2\omega a} 3^{D-a} \sigma(T) \omega^{-D+a}.$$

This explains the expression we put earlier in Lemma 2.

Let X_i be the number of vertices consumed while embedding T_i (or reaching a failure while embedding T_i). Conditioning on $X_1 + \dots + X_{i-1} \leq n/2$, the random variable X_i is dominated by k times the geometric random variable with probability of success

$$\rho_i = (1 - e^{-\omega/10})^k \prod_{i=1}^a \psi(d_i) \geq (1 - e^{-\omega/10})^k \frac{\omega^{D-a}}{3^{D-a} \sigma(T_i)} e^{-2\omega a}.$$

Denote $\mu_i = k/\rho_i$. Then

$$\Pr[X_i \geq (\log^{1.5} n) \mu_i | X_1 + \dots + X_{i-1} \leq \frac{n}{2}] \leq (1 - \rho_i)^{\log^{1.5} n / \rho_i} = o(1/n).$$

Observe that by Lemma 2, **whp**,

$$\sum_{i=1}^r \mu_i \leq \frac{n}{\log^2 n}.$$

It thus follows:

$$\begin{aligned} & \Pr[X_1 + \dots + X_r \geq \frac{n}{\log^{1/2} n}] \\ & \leq o(1) + \Pr[X_1 + \dots + X_r \geq \log^{1.5} n \sum_{i=1}^r \mu_i] \\ & \leq o(1) + \Pr \left[\bigcup_{i=1}^r \left(X_i \geq (\log^{1.5} n) \mu_i | X_1 + \dots + X_{i-1} \leq \frac{n}{2} \right) \right] \\ & = o(1) + o(r/n) \\ & = o(1). \end{aligned}$$

Thus, **whp** Phase 1 completes successfully, having used at most $n/(\log n)^{1/2}$ vertices, and the remaining trees of G_1 have their maximum degree at most $\omega/20$.

3.2 Phase 2: Embedding trees with no large vertices

We use the same embedding algorithm as before and a similar argument to analyze it. Note that **whp** $\sum_{i=r_1+1}^s |V(T_i)| \leq (1-2\alpha)n$ where $\alpha = e^{-1}/2$. We declare Phase 2 to be a failure if it uses at least $(1-\alpha)n$ vertices. At each stage of the embedding $|U| \geq \alpha n/2$ for a tree T_i with k vertices, the probability of a successful attempt is at least

$$\Pr\left(\text{Bin}\left(\frac{\alpha n}{2}, p_2\right) > \omega/20\right)^{k-1} \geq \pi_k = (1 - e^{-\omega/10})^{k-1}. \quad (7)$$

Let τ_k denote the number of trees of size k in G_1 and let $Z_{i,k}$ denote the number of vertices used in attempts to embed $T_{i,k}$, the i th tree of size k . Then

1. $Z_{i,k}$ is dominated by k times the geometric random variable $\Gamma_{i,k}$ which has probability of success π_k , see (7). We can couple $Z_{i,k}$ with a copy of $\Gamma_{i,k}$ so that $Z_{i,k} \leq \Gamma_{i,k}$.
2. If $Z = \sum_{k=2}^L \sum_{i=1}^{\tau_k} Z_{i,k} < (1-\alpha)n$ then the embedding succeeds.

Let $Z_2 = \sum_{k=(\log \log n)^2}^L \sum_{i=1}^{\tau_k} Z_{i,k}$. Observe that by Lemma 3, the expectation of Z_2 does not exceed $n/\log^2 n$. Thus we can apply the argument similar to that of Phase 1 to show that **whp** it takes $O(n/\log^{1/2} n)$ vertices to embed tree components of G_1 with at least $(\log \log n)^2$ vertices.

We now turn to $Z_1 = \sum_{k=2}^{(\log \log n)^2} \sum_{i=1}^{\tau_k} Z_{i,k}$. Consider the number of failures as we try to embed trees with at most $(\log \log n)^2$ vertices. Let $\nu = \frac{n}{(\log \log n)^3}$. The probability of failing at any attempt is at most $1 - \pi_{(\log \log n)^2} \leq (\log \log n)^2 e^{-\omega/10} < e^{-\omega/20}$. Furthermore, there will be less than n attempts embedding a tree and so the probability that there are ν failed attempts or more is at most

$$\binom{n}{\nu} (e^{-\omega/20})^\nu \leq \left(\frac{ne^{-\omega/20}}{\nu}\right)^\nu = o(1).$$

Each failed attempt consumes at most $(\log \log n)^2$ vertices. Thus **whp**

$$Z_1 = \sum_{k=2}^{(\log \log n)^2} \sum_{i=1}^{\tau_k} k Z_{i,k} = \sum_{k=2}^{\omega} k \tau_k + O(n/\log \log n) < (1-2\alpha)n.$$

Hence we conclude that **whp** $Z = Z_1 + Z_2 < (1-\alpha)n$ and the embedding succeeds.

3.3 Phase 3: Embedding unicyclic components

Recall that by Lemma 1 G_1 contains at most $\omega/120$ vertices altogether in unicyclic components. We divide the unused vertices of U into two sets U_1, U_2 of approximately equal size, which **whp** are at least $\alpha n/2$. Then **whp** U_1 will, in G_2 , contain at least ω cycles of size j for each $3 \leq j \leq \omega$. So, we will **whp** be able to choose a collection of cycles in U_1 to match with the cycles of G_1 . We can then use U_2 to embed the trees attached to these cycles which go to make up each of the unicyclic components. Since these trees are without large vertices, we use the analysis of the previous section and argue that the expected time to embed these trees is $o(n)$. This completes the analysis for $c < 1$. \square

4 $c > 1, m \leq C_1 n^{2-2/(c+x_c)}$

Let $x = x_c$ be the unique solution in $(0, 1)$ to $xe^{-x} = ce^{-c}$. It is easy to verify that $c + x_c > 2$ for all $c > 1$. Let $\alpha = 1 - \frac{x}{c}$ and $\beta = \frac{c}{2} \left(1 - \frac{x^2}{c^2}\right)$. **Whp** $G_{n, cn/2}$ contains a giant component with A vertices and B edges, where $|A - \alpha n|, |B - \beta n| = O(n^{1/2} \log n)$ vertices, [3].

We associate with each graph G_1 in $G_{n, \frac{cn}{2}}$, containing such a giant component, a graph H with A vertices and B edges. Then

$$\begin{aligned} \Pr[G_{n,m} \supseteq G_1] &\leq \Pr[G_{n,m} \supseteq H] \leq (n)_A \left((1 + o(1)) \frac{2m}{n^2} \right)^B \\ &\leq n^{\alpha n + O(\sqrt{n} \log n)} \left(\frac{2m}{n^2} \right)^{\beta n + O(\sqrt{n} \log n)} \\ &= \left[n^{(1+o(1))\alpha} \left(\frac{2m}{n^2} \right)^{(1+o(1))\beta} \right]^n. \end{aligned}$$

Hence if $m \leq c_1 n^{2-2/(c+x_c)}$ for $c_1 > 0$ small enough, the above expression tends to 0 as n grows, implying that **whp** $G_{n,m}$ does not contain most of the graphs from $G_{n, cn/2}$.

Comments. (a) Since $c + x_c > 2$, we get that for every $c > 1$ there exists $\epsilon = \epsilon(c) > 0$ such that $\Pi(c, m) = o(1)$ for $m \leq n^{1+\epsilon}$. This, combined with the result of Theorem 1 (a), shows that the function $\Pi(c, m)$ has sharp threshold at $c = 1$;

(b) As pointed out by the referee, we can get a somewhat weaker bound $\Pi(c, m) = o(1)$ for $m = n^{2-\frac{2}{c}-o(1)}$ by the following simple counting argument: the number of subgraphs with $\frac{cn}{2}$ edges of a graph G_2 with m edges is $\binom{m}{\frac{cn}{2}}$. The total number of non-isomorphic

graphs on n vertices with $\frac{cn}{2}$ edges is at least $\binom{\binom{n}{2}}{\frac{cn}{2}}(n!)^{-1}$. Thus, if

$$\binom{m}{\frac{cn}{2}} = o\left(\binom{\binom{n}{2}}{\frac{cn}{2}}(n!)^{-1}\right),$$

any graph G with m edges does not contain most of the graphs with $cn/2$ edges. Solving the above inequality for m we get the claimed bound.

5 Upper bound

We now show that if $c > 1$ and $p_2 = n^{-1/((1+o(1))c)}$ then **whp** $G_{n,p_2} \supseteq G_{n,p_1}$.

We note first that Pittel, Spencer and Wormald [7] have shown that the threshold probability c_k/n for having a k -core satisfies $c_k = k + \sqrt{k \log k} + O(\log k)$ and so for large c we find that **whp** $G_{n,c/n}$ is d -degenerate for $d = c - \sqrt{c \log c} + O(\log c)$. (A graph is d -degenerate if its vertices can be ordered as v_1, v_2, \dots, v_n so that v_i has at most d neighbours in $\{v_1, v_2, \dots, v_{i-1}\}$.) Also, G_1 is $K_{2,3}$ -free **whp**. Finally, **whp** G_1 has $(e^{-c} + o(1))n$ isolated vertices. We will therefore be able to use the following:

Theorem 2 *Let $d > 3, 0 < c_0 < 1$ be constants. Let $\delta = \max\left\{\frac{1}{d-1}, \frac{1}{d(d-3)}\right\}$. Let $p(n) = An^{(-1+\delta)/d}$. Let H be a d -degenerate $K_{2,3}$ -free graph on c_0n vertices, of maximum degree $\Delta(H) \leq \Delta_0 = 1/(4dp)$. If $A = A(c_0, d)$ is large enough, then **whp** the random graph $G(n, p)$ contains a copy of H .*

Proof Fix an ordering $\sigma = (v_1, \dots, v_{c_0n})$ of the vertices of H such that every vertex of H has at most d neighbors preceding it in σ . Denote $N_H^-(v_i) = \{v_j : j < i, (v_i, v_j) \in E(H)\}$. We will embed H vertex by vertex, according to σ .

Partition the vertex set of $G(n, p)$ into two parts: V_0 of size $|V_0| = \frac{1-c_0}{2}n$, and V_1 . We will use V_0 to embed the first $\nu_0 = A_1n^\delta$ vertices of σ , and V_1 to embed the rest, the value of $A_1 = A_1(d)$ will be chosen later to satisfy inequality (8).

Observe that if A is large enough then **whp** in $G(n, p)$ every d vertices of V_0 have at least A_1n^δ common neighbors in V_0 . Thus we can easily embed the first A_1n^δ vertices of H according to σ in V_0 .

Now, we will use Theorem 2 of Fernandez de la Vega and Manoussakis [4] (or rather its proof) to embed the rest of H in V_1 . We start by adding some edges to it to form a new graph H' . Specifically, for each $i > A_1n^\delta$, if $v_i \in V(H)$ has less than d neighbors preceding it in H , we add to H $d - |N_H^-(v_i)|$ edges connecting v_i to random vertices before it in the order. Let H' be the (random) graph obtained. Denote by U_i the set of neighbors of v_i in H' preceding v_i .

Let $A_1 n^\delta < i < j$. We will estimate the probability of $U_i = U_j$. Assume $|N_H^-(v_i)| = s_1$, $|N_H^-(v_j)| = s_2$, $|N_H^-(v_i) \cap N_H^-(v_j)| = t$. We first expose $d - s_1$ random neighbors of v_i ; for $U_i = U_j$ they should contain all $s_2 - t$ vertices of $N_H^-(v_j) \setminus N_H^-(v_i)$, this happens with probability at most:

$$\frac{\binom{i-1-s_2+t}{d-s_1-s_2+t}}{\binom{i-1}{d-s_1}} < \frac{\binom{i-1}{d-s_1-s_2+t}}{\binom{i-1}{d-s_1}} \leq \frac{(1+o(1))(d-s_1)^{s_2-t}}{i^{s_2-t}}.$$

Now expose $d - s_2$ random neighbors of v_j , they should coincide with $U_i \setminus N_H^-(v_j)$, the probability of this to happen is

$$\frac{1}{\binom{j-1}{d-s_2}} = \frac{(1+o(1))(d-s_2)!}{j^{d-s_2}}.$$

So altogether the probability that $U_i = U_j$ is less than $\frac{d^d}{i^{d-t}}$.

Recall that H is $K_{2,3}$ -free and thus $t \leq 2$. Also, $t > 0$ only for those $i < j$ who have a common neighbor in front of them in σ , and this happens for at most $d\Delta(H)$ values of j , for a given i . Therefore the probability that there exist i, j such that $U_i = U_j$ is at most:

$$\begin{aligned} & \sum_{\substack{A_1 n^\delta < i < j \\ N_H^-(v_i) \cap N_H^-(v_j) \neq \emptyset}} \frac{d^d}{i^{d-2}} + \sum_{\substack{A_1 n^\delta < i < j \\ N_H^-(v_i) \cap N_H^-(v_j) = \emptyset}} \frac{d^d}{i^d} \\ & \leq d^d \left[d\Delta(H) \sum_{i > A_1 n^\delta} \frac{1}{i^{d-2}} + n \sum_{i > A_1 n^\delta} \frac{1}{i^d} \right] \\ & < d^d \left[\frac{d\Delta(H)}{(d-3)(A_1 n^\delta)^{d-3}} + \frac{n}{(d-1)(A_1 n^\delta)^{d-1}} \right] \\ & < \frac{1}{2} \end{aligned} \tag{8}$$

for large enough $A_1 = A_1(d)$.

Now we argue that **whp**

$$\text{each } U_k \text{ intersects at most } 2d\Delta_0 \text{ of the sets } U_j \text{ with } j < k. \tag{9}$$

Clearly $N_H^-(v_k)$ intersects at most $d\Delta_0$ sets $N_H^-(v_j)$ with $j < k$. We thus need to show that **whp** the random edges of $H' - H$ do not add $d\Delta_0$ sets U_j , $j < k$, intersecting U_k , for any given k . To do so, we fix v_k , condition on U_k , fix $u \in U_k$. Then the number Z_u of vertices v_j that choose to add a random edge (v_j, u) is dominated by $X_1 + X_2 + \dots + X_n$ where X_1, X_2, \dots, X_n are independent and X_i is 0/1 and $\Pr(X_i = 1) \leq \frac{d}{v_0 + i - 1}$ for $i = 1, 2, \dots, n$. Then $\mathbf{E}(X_1 + X_2 + \dots + X_n) \leq O(\log n)$ and by Theorem 1 of Hoeffding [5] we see $\Pr(Z_u \geq \Delta_0)$ is exponentially small. We conclude then that the probability that (9) is violated is less than 1/2, too. Thus, when $A > 0$ is large enough, there exists a supergraph $H' \supseteq H$ satisfying

(i) For each pair $A_1 n^\delta \leq i < j$, $U_i \setminus U_j \neq \emptyset$.

(ii) Each U_k intersects at most $2d\Delta_0$ sets $U_j, j \neq k$.

Now consider continuing an embedding of v_1, v_2, \dots, v_{j-1} in G_2 where $j > \nu_0$. As in [4] we define, for each $i < j$ a vertex $x_i \in U_i \setminus U_j$. Suppose our embedding has assigned $v_i \rightarrow w_i$ for $i < j$ and that $W = [n] \setminus \{w_1, w_2, \dots, w_{j-1}\}$. Let the correspondence v_i, w_i map U_i to W_i and for $w \in W$ let $N_2(w)$ be the neighbours of w in G_2 . Suppose our embedding algorithm checks each $w \in W$ as a candidate in increasing value of w . Let $J = \{i < j : U_i \cap U_j \neq \emptyset\}$. Then arguing as in [4] we see

$$\Pr(N_2(w) \supseteq W_j \mid \text{history of process}) = \Pr(N_2(w) \supseteq W_j \mid N_2(w) \not\supseteq W_i, i \in J')$$

for some $J' \subseteq J$

$$\begin{aligned} &\geq \Pr(N_2(w) \supseteq W_j \text{ and } x_i \notin N_2(w), i \in J') \\ &\geq p^d (1-p)^{2d\Delta_0} \\ &\geq p^d/2. \end{aligned}$$

Thus,

$$\Pr(\exists w : N_2(w) \supseteq W_j \mid \text{history of process}) \leq (1 - p^d/2)^{(1-c_0)n/2} \leq e^{-An^\delta/4}.$$

Thus **whp** the embedding succeeds for all $j \leq c_0 n$. □

We apply the above theorem with H equal to G_1 minus its isolated vertices. Then **whp** H has at most $(1 - e^{-c} + o(1))n$ vertices. Finally, we will need

$$p_2 = n^{-c^{-1}(1+\sqrt{(\log c)/c+O(c^{-1})})}$$

which is somewhat better than claimed in Theorem 1(b).

Comment 1 *As pointed out by the referee, the techniques of [2] can be used to get an explicit construction of a graph G with $\tilde{O}(n^{2-1/(c'+1)})$ edges, where $c' = \lceil c \rceil$, containing almost all graphs with $cn/2$ edges. This is about the same as the upper bound of $n^{2-(1+o(1))/c}$ derived in this paper, our argument shows however that **almost all** graphs with that many edges are almost universal.*

Comment 2 *At a point quite late in the reviewing process, Andrzej Ruciński became cognisant of and reminded us of the relevance of the paper by Riordan [8]. Theorem 2.1 of that paper is similar to Theorem 1(b). He shows that if*

$$\gamma(G) = \max_{H \subseteq G} \{|E(H)| / (|V(H) - 2)\}$$

then our $1/(c - y_c)$ can be replaced by the likely value of $1/\gamma(G_{n,p})$. Now the value for $\gamma(G_{n,p})$ is not easy to estimate although it does seem likely that $\gamma(G_{n,p}) = c + o(c)$. Furthermore, Riordan's proof is very different relying on the second moment method, whereas ours is constructive.

Acknowledgement. The authors wish to thank Itai Benjamini for posing this problem. They are also thankful to the anonymous referee and to Andrzej Ruciński for their helpful criticism on the first version of the paper.

References

- [1] N. Alon, M. Capalbo, Y. Kohayakawa, V. Rödl, A. Ruciński and E. Szemerédi, Universality and Tolerance, *Proceedings of the 41st IEEE Symposium on Foundations of Computer Science*, IEEE (2000), 14-21.
- [2] N. Alon, M. Capalbo, Y. Kohayakawa, V. Rödl, A. Ruciński and E. Szemerédi, Near-optimum universal graphs for general bounded-degree graphs, *Proceedings of the conference on approximation, randomization and combinatorial optimization (RANDOM-APPROX'2001)*, Lecture Notes in Computer Science 2129, Springer (2001), 170–180.
- [3] P. Erdős and A. Rényi, *On the evolution of random graphs*, Publ. Math. Inst. Hungar. Acad. Sci. 5 (1960) 17-61.
- [4] W. Fernandez de la Vega and Y. Manoussakis, Grids in random graphs, *Random Structures and Algorithms* 5 (1994) 329-338.
- [5] W. Hoeffding, Probability inequalities for sums of bounded random variables, *Journal of the American Statistical Association* 58 (1963) 13-30.
- [6] S. Janson, T. Łuczak and A. Ruciński, **Random graphs**, Wiley, New York, 2000.
- [7] B. Pittel, J. Spencer and N. Wormald, *Sudden emergence of a giant k -core in a random graph*, Journal of Combinatorial Theory (B) **67** (1996) 111–151.
- [8] O. Riordan, Spanning subgraphs of random graphs, *Combinatorics, Probability and Computing* 9 (200) 125-148.